## TITLE PAGE

# ASSESSMENT OF THE EQUIVALENCE OF WAEC AND NECO MATHEMATICS MULTIPLE-CHOICE TESTS USING ITEM RESPONSE THEORY 

BY

OGUOMA, CHINYERE CHINWEIKE 2010177002F

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## CERTIFICATION

This is to certify that Oguoma Chinyere Chinweike with Registration Number 2010177002F is responsible for the work submitted in this dissertation, that the work is original except as specified in the acknowledgements and references.

Oguoma, Chinyere Chinweike

## APPROVAL PAGE

This Dissertation has been approved for the Department of Educational foundations, Faculty of Education Nnamdi Azikiwe University, Awka.

Prof. Romy Okoye
Supervisor

Prof. N.N. Agu
Head of Department

Prof. O.S. Abonyi
External Examiner

Prof. O. Ibeneme
Dean, Faculty of Education

Date
Dean, School of Post Graduate Studies

## DEDICATION

This research work is dedicated to my husband Hon. Chinweike Oguoma.

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#### Abstract

The study assessed the equivalence of West African Examination Council and the National Examination Council mathematics multiple-choice tests, from 2011 to 2014 using item response theory (IRT). The study was necessitated by the recurring variation in performance levels of candidates in the examinations conducted by the two examining bodies. Six (6) research questions and two (2) hypotheses guided the study with descriptive survey as the research design. The study was carried out in Imo-State. A sample of 1051 students in SS 3 was chosen through a combination of non-proportionate random sampling and cluster sampling techniques for the study. The research instruments were the mathematics multiple-choice tests administered by WAEC and NECO in 2011, 2012, 2013 and 2014. The pair of NECO 2011 and WAEC 2011 tests were separately administered within one week simultaneously in the 30 selected schools. The NECO 2012 and WAEC 2012, NECO 2013 and WAEC 2013 and NECO 2014 and WAEC 2014 tests were similarly administered at two weeks interval between each pair. The method of data analysis for the study involved the use of factor analysis model of SPSS version 21 for research question 1, tetrachoric correlation module of LISREL version 8.8 and frequency count for research questions 2 and 3, calibration module of BILOG MG version 3.0 for research question 4, empirical reliability of the calibration model, of BILOG MG version 3.0 for research question 5 and linear equating transformation equations and scatter plots for research question 6 . The Z-test of correlation analysis was used to test the null hypotheses. The major findings of the study were that (v) NECO and WAEC mathematics tests of 2011, 2012, 2013, and 2014 did not fulfill all the conditions that are required for test scores obtained from two tests designed to measure the same ability of examinees to be used interchangeably, therefore they are not equivalent. Based on the findings, it was recommended among others that (i). Education authorities should review their stands on the equivalence by government fiat placed on the two examinations conducted by WAEC and NECO mathematics test-items.


## CHAPTER ONE

## INTRODUCTION

## Background to the Study

Evaluation is a major educational tool used in the identification of individual talents and skills for placement of students at appropriate learning programmes or vocations. This, in effect, is the idea behind the establishment of evaluation agencies which also act as examining bodies that maintain a common standard in test development and administration of public examinations. Nworgu (2006) maintains that these agencies were set up to promote education, coordinate educational programmes, control and monitor the quality of educational institutions.

These agencies organize public examinations which provide uniform standards to all test takers, irrespective of the type or method of instruction they have received. Some of these examination bodies in Nigeria are; the West African Examination Council (WAEC), National Examination Council (NECO), Joint Admissions and Matriculation Board (JAMB), National Business and Technical Examination Board (NABTEB) among others. A close look at the functions of these boards reveals that some of them perform similar functions. For instance, WAEC, NECO, and NABTEB all conduct senior secondary school leavers' certificate examinations but in the case of NABTEB, the examination is targeted only at secondary school leavers of technical and vocational colleges in Nigeria.

Nkwocha (2015) states that, just three examining bodies; WAEC, NECO and NABTEB are responsible for the award of senior school certificate in Nigeria. They conduct parallel or equivalent senior school certificate examinations in the country. They maintain high standards in the development and administration of the examinations, of which performances in the examinations should be good indicators of individuals' standing in any subject area of interest. Both examining bodies follow uniform mode of test construction, following the rigorous standardization procedures, administration, scoring and interpretation.

WAEC was established in 1952 by acts of the British West African Colonies with headquarters in Accra (Ghana) and national offices in the capital cities, viz: Lagos (Nigeria), Freetown (Sierra-Leone) and Bathurst (nowBanjul Gambia), and each national office is headed by a very senior WAEC staff from that country.

WAEC carries out the following major examining functions:-
i. Conducting national examinations which are unique to and required by each of the member countries. These include common entrance examinations to secondary schools, nurses selection tests, entrance examinations to the Nigeria Defence Academy and variety of selection tests for various trades and vocations.
ii. Conducting international examinations: It includes principally the West African Senior School Certificate Examination (WASSCE) in the

English speaking West African countries of Nigeria, Ghana, SierraLeone, Gambia and Liberia.
iii. Conducting examinations in collaboration with other examining bodies: These include the General Certificate of Education (GCE) of the University of London School Examination Council (popularly known as the 'O' Level) as well as the Advanced Level Examinations ('A' level) of the same body. They also include the examinations of the Royal Society of Arts (RSA) of London, the City and Guilds of London, etc. Conducting examinations on behalf of other examining bodies: These include Test of English as a Foreign Language (TOEFL) administered on behalf of United States Foreign Missions, Embassies and Universities and the Scholastic Aptitude Tests.(SAT) etc, administered on behalf of Educational Testing Service (ETS) of the USA.

Following an increasing agitation for the creation of a National Examination body that will be equivalent to WAEC in functions and to minimise the frustration experienced by candidates due to WAEC monopoly, the National Examination Council, NECO was established. According to Nwana (2007), the National Examinations Council (NECO) bill was passed into law in 2001 and between its creation in 1999 and passing of the bill, it operated under the National Board for Educational Measurement (NBEM) Act 69 of 1993 having its Headquarters in Minna. The establishment of NECO was not to compromise the place of WAEC, rather the efforts and policy
directives that brought it into being ensured that the academic standard of the NECO Senior School Certificate was to be the same as that of WAEC at that level, including such examining features as the style of developing the examination papers, the marking and grading of papers etc. The National Examination Council conducts national examinations such as:
i. Junior School Certificate Examination (JSCE),
ii. Examinations into schools for the gifted children,
iii. Senior School Certificate Examination (SSCE),
iv. National Common Entrance Examination (NCEE).

Generally, stakeholders in education (Government, teachers and parents) have it that NECO and WAEC tests are equivalent. Evidence of this equivalence is shown in the requirement for admission of candidates into tertiary institutions in Nigeria. These higher institutions require that a student must possess a minimum of five credit passes including mathematics and English language in either WAEC or NECO or a combination of the two.

Of the five required credit passes in school subjects, Mathematics is known to be the most fundamental and useful tool in the technological advancement of any nation (Abiodum, 2005). Furthermore, it has been described as the most useful instrument in commerce, physical sciences, engineering, social sciences, industry, medicine and biological sciences (George, 2007). Considering the fact that NECO was established to be equivalent to

WAEC, it is reasonable to expect that the performance of examinees in the test being conducted by the two examining bodies should be comparable.

However, the statistics of students' performance in mathematics tests from year 2005 to year 2014 of the two examining bodies appeared to be at great variance (National Bureau of Statistics, 2015). For example, in year 2008, the statistics show that WAEC recorded $57.24 \%$ passes at credit level (i.e., A1 C6), while NECO recorded more than $70 \%$ credit level passes. In fact since the inception of NECO till date, the results of the two examining bodies have always been characterized by variation in performance levels of candidates in the examinations. One then is left to question the equivalence of NECO and WAEC tests since the disparity between the performances of students in the two tests of the examining bodies are so much, perhaps, the tests are not comparable.

Evaluation of tests comparability, especially tests that are designed to measure the same construct (e.g., students' abilities) are usually achieved by empirically comparing the test scores produced by the two tests. Generically, evaluation of the comparability of test scores on two tests is done through a process called linking (Kolen \& Brennan, 2004). According to Holland (2007), linking refers to the general class of transformations between the scores from one test and those of the other. Holland added that linking transformations can be developed in a variety of methods that reflect the similarities and differences between the tests as well as the uses to which the links are to be put.

Holland and Dorans (2006) divided linking methods into three basic categories: predicting, scale aligning and equating. Among these methods of linking, only the equating method evaluates the equivalence of tests whose test scores are intended to be used interchangeably (Dorans \& Walkers, 2007; Hambleton, Swaminathan \& Rogers, 1991; Holland, 2007; Kolen \& Brennan, 2004).

In order to conclude that two tests are equivalent and that the test scores obtained from the tests can be used interchangeably, five conditions are widely required to be fulfilled by the tests (Holland \& Dorans, 2000; 2006; Petersen, 2007). These conditions include:
i. The equal ability requirement;
ii. The equal reliability requirement;
iii. The symmetry requirement;
iv. The equity requirement; and
v. The population invariance requirement.

However, Petersen (2007) showed that requirements "iv" and "v" respectively explain why both requirements "i" and "ii" should be used in the evaluation of tests equivalence. As a result, this study evaluated the equivalence of WAEC and NECO mathematics tests using requirements i, ii, and iii.

The equal ability condition requires that the tests should measure the same ability of examinees. In the context of this study, this requirement implies that WAEC and NECO mathematics tests should measure the same 'mathematics
ability' of the students. Evaluation of this criterion/requirement can be achieved through several approaches. Prominent among these approaches is the Reduction in Uncertainty index (RiU) (Dorans, 2000). This index is estimated using the correlation coefficient of the test scores obtained by examinees on the two tests. Dorans (2000) state that when the index of RiU estimated for two test scores is greater than $50 \%$, it is concluded that the tests are close enough to measure the same ability.

Test scores obtained by examinees in achievement tests can be estimated under the two major measurement frameworks used in educational testing- the classical test theory (CTT) and item response theory (IRT). The CTT approach estimates observed scores of examinees while IRT approach estimates ability scores of examinees (Hambleton \& Jones, 1993; Kolen \& Brennan, 2004).

In the present study, the IRT approach to test score estimation was adopted. This is because the ability estimates under IRT are independent of the number of items contained in the test that are used to compute their values. This is unlike the CTT observed scores that depend largely on the number of items in the test that are used in the estimation of their values (Hambleton, Swaminathan \& Rojers, 1991). More importantly, the IRT approach to test score estimation has been found to work better than the CTT counterpart (Adegoke, 2014). In using IRT framework, it is important to evaluate which of the IRT models is the most appropriate for the data of interest. Item response theory models have been developed for tests whose items are scored dichotomously (0 or 1) as well as
tests whose items are scored polytomously (e.g., essay tests in which examinees can earn a score of 0,1 , or 2 or more on each item) (Kolen \& Brennan, 2004). Because this study emphasizes multiple choice tests whose items are scored dichotomously, the IRT models for dichotomously scored items was adopted.

According to Kolen and Brennan (2004), item response theory models for dichotomously scored test items assumes that examinee ability is described by a single latent variable referred to as theta $(\theta)$, defined so that theta lies between $\infty$ and $+\infty$. The use of a single latent variable implies that the construct being measured by the test is unidimensional. For test items that are dichotomously scored, there are three parameter models. These are: three-, two- and one parameter logistic models. These models provide mathematical equation for the relation of the probability of an examinee answering an item correctly as a function of the item's parameter(s) and the underlying ability that the test intends to measure (Baker, 2001).

The use of the models is governed by two basic assumptions which must be fulfilled for accurate item and person parameter estimates to be obtained. These assumptions include; unidimensionality, and local independence (Demars, 2010). Furthermore the choice of which of the models that is the most appropriate for item and person estimates depends largely on the model-data fit (Demars, 2010).

Unidimensionality means that the model has a single $\theta$ for each examinee, and any other factors affecting the item response are treated as random error or
nuisance dimensions unique to the item and not shared by other items. According to Hambleton et. al (1991), what is required for this assumption to be met adequately by a set of test data is the presence of a "dominant" component or factor that influences test performance. Many methods have been proposed for testing unidimensionality. Prominent among the methods is analysis of the eigenvalues (roots) of the inter-item correlation matrix (Demars, 2010). According to Lord (1980), when the eigenvalue of the first factor is substantially greater than the second then the data can be assumed to be approximately unidimensional. Specifically a test is adjudged unidimensional when the ratio of the eigenvalue of the first component to the second component is probably 2 to 1 (2:1) or more. However, the extent of unidimensionality of a test increases as the ratio of the eigenvalue of the first component increases from 2 . This implies that a test with eigenvalues whose ratio of the first component to the second component is 3 to 1 (3:1) is more unidimonsional when compared with a test with eigenvalue whose ratio of the first component to the second component is 2 to 1 (2:1) (Demars, 2010).

On the other hand, item local independence states that the probability of an examinee answering a test item correctly is not affected by his/her performance on any other item in the test (Nenty, 2004). Assessment of item local independence is usually achieved by outright tetrachoric/polychoric correlation among items response on a test (Ubi, 2006 cited in Ubi, Joshua, \& Umoinyang, 2011). According to Ubi et al (2011), an item is considered locally
independent, if the tetrachoric/polychoric correlations among the items are not significantly different from zero.

Model- data fit issues are of major concern when applying item response theory models to real data (Steven, 1990). In fact when a model is not appropriate or does not fit the data, use of estimated parameters may be compromised (Stone \& Zhang, 2003). Typically, IRT practitioners focus on the fit of individual items, not on overall fit of the model across all items. This is usually assessed by using all the models for calibrating the data of interest. The model that best fits the data is obtained by the count of the number of items deleted by each model. The model that deletes the smallest number of items is usually adjudged the model that fits the data (Wiberg, 2004). As noted by Demars (2010), violation of these assumptions may lead to misestimation of parameters (items and examinees ability).

It is reasonable to adjudge two tests that are able to fulfil the equal ability requirement as being equivalent. However, researchers (Dorans \& Holland, 2000; Petersen, 2007) showed that tests that measure the same construct but differ in reliability should not be considered equivalent. This implies that the equal reliability requirement is equally important in evaluating the equivalence of two tests.

Reliability, the extent to which an instrument measures consistently what it is designed to measure is important for two reasons. These reasons are; (1), reliability provides a measure of the extent to which an examinee's test score
reflects random measurement error and (2), reliability is a precursor to test validity. That is, if test scores cannot be assigned consistently, it is impossible to conclude that the test scores accurately measure the domain of interest (Wells \& Wollack, 2003). Thus, reliability is a central notion in educational measurement and classical test theory. Within item response theory, test reliability is estimated using the information function of the individual item that makes up the test (Baker, 2001). This reliability in the parlance of IRT is usually estimated as empirical reliability by IRT softwares (e.g BILOG MG) ( De Ayala, 2009).

Evaluation of reliability of two tests for the purpose of equivalence assessment requires that the two tests should have equal reliability estimates. However, LiU and Walker (2007), stated that equality of reliability is a necessary but not sufficient condition for tests equivalence. In addition, high reliability is needed to ensure that the equivalent test scores are informative enough to be acceptable (Dorans, 2004). According to Cohen and Swerdlik (2009); Kline (2000) and Kline (2005), this value of reliability should not drop below 0.7.

The third requirement that must be fulfilled by two tests in addition to equal ability and equal reliability requirements is the symmetry requirement. Specifically, symmetry condition requires that the transformation used in equating tests must be symmetric (Lord, 1980). This implies that, the function used to transform test scores on test 1 to the scale of test 2 must be the inverse of the function used to transform test scores on test 2 to the scale of test 1 .

Evaluation of this requirement is determined using scatterplots of the equating functions. According to Kolen and Brennan (2004; 2014), the functions are symmetrical when for example, an ability score of +2 on test 1 converts to ability score of +3 on test 2 scale, then an ability score of +3 on test 2 must convert to ability score of +2 on test 1 scale.

Transformation of tests' scores of one form to scale of another form can be achieved through tests equating procedures of the classical test theory (CTT) and item response theory (IRT). The CTT approach makes use of observed scores of the testees while the IRT makes use of the ability estimates of the examinees (Kolen \& Brennan, 2004). However, in the present study the IRT approach was adopted. This is because the ability estimates under IRT are independent of the samples (number of items) that are used to compute their values. More importantly, the IRT approach to test equating has been found to work better than the CTT counterpart (Fan, 1998; Harmbleton, Swaminathan \& Rojers, 1991).

Under IRT, four methods are used in the transformation of ability estimates on one test form to the scale of another test form (Kolen \& Brennanm, 2004). These methods of transformation include: mean/sigma, mean/mean, Haebara Test Characteristics Curve, and Stocking and Lord Test Characteristics Curve. The choice of any of the transformation methods for test score transformation is often dictated by the data collection design used for the test forms to be transformed. Majorly, there are three data collection designs used in

IRT equating; they are (a) single group, (b) random group and (c) anchor test designs (Kolen \& Brennan, 2004). In the single group design, one group of testees taken from one population takes the test forms to be equated. In the random group design, two randomly selected groups, of equivalent ability taken from the same population, take different forms of the test. In the anchor test design two groups of examinees taken from two different populations take different forms of the test, with each form containing a set of common items.

According to Kolen and Brennan (2004), the mean/sigma method of transformation is used when the data collection design employed for test equating is the single group design while the mean/mean or mean/sigma method is used when the random group design is used for data collection. But any of the mean/mean, mean/sigma, Haebara test characteristics curve, stocking and Lord test characteristics curve methods can be used when the anchor test design is employed for data collection. Furthermore, the choice of design-single group, random group, or anchor test- is often dictated by practical constraints of the testing program (Cook \& Eignor, 1991).

In the present study which emphasized NECO and WAEC tests, the single group design was adopted. This is because of the need to ensure that variation in test scores, if any, is not a function of heterogeneity of samples. Consequently, the mean/sigma transformation method was adopted in the study.

Mean/sigma transformation method under the single group makes use of the means and standard deviations of the ability estimates of examinees on test
forms to transform the examinees' ability estimates on test forms from the scale of one form to the other form. According to Kolen and Brennan (2004), examinees' ability estimates on one test form can be placed on the scale of another test form by equating the standardized ability estimates of examinees on the two test forms.

Quite a number of empirical enquiries in Nigeria have been advanced to assess the extent of comparability of WAEC and NECO tests. For example, Kolawole (2007) compared the psychometric properties of NECO and WAEC mathematics multiple choice tests and found that the reliability and validity coefficients of the two tests were similar. Other empirical studies (Bamidele and Adewale, (2013) and Obinne, (2008) laid credence to the finding of Kolawole (2007). Although Anagbogu (2009), found that NECO items were more difficult than WAEC items, this result was based on the discrimination parameters of the two tests used in the empirical study.

Based on available empirical studies, it appears that the focus of research on establishing the comparability of WAEC and NECO at least in Nigeria have always been centred on only the psychometric properties of the tests of the examining bodies. To a large extent, a lot of features regarding the extent of the equivalence of WAEC and NECO mathematics tests are empirically unknown. Specifically, there is no known study in Nigeria that has assessed the equivalence of WAEC and NECO mathematics tests in terms of the traits which the tests purport to measure; the equivalence of the reliability of the tests, the
equivalence of the local independence of the tests' items. More importantly, the application of item response theory in assessing the equating of the test scores obtained from the two tests is indeed rare. Furthermore, whether or not scores obtained from the two tests could be used interchangeably was unknown.

## Statement of the Problem

The West African Examination Council (WAEC), in the early days of its inception, was the only examining body in Nigeria saddled with the responsibility of measuring, assessing and certifying senior secondary students' proficiency in school subjects.

Nwana (2007) noted that the establishment of National Examination Council (NECO) was not to compromise the place of WAEC, the efforts and policy directives that brought it into being ensured that the academic standard of the NECO Senior School Certificate was to be the same as that of WAEC at that level, including such examining features as the style of developing the examination papers, marking and grading of papers, etcetera. Ever since then tertiary institutions in Nigeria require for admission minimum of five credit passes which include mathematics and English language in either WAEC or NECO or a combination of the two. This suggests that the WAEC and NECO tests (especially mathematics) are equivalent and the test scores can be used interchangeably.

Although stake holders in education believe that tests of the two examining bodies are equivalent, the statistics of results of students that sat for the examinations show different picture as a lot of disparities were observed between levels of students' performance in the results of the two examining bodies' examinations. However, to a large extent, there is no empirical justification for this claim of equivalence by stake holders as there is a dearth of literature on the extent of equivalence of WAEC and NECO tests. The few studies done on the comparison of WAEC and NECO mathematics tests and other senior secondary certificate examination subjects limited their enquiries to reliability estimates of the two tests with no attention given to the extent of equivalence of the two tests and the extent to which the tests can be used interchangeably. Furthermore, test conducted by WAEC is often perceived by the public as being superior to those conducted by NECO because of its international outlook and years of operation. However, this assertion is yet to be proved empirically. Hence, the extent of equivalence of tests of NECO and WAEC appears to be largely unknown. There is need for more in-depth studies.

This study therefore, assessed the equivalence of WAEC and NECO mathematics tests using the tests interchangeability criteria of test equating in the measurement framework of item response theory.

## Purpose of the Study

The main purpose of the study was to assess the equivalence of mathematics multiple choice tests of NECO and WAEC from 2011 to 2014; ascertain if the test scores obtained on the tests could be used interchangeably using the criteria for assessing tests score interchangeability and finally ascertain the unidimensionility and local independence of the tests.

Specifically, the study set out to:

1. assess the unidimensionality of NECO and WAEC mathematics multiple choice tests from 2011-2104
2. assess the local independence of NECO and WAEC mathematics multiple choice test items from 2011-2014
3. compare the local independence of NECO and WAEC mathematics multiple choice test items from 2011 - 2014
4. ascertain if the NECO and WAEC mathematics multiple choice tests of 2011-2014 measured the same mathematics ability of examinees.
5. estimate the reliability estimates of NECO and WAEC mathematics multiple choice tests from 2011-2014
6. assess how symmetrical the equating functions for placing the ability estimates of examinees in WAEC test on the scale of NECO test are and vice versa.

## Significance of the study

This study which was designed to assess the equivalence of mathematics multiple choice tests of WAEC and NECO, is significant in many ways to examining bodies, parents, Ministry of Education and stakeholders.

The result would help examining bodies to appreciate the usefulness of test equating in determining the extent of equivalence of achievement tests developed from the same content area and intended for the measurement of the same construct.

The findings of the study would help to disabuse the minds of parents that the results obtained by their wards on WAEC and NECO mathematics tests are equivalent, despite the order/claim by government that they are equivalent. The findings of the study would inform the Federal Ministry of Education, on the need to establish a central body that will be saddled with the responsibility of ensuring that WAEC and NECO tests are as equivalent as possible. This body will serve as monitoring team for the maintenance of the expected standards in setting of the examination, conduct of examinations and certification.

More importantly, the findings of this study would provide information to stakeholders regarding, the equivalence of WAEC and NECO mathematics tests and whether the test scores of mathematics tests of the two bodies can be used interchangeably.

The study would also be a very useful reference material to future research students who may wish to investigate this issue further, as well as those who may which to carry out other comparative studies.

## Scope of the Study

The item statistics of any test item are many but the researcher intends to focus on the unidimensionality of the test item, local independence of test items, test score equating of the examination papers. The study restricted it's scope to the Senior Certificate Examinations that were conducted by two examination bodies (WAEC and NECO) from 2011 to 2014.

## Research Questions

The following research questions were answered in the course of this study.

1. How unidimensional are the SSCE mathematics multiple choice test items of NECO and WAEC?
2. How locally independent are SSCE mathematics multiple choice test items of NECO and WAEC?
3. How equivalent are the local independence of SSCE mathematics multiple choice test items of NECO and WAEC?
4. To what extent do SSCE mathematics multiple choice tests of WAEC and NECO measure the same construct (i.e., mathematics proficiency)?
5. How reliable are NECO and WAEC mathematics multiple choice tests?
6. Are the equating functions for placing the ability estimates of examinees in mathematics test of WAEC on the scale of mathematics test of NECO and vice versa symmetrical?

## HYPOTHESES

The following null hypotheses were tested at 0.05 level of significance

1. There is no significant relationship between the unidimensionality of NECO and WAEC Mathematics multiple- choice test
2. There is no significant difference between the reliability estimates of NECO and WAEC Mathematics multiple -choice test.

## CHAPTER TWO REVIEW OF RELATED LITERATURE

Literature related to this study was reviewed under the following headings:-

## Conceptual Framework

Unidimensionality of test items

Local independence and test items
Reliability of a test
Test score equating

## Theoretical Framework

Item response theory (IRT)

## Theoretical Studies

IRT models

Applications of IRT

Model-data fit

Different forms of estimating reliability

Test equating requirements

## Empirical Studies

Unidimensionality of tests
Local independence of test items
Reliability estimates of tests

Test equating
Summary of Review of Related Literature

## Conceptual Framework

Unidimensionality of test items. Item response theory (IRT) assumes that there is a single latent trait which is sufficient to explain or account for examinee's performance, referred to as one dimensional. One dimensional test may be defined simply as a test in which all items are measuring the same trait. Unidimensionality, therefore means that any item developed should test one area of knowledge and nothing else (Joshua, 2005). This in actual sense implies that there are divergence of test items, which should all converge to measure a particular area of knowledge. For example, a mathematics test that contains some items that are strictly computational and other items that involve verbal material likely is not unidimensional.

According to Sitjsma and Bram (2006), unidimensionality assumes that a single latent ability or a homogenous set of items to be tested is sufficient to explain examinee performance. This means that the items that collectively measure a unique underlying latent trait for each examinee and that only one latent trait influences the item responses. Other factors affecting these responses are treated as random error (Demars, 2010). This assumption relates to construct validity, since the test must only measure that particular construct, it is designed to measure. For instance, a test designed to measure depression must only measure that particular construct, not closely related ideas such as anxiety or stress. Nworgu (2015) stated that construct validity is concerned with the extent to which research instrument measures one particular psychological (social)
construct. This is a strong assumption and may not be reasonable in many situations as tests or survey instruments may be designed to measure multiple traits.

According to Hambleton et al (1991), what is required for this assumption to be met adequately by a set of test data is the presence of a "dominant" component or factor that influences test performance. Many methods have been proposed for testing unidimensionality. Prominent among the methods is analysis of the eigenvalues (roots) of the inter-item correlation matrix (Demars, 2010). According to Lord (1980), when the eigenvalue of the first factor is substantially greater than the second then the data can be assumed to be approximately unidimensional.

Local independence of test item. Item local independence states that the probability of an examinee answering a test item correctly is not affected by his/her performance on any other item in the test (Nenty, 2004). In other words, the knowledge of item 1 should not aid an examine to answer item 2. However this assumption has never been met precisely, for this reason, Joshua (2005) suggested that chain items should be avoided.

In applying the IRT models, the local independence assumption is one of the important features. This assumption means that for every examinee's response $\mathrm{Y}_{\mathrm{pi}}=0$ or 1 (where 0 denotes an incorrect response and I denotes a correct response) to the items I are statistically independent. In other words, the examinee's performance on one item is not influenced by the correctness of
answering the other items (Sitjsma \& Brain, 2006). The local independence assumption implies that there are no dependencies among items other than those that are attributable to latent ability. One example where local independence likely would not hold is when tests are composed of sets of items that are based on common stimuli, such as reading passages or chapters. In this case, local independence probably would be violated because items associated with one stimulus are likely to be more related to one another than to items associated with another stimulus. (Kolen \& Brennan, 2004).

Local independence, also indicates that if the assumption of unidmensinality holds, then the response of a subject to one item will be independent of his or her response to another item, conditional on the latent trait. In other words, if items are locally independent, they will be uncorrelated after conditioning on $\theta_{\mathrm{j}}$ (Demars, 2010). Assessment of item local independence is usually achieved by outright tetrachoric/polychoric correlations among item responses on a test (Ubi, 2006 cited in Ubi et al. 2011). According to Lord (1978) in Ubi et al (2011), an item is considered locally independent, if the tetrachoric/polychoric correlations among the items are not significantly different from zero.

Reliability of a Test. Reliability is an essential quality in any kind of measurement. It is the degree of consistency with which it measures whatever it is measuring. According to Okoye (2015), the reliability of an instrument is a measure of the extent to which the instrument consistently measures what it intends
to measure. Cozby and Bates (2012) defined reliability as the consistency or stability of a measure of behavior. Abonyi (2011) viewed reliability as the extent to which we can attribute individual differences in test scores to true difference. Test reliability can also be defined as the degree to which test items consistently measure any phenomenon they measure and the degree to which the same responses repeatedly given to the same questions repeatedly attract the same score (Nkwocha, 2007). This implies that it is repeatability or replicability of measurement.

Test Score Equating. Test equating is a statistical procedure, used to ensure comparability of test forms built with the same content area for the measurement of the same construct. According to Michealides and Haertel (2004), test equating is the statistical process that establishes comparability between alternate forms of test built to the same content and statistical specifications by placing scores on a common scale; thus allowing interchangeable use of scores on these forms .

Equating could be described as the relationship between scores of different forms that are constructed according to the same content and statistical specifications (Kolen \& Brennan, 2004). Equating adjusts for difference in difficulty among forms that are built to be similar in difficulty and content.

## Theoretical Framework

## Item response theory

According to Baker (2014), item response theory is the study of test and item scores based on the assumptions concerning the mathematical relationship between abilities and item responses. This implies that the item response seeks to establish a relationship between the characteristics of the examinees, the individual test item parameters and the response to each item. It holds that there some latent traits in examinees that are not directly measurable or observable but are assumed to underlie a test performance. We can therefore say that the IRT states that the probability that an examinee with a given latent trait $(\theta)$ gets an item right is dependent on the examinee's ability and item characteristics (parameters).

Item response theory is not single theory postulated by a single individual. It evolved as a result of the works of a number of individuals. According to Ainsworth, (2016), IRT refers to a family of latent trait models used to established psychometric properties of items and scales. Sometimes it is referred to as modern test theory to differentiate it from classical test theory (CTT).

Paul Lazardsfeld, a mathematical sociologist was the first to initiate and introduce the term "Latent - Trait" in measurement theory in 1950. Lazarsfeld developed a set of mathematical models closely akin to factor analysis models. His work included both models that could be used with continuous and categorical variables, as in the case of factor analysis. Lazarsfeld's models
attempt to explain relationships among observed variables in terms of one or more unobserved (Latent) variables which is referred to as traits, while the developmental level score is latent - trait score, which is a special kind of factor score. Lazarsfeld called both the categorical and continuous variable models latent - trait models.

The work of Frederic Lord in 1952 gave birth to Latent - trait theory. Lord in various ways sustained the activities of IRT since he was interested in the relationship between an individual's response to test items and characteristics of the items. Then, until when Richardson (1956) derived relationship between item response theory model parameters and classical item parameters, which provided an initial way for obtaining item response theory parameter estimate. Subsequently, Birnbaum (1965), Rasch (1960), Wright (1967), Lord and Norvick (1968), Wight and Panchapa, Kesan (1974), Akpan (1989), in one way or the other contributed to the development and popularity of items response theory. These experts contributed in their own ways for the development and sustenance of item response theory.

Thus, IRT is the most significant and popular development in psychometrics to overcome the shortcoming of classical test theory and maximize objectivity in measurement (Joshua, 2005). Anderson \& Morgan (2008) submits that IRT allows an item to be characterized independently of any sample of items administered to the person. Thus, IRT is very useful when multiple set of items are administered to students in an assessment. It is against
this background that IRT is the theoretical framework on which the study was anchored. Therefore, the present study relates to this theory because it aims at assess the equivalence of WAEC and NECO mathematics multiple-choice tests using item response theory (IRT) from 2011 to 2014. Also assess the extent to which the test scores obtained on the tests can be used interchangeably using the criteria for assessing tests unidimensionability and local independence.

## Theoretical Studies

IRT models. IRT models are mathematical functions that specify the probability of a discrete outcome, such as a correct response to an item, in terms of person and item parameter (Zairul \& Adibah, 2010). IRT models show the relationship between the ability or trait (symbolized $\theta$ ) measured by the instrument and an item response. The item response may be dichotomous (two categories), such as right or wrong, yes or no, agree or disagree. Or, it may be polytomous (more than two categories), such as a rating from a judge or scorer or a likert - types response scale on a survey. The construct measured by the items may be on academic proficiency or aptitude, or it may be on attitude or belief (Demars, 2010). IRT models are commonly used to model the latent traits associated with a set of items or questions in a test or survey.

IRT models describe the interactions of person and test items (Reckase, 2009). Therefore, to describe a relationship between the examinees' ability and performance on an item, one employs one or more parameters depending on the

IRT model appropriate in use. The models are referred to as logistic parameter models. These models are one-parameter logistic model (1PL), two-parameter logistic model (2PL) and three- parameter logistic model (3 PL) (Baker, 2001).

One-parameter Logistic model. This model is also called the Rash model, which was a major focus of the Danish mathematician, George Rasch. The Rash model involves a rigorous perspective that is distinguished not only by its adherence to the IRT assumptions, but also by its emphasis on high quality items and internal scales (Bond \& Fox, 2007; Embretson \& Reise, 2000). The Rash model estimates the probability of answering the item correctly as a logistic function of the difference between the individual's ability and the item difficulty. Thus, the Rash model allows us to create an internal scale of scores for both the item's difficulty and individual's ability, and these scores are scaled in logits (Baker, 2004). The Rash model, or the one parameter logistic (IPL) model, specifies that the item difficulty is the only item characteristic that varies from item to item holding the item discrimination values equal for all items. Because there is only one parameter to be estimated, this model does not require large sample size. Previous research suggested that a sample size of as large as 200 examinees would be sufficient to accurately estimate item parameters of the IPL model (Chang, Hanson \& and Harris, 2000).

One- parameter logistic model (1PL)
$P(\theta)=\frac{1}{1+e^{-1\left(\theta-b_{i}\right)}}$
eqn 1.1

Where $\mathrm{P}(\theta)$ is the probability of correct response to an item,
$\mathrm{e}=2.72828 \ldots,($ exponential constant $)$
$\mathrm{b} i=$ difficulty parameter of item $i$ known under 1RT as item location parameter

The Rash analysis is centered on estimating the probability of success of the testtaker on a specific item. The use of Rash analysis contributes to the normalization of test curves. Just as CTT has four indicators to assess the adequacy of the items, Rash models have statistics to evaluate the fit of the item into the model (Limacre, 2008). The idea underlying these statistics is that correct answers in more difficult item are accompanied by people with higher ability. At the same time, these people will have a greater probability of attaining higher scores on easier items than on more difficult ones.

Two parameter logistic model:- Orluene (2010) records that this model proposed by Birnabaum (1968) contains two item parameter (difficulty and discrimination indices) and latent ability omitting the guessing parameter, which is designated as "a" for discrimination and "b" for difficulty as opposed to one parameter model which was limited to difficulty "b". The function of the discrimination parameter is that, it allows the item characteristics curve (ICC) for various items to exhibit different slopes and further confirms that some items
have a stronger or weaker relationship with the underlying construct $(\Theta)$ being measured than others. Due to the item characteristic curve, the slope of the curve is a function of the ability level and reaches the maximum value difficulty (Baker, 2001).

The equation for the two - parameter logistic model is given by:

Two-parameter logistic model (2PL)

$$
P(\theta)=\frac{1}{1+e^{-a_{i}\left(\theta-b_{i}\right)}}
$$

Where $\mathrm{P}(\theta)$ is the probability of correct response to an item, $\mathrm{e}=2.72828 \ldots,($ exponential constant $)$
$\mathrm{a}_{\mathrm{i}}=$ discrimination parameter of item $i$, conceptually known under 1RT as item slope
$\mathrm{b} i=$ difficulty parameter of item $i$ known under 1RT as item location parameter, and

Three parameter logistic model: This model adds a guessing parameter as the third descriptor of the item characteristic curve. When the guessing parameter is taken into account, it shows that in many items, even if the examinee does not know anything about the subject matter $(\Theta=-5)$, he or she can still have some chance ( $\mathrm{p}>\mathrm{o}$ ) to get the right answer ( $\mathrm{Yu}, 2013$ ). The 3- PL model is a more general model where the discriminating power is allowed to vary among items and guessing is allowed to occur for the examinees. However, in order to accurately estimate the 3 pl item parameters, previous research
suggested that at least 1,000 (Darmas, 2010; Reckase, 1979, Skaggs \& Lissitz, 1986) to 10,000 (Thissen \& wainer, 1982) examinees would be needed.

Estimating IRT item discriminating parameter requires larger sample sizes than estimating item difficulty parameters (Barnes \& Wise, 1991)

The formula is given by

P Three-parameter logistic model (3PL)
$P(\theta)=c_{i}+\left(1-c_{i}\right) \frac{1}{1+e^{-a_{i}\left(\theta-b_{i}\right)}} \quad--\cdots-\cdots \cdots-\cdots-------\quad$ eqn1.3.
Where $\mathrm{P}(\theta)$ is the probability of correct response to an item,
$\mathrm{e}=2.72828 \ldots,($ exponential constant $)$
$\mathrm{a}_{\mathrm{i}}=$ discrimination parameter of item $i$, conceptually known under 1RT as item slope
$\mathrm{b} i=$ difficulty parameter of item $i$ known under 1RT as item location parameter, and
$\mathrm{C} i=$ guessing parameter of the item $i$,

## Application of Item Response Theory.

Umobong (2004) outlined the following as the applications of IRT in testing situations.

Item Banking:- Involves collection of test items, "stored" with known item characteristics and made available to test constructors. The invariance property of IRT parameter makes it possible to feed items into item banks, and recall
them at will. With the objective nature of IRT, items are constructed, administered and tested for goodness of fit, and items found not to fit the model are eliminated while those that fit the model are stored.

Tailored Testing:- is the matching of the test item to the ability of the examinee, in other words, attempting to save the high ability examinee from too many trivially easy items and low ability examinee from too many hard items. IRT is important in tailored testing because it makes for the estimation of ability that are independent of the particular test items administered as such examinees can be compared even when they had taken different sets of test items.

Test Construction / Development. Test construction and development under IRT is more advantageous particularly in the aspects of item analysis, selection of items, test validity and reliability assessment. Items in IRT are said to be calibrated without reference to the items by a test of fit of the model.

IRT has been applied to the types of tests encountered in the areas of educational and psychological measurement. In developing tests using latent trait theory, it provides the test developer with sample invariant item parameters and a powerful method of item selection which depends on the test information function.

Test Bias:- Item bias is defined as a test item in which all individuals having same underling ability have equal probability of getting the item correct, regardless of subgroup membership. Items in a test that measure a single trait must measure the same trait in all subgroups of the population to which the test
is administered. Items that fail to do so biased for or against a particular subgroup (Lord \& Novick, 1991).

Criterion References Testing (Mastery Testing):- Criterion referenced tests are designed to determine the extent to which an examinee has reached a specific level of achievement, in which case, the examinee is said to be proficient or "master" of a particular skill tested. Particularly, a criterion - referenced test is based on a well defined set of tasks known as a universe of items or a domain.

In evaluating instruction, test developers are concerned with the proportion of tasks in the domain that a student can correctly perform. This proportion of task that a student can correctly master is defined as a student domain score or mastery level. The test developer then selects items from the pool that have the highest item information functions at that level.

This type of test measures precisely at the mastery level such that errors in classifying people are minimized. The advantages of application of IRT over CRT is that due to the invariance property of the parameter, it is possible to construct test tailored to meet specific ability level.

Test Score Equating. Test equating is a statistical procedure, used to ensure comparability of test forms built with the same content area for the measurement of the same construct. According to Michealides and Haertel (2004), test equating is the statistical process that establishes comparability between alternate forms of test built to the same content and statistical specifications by placing scores on a common scale; thus allowing interchangeable use of scores on these forms.

In large scale testing situation multiple and interchangeable forms of test are used. Classical test construction techniques do not assure that two or more forms of a test can be made equivalent in level and range of difficulty. At times the need for equating test from either horizontally or vertically using the same unit for the two forms arises. Hambleton et al (1978) opined that latent trait models allow equating scores using ability estimates, which helps to control for the problem of non - equivalent groups and that of constructing parallel forms of tests. This is due to its characteristics of parameter variance.

## Model-data fit

In using IRT framework, it is important to evaluate which of the models is the most appropriate for the data of interest. The choice of which of the models that is the most appropriate for item and person estimates depends largely on the model-data fit (Demars, 2010). Model- data fit issues are of major concern when applying item response theory models to real data (Steven, 1990). In fact when a model is not appropriate or does not fit the data, use of estimated parameters may be compromised (Stone \& Zhang, 2003). The above statement implies that the use of a model that does not fit the data well cannot provide good answer to the underlying scientific questions under investigations.

Typically, IRT practitioners focus on the fit of individual items, not on overall fit of the model across all items. This is usually assessed by using all the models for calibrating the data of interest. The model that best fits the data is
obtained by the count of the number of items deleted by each model. The model that deletes the smallest number of items is usually adjudged the model that fits the data (Wiberg, 2004). As noted by Demars (2010), violation of these assumptions may lead to misestimation of parameters (items and examinees ability).

In assessing model-data fit, the best approach involves (i) designing and conducting a variety of analysis designed to detect expected types of misfit, (ii) considering the full set of results careful and (iii) making a judgment about the suitability of the model for the intended application. (Hambleton, 1991).

Different forms of estimating reliability: A reliability measure does not fluctuate from one reading to the next, if the measure does fluctuate, there is error in the measurement device. Therefore, a more formal way of understanding reliability is to use the concepts of true score and measurement error. Any measure that you make can be thought of as comprising two components: 1) a true score, which is the real score on the variable 2) measurement error. An unreliable measure of intelligence contains considerable measurement error and so does not provide an accurate indication of an individual's true intelligence (Cozby \& Bates, 2012).

The question is, how can we assess reliability? We cannot directly observe the true score and error components of an actual score on the measure. However, one can assess the stability of measures using correlation coefficients.

According to Anastasi (1988) correlation coefficient (r) expresses the degree of correspondence, or relationship, between two sets of scores. It is also described as a number that tells us how strongly two variables are related to each other (Cozby \& Bates, 2012). Correlation coefficients have many uses in the analysis of psychometric data. The measurement of test reliability represents in application of such coefficients. Therefore the most common correlation coefficient when discussing reliability is the Pearson product-moment correlation coefficient. The Pearson correlation coefficient (symbolized as r ) can range from 0.00 to +1.00 and $0-00$ to -1.00 . A correlation of 0.00 tells us that the two variables are not related at all. The closer a correlation is to 1.00 , either +1.00 or -1.00 , the stronger is the relationship. The positive and negative signs provide information about the direction of the relationship. When the correlation coefficient is positive; there is a positive linear relationship - high scores on one variable are associated with high scores on the second variable. A negative linear relationship is indicated by a minus sign - high scores on one variable are associated with low scores on the second variable.

Therefore to assess the reliability of a measure, one needs to obtain at least two scores on the measure from many individuals. If the measure is reliable, the two scores should be very similar. The index of reliability is usually the reliability coefficient. Hence, reliability coefficient is the statistic or an index that tells us the degree of consistency between two sets of
scores obtained from the same group with one test (Onunkwo, 2002). There are various methods of establishing the reliability of a test. These include the following; test - retest, equivalent forms method, split half, Kuder Richardson and Cronbach alpha coefficients.

Test-retest Method. The test - retest method involves administration of one test twice on the same group of people, at a reasonable interval of time. The time interval could be short or long, depending on the nature of the attribute being measured. The scores they obtain at the two different occasions are correlated using an appropriate type of correlation. Pearson's product moment correlation is used for interval scores, while Spearman's Rank Correlation is used for ordinal data. A reliability coefficient estimated through the process of test retest is regarded as a measure of stability.

Equivalent or Parallel Form or Forms Method. Here, two forms of a test are administered to the same group of students in close succession, and the resulting test scores are correlated. This correlation coefficient provides a measure of equivalence. This implies the degree to which both forms of the test are measuring the same aspects of behavior. A test developer or expert may have good reason to construct not just one test but two tests, one test being similar or equivalent or identical to the other in nearly all respects. Such tests according to Nwana (2007) are referred to as parallel tests. Therefore, they will be regarded as parallel if:

1. they both cover the same content and construct
2. each item in one test has its match in the other test in terms of difficulty and discrimination indices
3. they have the same rubric for administering, scoring, grading and scaling the answers.

Many standardized tests have more than one form. These forms of the same tests are said to be equivalent, parallel, or alternate forms. All forms of the same tests are said to be equivalent or parallel; that is, when they yield comparable raw scores from the same group.

Split half Method. This involves a single administration of one form of a test, where the test is administered to a group of subjects and later the items are divided into two comparable halves. Scores are obtained for each individual on the comparable halves and a coefficient of correlation is calculated for the two sets of scores. If each subject has a very similar position on the two sections, the test has high reliability. If there is little consistency in positions, the reliability is low. Coefficient obtained through this process is regarded as a measure of internal consistency.

However, this method of establishing reliability has a limitation and that is, the reliability coefficient so obtained is that of a half test because the full test is split into two. Therefore, to raise it to that of a full test the Spearman Brown Prophecy formula is applied. The formula is meant to raise the reliability coefficient of a half test to that of a full length test. The formula is given below

Reliability of full test $=\quad \underline{2} \times$ reliability of $1 / 2$ test

$$
1+\text { reliability of } 1 / 2 \text { test }
$$

Kuder - Richardson ( $\boldsymbol{K}-\boldsymbol{R}$ ) Reliability Method. This involves using a single administration of a single form, which is based on the consistency of responses to all items in the test. Kuder - Richardson procedures stress the equivalence of all the items in a test, which are appropriate when the intention of the test is to measure a single trait. For a test designed to measure several traits, the Kuder-Richardson reliability estimate will usually be lower than reliability estimates based on a correlational procedure.

The two procedures for determining K-R reliability indices are $\mathrm{K}-\mathrm{R}_{2} \mathrm{O}$ and K $\mathrm{R}_{21}$.
$\mathrm{K}-\mathrm{R}_{20}$ is applicable to tests whose items are scored dichotomously, that is, as either right or wrong; pass or fail. This technique does not required two halves of the test, rather it involves a thorough examination of testee's response or performance on each item of the test. This is applicable in objective test and the formula is

$$
\mathrm{r}_{\mathrm{xx}} \quad \frac{\mathrm{~K}}{\mathrm{~K}-1}\left[\frac{\mathrm{~S}_{\mathrm{x}}^{2}-\Sigma_{\mathrm{pq}}}{\mathrm{~S}_{\mathrm{x}}^{2}}\right]
$$

where
$\mathrm{K}=$ number of items on the test
$\mathrm{S}_{\mathrm{x}}^{2}=\quad$ variance of scores on the total test
$\mathrm{P}=$ proportion of correct responses on a single item
$\mathrm{q}=$ proportion of incorrect responses on the same item
$\mathrm{K}-\mathrm{R}_{2} \mathrm{i}$ can be used in both essay and objective tests. It is computationally simpler but requires the assumption that all items in the test are of equal difficulty, which is practically unattainable in test development. Although it is simple in operation, its use is bound to carry along a number of errors, which leads to the reduction of the reliability estimate. It is given by

$$
\mathrm{r}_{\mathrm{xx}}=\frac{\mathrm{S}_{\mathrm{x}}^{2}-\overline{\mathrm{X}}(\overline{\mathrm{~K}-\mathrm{X})}}{\mathrm{S}_{2}^{2}(\mathrm{~K}-1)}
$$

where
$\mathrm{r}_{\mathrm{xx}}=$ the reliability of the whole test
$\mathrm{K}=$ the number of items in the test
$\mathrm{S}_{\mathrm{x}}^{2}=$ the variance of the scores
$\overline{\mathrm{X}}=$ the mean of the scores

Cronbach's Coefficient Alpha. This is applied in estimating internal consistency of tests that are not dichotomously scored. Conbach alpha is used when measures have multiple scored items, such as attitude scales or essay tests. For example, on a Likert attitude scale the individual may receive a score from 1 to 5 depending on which option was chosen. Similarly, on essay tests a different number of points may be assigned to each answer. Cronbach's alpha provides us with the average of all possible split-half
reliability coefficients. To actually perform the calculation, scores on each item are correlated with scores on every other item.

Abonyi (2011), noted that Cronbach procedure is allergic to the following:

- instruments that are dichotomously scored
- instruments that are not balanced (i.e. number of positively directed and negatively directed items)
- poor representation of construct

The formula for alpha is given by:

$$
\propto \text { or } \mathrm{r}_{\mathrm{xx}}=\left[\frac{\mathrm{K}}{\mathrm{~K}-1}\right] \quad\left[\frac{\mathrm{s}_{\mathrm{x}}^{2}-\sum \mathrm{si}^{2}}{\mathrm{~S}_{\mathrm{x}}^{2}}\right]
$$

Where
$\mathrm{K}=$ number of items on the test
$\sum \mathrm{si}^{2}=\quad$ sum of the variance of the item score
$S_{\mathrm{x}}^{2}=\quad$ variance of the test scores (all K items)
Inter-rater/Scorer/Marker's/Reader Reliability. This approach to estimating test reliability is applicable, especially for determining the reliability of scorers of essay tests. It requires several readers (examiners) who will assign marks to the same essays on agreed basis (using the marking scheme) and their scores are correlated to determine the reader reliability (Anikweze, 2010). Iwuji (1997) sees it as the degree of consistency with which a marker marks an essay script at different occasions.

It is also the level of agreement among group of markers over awards of marks in a given essay test script on different occasions. This type of reliability is assessed by having two or more independent judges score the test. The scores are then compared to determine the consistency of the raters' estimates. One way to test inter-rater reliability is to have each rater assign each test item a score. For example, each rater might score items on a scale from 1 to 10 . Next, .you would calculate the correlation between the two ratings to determine the level of inter-rater reliability. Another means of testing inter-rater reliability is to have raters determine which category each observation falls into and then calculate the percentage of agreement between the raters, so if the raters agree 8 out of 10 times, the test has an $80 \%$ inter-rater reliability rate.

The above approaches in applying the reliability of tests have been in classical test theory, but this study adopted the empirical reliability approached to IRT. Thus, reliability is a central notion in educational measurement and classical test theory. Within item response theory, test reliability is estimated using the information function of the individual item that makes up the test (Baker, 2001). These functions provide a sound basis for choosing items in test constructions. The item information function takes all item parameters into account and shows the measurement efficiency of the item at different ability levels. This reliability in the parlance of IRT is usually estimated as empirical reliability by IRT softwares (e.g BILOG MG) ( De Ayala, 2009).

Under the item response theory framework, reliability is estimated by the volume of information provided by the individual item that makes up a test. As a result, the volume of information, based on a single item can be computed at any ability level and is denoted by $\mathrm{I}_{\mathrm{i}}(\theta)$, where $i$ indexes the item" (Baker, 2001). When this volume of information is plotted against ability the resulting graph is called the item information (Baker, 2001). A very useful IRT property is that individual item information functions sum to test information,

$$
\mathrm{I}(\theta)=\sum_{t=1}^{n} I_{i} \theta \quad--------- \text { Eqn. 2. } 31
$$

Symbolically, item information and Test information are defined respectively by:

$$
\begin{aligned}
& \mathrm{I}_{i}(\theta)=\frac{p i(\theta)^{2}}{p(\theta)[1-p(\theta)} \cdots------ \text { Eqn. } 2.32 \\
& \text { and, } \mathrm{I}(\theta)=\sum_{t=1}^{n} I_{i}(\theta)--------- \text { Eqn, } 2.33
\end{aligned}
$$

Where $\mathrm{P}^{1}(\theta)$ is the first derivative of $\mathrm{P}(\theta)$, and $\mathrm{p}(\theta)$ is the probability of correct response to item $i$ at some $\theta$ level (Leucht \& Hirsch, 1992).

According to Zimowski et al (2003), the volume of information provided by a set of test items at any level is inversely related to the error associated with ability estimates at the ability level. The standard error of the ability estimates at ability level $\theta$ can be written as:
$\operatorname{SE}(\theta) \quad=\quad \frac{1}{\sqrt{1(\theta)}} \ldots \ldots .$. Eqn. 2. 33
This standard error of measurement defines the precision with which an item measures examinee ability at a particular ability level.

Test equating requirement: Equating is the strongest from of linking between the scores on two tests. The goal of equating is to produce a linkage between scores on two test forms such that the scores from each test form can be used as if they had come from the same test (Dorans, Moses \& Eignor, 2010). The above statement implies that score equating allows the scores from both tests to be used interchangeably. Strong requirements must be put on the blueprints for the two tests and on the method used for linking scores in order to establish an effective equating. Therefore, five requirements are widely viewed as necessary for a linking to be an equating (Holland \& Doran, 2006). They are:
a. The equal ability requirement:- The two tests should both be measures of the same construct (latent trait, skill, ability). Evaluation of this criterion requirement can be achieved through several approaches. Prominent among these approaches is the Reduction in Uncertainty index (RiU) (Dorans, 2000). This index is estimated using the correlation coefficient of the test scores obtained by examinees on the two tests. Dorans (2000) stated that when the index of RiU estimated for two tests' scores is greater than $50 \%$, it is concluded that the tests are close enough to measure the same ability. Mathematically, RiU is expressed by the relation:
$\operatorname{RiU}=1-\sqrt{1-\mathrm{r}^{2}}$
Where, $\mathrm{r}=$ correlation coefficient between the two sets of test scores.
b. The equal reliability requirement:- The two tests should have the same level of reliability. Evaluation of reliability of two tests for the purpose of
equivalence assessment requires that the two tests should have equal reliability estimates. However, LiU and Walker (2007), stated that equality of reliability is a necessary but not sufficient condition for tests equivalence. In addition, high reliability is needed to ensure that the equivalent test scores are informative enough to be acceptable (Dorans, 2004). According to Kline (2000); Kline (2005) and Cohen and Swerdlik (2009), this value of reliability for KR20 should not drop below 0.7.
c. The symmetry requirement:- The third requirement that must be fulfilled by two tests in addition to equal ability and equal reliability requirements is the symmetry requirement. Specifically, symmetry condition requires that the transformation used in equating tests must be symmetric (Lord, 1980). This implies that, the function used to transform test scores on test 1 to the scale of test 2 must be the inverse of the function used to transform test scores on test 2 to the scale of test 1 . Evaluation of this requirement is determined using scatterplots of the equating functions. According to Kolen and Brennan (2004; 2014), the functions are symmetrical when for example, an ability score of +2 on test 1 converts to ability score of +3 on test 2 scale, then an ability score of +3 on test 2 must convert to ability score of +2 on test 1 scale.
d. The equity requirement:- It should be a matter of indifference to an examinees as to which of the tests the examinee actually takes.
e. The population invariance requirement:- The equating function used to link the scores of $x$ and $y$ should be the same regardless of the choice of (sub) population from which it is derived. Dorans and Holland (2000) developed quantitative measures of equitability that indicate the degree to which equating functions depend on the subpopulations used to estimate them.

Data collection design used in test score equating. A variety of designs can be used for collecting data for equating. The group of examines included in an equating study should be reasonably representative of the group of examinees who will be administered the test under typical test administration conditions. The choice of a design involves both practical and statistical issues. The linking designs that permit the scaling of item parameters are the following:

Single-Group Design:- The single group design is the simplest data collection design. In the single group design, all examinees in a single sample of examinees from population (P) take both tests. The single group design can provide accurate equating results with relatively small sample sizes (Von Davier, Holland, Thayer, 2004).

Table 1: Design Table for the Single Group (SG) Design

| Population sample | x | y |  |
| :--- | :--- | :--- | :--- |
| p | I | (a) | (a) |

Note (a) indicates examinees in sample for a given row take tests indicated in a given column.

Lack of (a) indicates score data were not collected for that combination of row and column.

The single group design controls for any possibility of different examinee proficiency by having the same examinees take both tests. It has several major uses in practice of scaling and equating. In using this design, however, it is necessary to assume that an examinee's score on the second test form is unaffected by the fact that he or she previously has taken the first form. That is, it must be plausible that practice and other types of order effects can be ignored. In the present study which emphasized NECO and WAEC tests, the single group design was adopted. This is because of the need to ensure that variation in test scores, if any, is not a function of heterogeneity of samples. Consequently, the mean/sigma transformation method was adopted in the study.

According to Von Davier et al, (2004), the design table for the single group design is given in Table

Equivalent group Design ( $E G$ ):- The two tests to be linked are given to equivalent but not identical groups of examinees, chosen randomly (Kolen \& Brennan 2004). In most equating situations, it is impossible to arrange for enough testing time foe every examinee to take more than one test. The simplest solution is to have two separate samples to take each form of the test. In the equivalent group design the equivalent samples are taken from a common population p , one tested with x and the other with y . Because examinees take only one test, the issue of order effects does not arise with the equivalent design.

The table for the equivalent group (EG) design is shown in Table 2

Table 2: Equivalent group design

| Population | Sample | x | y |
| :--- | :--- | :--- | :--- |
| P | 1 | (a) |  |
| P | 2 |  | @ |

Note: @ indicates examinees in sample for a given row take tests indicated in a given column.

Lack of (a) indicates score data that were not collected for that combination of row and column

The equivalent group design is fairly convenient to administer. It does not require that the two tests have any items in common, but this design can be used even when they do have items in common. It also has some limitations. One limitation is that it requires large sample size to produce accurate equating results (Dorcans, Moses \& Eignor, 2010)

Counterbalanced (CB) Design: In one method for counterbalancing, test booklets are constructed that contain form $x$ and form $y$. One half of the test booklets are printed with form x following form y , and the other half are printed with form y following form x . In packaging, test booklets having form x first are alternated with test booklets having form y first. When the test booklets are
handed out, the first examinee takes form x first, the second examinee takes form $y$ first, the second examinee takes form $y$ first, the third examinee takes form x first, and so on. When the booklets are administered, the first and second forms are separately timed. This spiraling process helps to ensure that the examinee group receiving form Y first is comparable to the examinee group receiving form X first (Kolen \& Brennan 2004)

Table 3: Counterbalanced (CB) Design

| Population | sample | x 1 | x 2 | Y 1 | Y 2 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| P | 1 | @a |  |  | (a) |
| P | 2 |  | @ | (a) |  |

Anchor Test or Non equivalent Groups with Anchor Test (NEAT) Design. In anchor test designs there are two populations, P and Q , with a sample of examinees from p taking test x , and a sample from Q taking test y . In addition, both samples take an achor test (Kolen \& Brennan, 2004). The role of the anchor test is to quantify the differences in ability between samples from P and Q that affect their performance on the two tests to be equated, $x$ and $Y$. The best kind of an anchor for equating is a test that measures the same construct, that is x and y measures. The anchor A is usually a shorter and less reliable test than the tests to be equated (Dorcans, Moses \& Eignor, 2010).

Table 4: None-Equivalent Groups with Anchor Test Design

| Population | Sample | x | A | Y |
| :--- | :--- | :--- | :--- | :--- | :--- |
| P | 1 | (a) | (a) |  |
| Q | 2 |  | (a) | (a) |

The anchor test designs requires sample sizes somewhere in between those of the single group and equivalent group designs, although the sample size requirements depend on how strongly correlated the anchor test is with the two tests to be equated and how similar the two populations are. The non-equivalent groups anchor test design is used to collect data in situations in which it is not possible to administer the tests to be equated to the same or random equivalent groups (Doran, Pomerich \& Holland, 2007)

## Empirical Studies.

The review was conducted under the following headings:

Unidimensionality of test items

Local independence of test items

Reliability of a test

Test equating

## Unidimensionality of test items

Metibemu (2016) carried out a study on the comparison of classical test theory and item response theory in the development and equating of physics
achievement tests on Ondo State, Nigeria, a sample of 1423 students were used, factor analysis was used for assessing the Unidimensionality of the test. The result showed that the ratio of eigenvalue to the second eigenvalue was $2: 1$. Therefore, it was concluded that the tests were Unidimensional.

In a related study, investigated by Aliyu and Uruemu (2015) on the development and validation of mathematics achievement test (MAT) using the 3 - parameter logistic model of the item response theory, a sample of 1000 students were randomly selected from the population of all SS 3 students in Delta State. Data were obtained using the mathematics achievement test developed by the research. Factor analysis of principle component analysis (PCA) and rotated component matrix (RCM) was used to established the Unidimensionality of the items. The results showed that rotated factor locally matrixes ranged between 0.706 and 0.993 were obtained indicating that the MAT has construct validity which measure Unidimensionality traits.

## Local independence of test item.

Oke (2012) investigated the local independence of West African Examination Economics test items using secondary school students from Ajeromi-Ifelodim Local Government Area of Lagos State. Tetrachoric correlation coefficient was used to find out the extent to which the items were locally independent. The result showed that there were 2450 correlations, out of which 1865 (76.1\%) had coefficient close to zero, implying that Economics
test items for 2010 Senior School Certificate Examination were locally independent.

In the same vein, a study was carried out by Ojerinde (2013) to ascertain local independence of University Tertiary Matriculation Examination (UTME) pre-test physics test items using vista-tetrachoric software. He grouped the tetrachoric correlation coefficient of physics test items into five groups, namely; (1), pairs of items having correlation coefficient that fell between $0.00-0.10$ (termed very close to zero), (2), pairs of items having correlation coefficients that fell between $0.200-0.299$ (termed close to zero), (3), pairs of items having correlation coefficient that fall between $0.300-0.449$, (4), pairs of items having correlation coefficients that fell between $0.450-0.500$, (5), pairs of items having correlation coefficients greater than 0.500 . The result showed that, the percentage of correlation coefficients that were close to zero was $64.35 \%$. Since a greater number of the correlation coefficients were close to zero, it was concluded that UTME pre-test physics items were locally independent.

Also Ubi et al (2011) carried out a study on item local independence of JAMB mathematics test items of years 2000 - 2003 in Cross River State. Tetrachoric correlation coefficient was used to locate the extent to which the items were locally independent. The results showed that in year 2000, out of 2601 inter-item correlations, only 17 (representing 56\%) which were up to 0.45 and 2334 correlations (representing 89.73\%), were approximately zero
(0). Similar correlations were observed for the other years, thus reflecting that the items were locally independent.

Olabode (2014) investigated on the comparison of local independence of 2012 NECO and WAEC mathematics tests' items. A sample of 500 SS3 students was selected from the population of all SS 3 students in Ogun State. Tetrachoric correlation coefficient was used to locate the extent to which the items were locally independent. The study gave rise to 2450 correlations and there appears not be any correlation in the pairs of items that fall all between 0.500 and above also $0.450-0.499$. The inter-item correlations of $0.500-$ .449 on items appears 1169 ( $96.54 \%$ ), while a larger correlation of 1505 (97.04\%) were close to zero (0). The result showed that WAEC and NECO test items of 2012 were locally independent. The study also found that the NECO 2012 mathematics test were more locally independent than the WAEC 2012 test items.

## Reliability of a test

In a related study, Kolawale (2007) carried out a study comparing the psychometric properties of WAEC and NECO mathematics multiple choice items so as to ascertain whether the two papers were equivalent tests. The samples of the study were made up of 500 senior secondary school students randomly selected from ten local government areas of Ekiti State. Cronbach Alpha reliability principles were applied to obtain reliability coefficients of 0.86
and 0.83 , respectively. The finding revealed that NECO and WAEC multiplechoice items were equivalent.

In a study investigated by Nnanemere, Nwaogu and Osunkwo (2010) on the reliability coefficient and validity indices of mathematics question papers set by NECO, a sample of 40 students that were purposively selected from the population of all SS III students in Government Secondary School, Owerri were used. Data were obtained using question papers .set by National Examination Council (NECO) in 2008. Using the split-half method, KuderRichardson's reliability, coefficient of internal consistency the value 0.75 and 0.69 were obtained indicating that test items set by NECO mathematics question papers were reliable.

Obinne (2008) also examined the psychometric properties of West African Examinations Council and National Examinations Council test items using item response theory. She used a sample of 1800 respondents drawn from the population of all year three SS 3 senior secondary school students who enrolled for the May/June/July 2006 biology senior school certificate examinations of WAEC and NECO in the three education zones of Benue State. Data were obtained using the WAEC and NECO 2000-2002 objective biology examination questions and analysed using the maximum likelihood estimation technique of the Bilog MG computer programme and t-test statistics. Result showed that the items were reliable; items of NECO examinations were relatively more difficult.

In a related study, Bamidele and Adewale (2013) did a comparative analysis of the reliability and validity coefficients of WAEC, NECO and NABTEB constructed mathematics examination questions, using final year students of secondary schools and technical colleges in Nigeria. A total of 600 students were used as sample. The researcher adopted past WAEC; NECO and NABTEB mathematics questions as instrument for the study. T - test, Fisher's transformation and Hotelling William test were techniques used for data analysis showing reliability coefficients of $0.89,0.89$ and 0.77 for WAEC, NECO and NABETEB respectively. The results of the findings showed that WAEC, NECO and NABETEB mathematics achievement tests were highly reliable. It also showed that the examination bodies are comparable and equivalent.

Anagbogu (2009) did a comparative analysis of WAEC and NECO Examinations and students' ability parameters in mathematics objective test using the SS III students that took the 2004 examinations. The researcher adopting a stratified random sampling technique in drawing sample of 873 students from a population of 6749 students. The instruments for the study are the WAEC question paper containing 50 items with four options and the NECO question paper containing 60 items with five options. The statistical techniques used for data analysis were the independent t-test factorial analysis of variance and simple graphical presentation. Result of the findings showed that there was a significant difference between the difficulty level of test items both in item and
person in WAEC and NECO examination instruments, with NECO having a more difficult level in their question paper. It was also found that there was no significant relationship between WAEC and NECO examination instruments and students ability parameter in mathematics.

## Summary of Review of related Literature.

The review covered the concept of unidinensionality of tests, local independence of test items, test reliability and test equating. Item response theory which guided the studies was also reviewed. Item response theory models, application of item response theory, model-data fit and various methods of establishing the reliability of a test were also reviewed under theoretical studies. Some of the empirical studies related to the topic of the research were also reviewed. The current researcher investigated on equivalence of two examining bodies; WAEC and NECO in relation to test equating which were not done or well detailed in all other work reviewed. This is why the researcher therefore deemed it necessary to fill these gaps by embarking on the assessment of the equivalence of WAEC and NECO mathematics multiple - choice tests using item response theory.

## CHAPTER THREE

## METHOD

This chapter discussed the research procedures that were used in this study, namely, research design, area of the study, population of the study, sample and sampling techniques, instrument for data collection, validation of the instrument, reliability of the instrument, method of data collection, method of data analysis.

## Research Design

The design of the study was descriptive survey design. Descriptive survey studies are those studies which aim at collecting data on, and describing it in a systematic manner the characteristics, features or facts about a given population (Nworgu, 2015). This study sought ascertain the comparability or equivalence of WAEC and NECO mathematics multiple - choice tests using the psychometric criteria of test equitability in the framework of item response theory.

## Area of the Study

The study was carried out in Imo State, Nigeria. The capital of the state is Owerri. The State has 27 local government areas, and 3 education zones namely Orlu (12 LGAs), Owerri (9 LGAs) and Okigwe (6 LGAs). The state is bounded in the East and North by Abia State, in the south by Rivers State, and in the West by Anambra State. The inhabitants of the state speak Igbo Language. The
major occupations of the people are farming, trading, teaching and other professions such as engineering, medicine, accounting. The Majority of educated people in the state are in the civil /public service.

## Population of the Study

The population of the study comprised all mathematics multiple-choice test items set by WAEC and NECO from 2000 to date. The population of the study also consisted of 37,036 senior secondary three (SS 3) students in the public secondary schools from the 3 education zones in the state that sat for the May/June 2015 Senior School Certificate Examination. There were 274 public Secondary schools in Imo State. The number of students in each local government area can be found in Appendix 3.

## Sample and sampling Technique

The sample comprised 1051 SSIII students. In obtaining the sample, a combination of non-proportionate stratified random sampling and cluster sampling techniques was used. First the state was stratified into three, according to the education zones. In each of the education zones, 10 public secondary schools were obtained through simple random sampling. Thus gave rise to 30 secondary schools. All the year III senior secondary school (SS III) students in the 30 schools constituted the sample.

## Instrument for data collection

The instrument of the study consisted of mathematics multiple-choice tests set by WAEC and NECO from 2011 to 2014. This consisted of 8 instruments for the 4 years of both examining bodies. The NECO Mathematics test was made up of 60 items. These items had five response options (A, B, C, D and E). The WAEC mathematics test was made up of 50 multiple-choice test items with 4 response options (A, B, C, and D).

## Validation of Instrument

The validity of the tests was not obtained because these instruments were adopted. However, it should be pointed out that validity assessment was part of the parameters the study was designed to determine.

## Reliability of the Instrument

The reliability of the tests was not obtained because these instruments were adopted. However, it should be pointed out that reliability assessments was part of the parameters the study was designed to determine.

## Method of Data Collection

WAEC and NECO mathematics multiple - choice test items of 2011, 2012, 2013 and 2014 were administered to the testees, with the cooperation and support of the principals and mathematics teachers of the selected schools and

30 research assistants, under the supervision of the researcher. The pair of NECO 2011 and WAEC 2011 tests were separately administered within one week simultaneously in the 30 selected schools. The NECO 2012 and WAEC 2012, NECO 2013 and WAEC 2013 and NECO 2014 and WAEC 2014 tests were similarly administered at two weeks interval between each pair. In all, the administration of the four pairs of tests took a total of 10 weeks to be completed.

The administration of the mathematics tests in all the stages were carried out during the mathematics periods of the selected schools. However, in some cases, the administration extended into periods allocated to other subjects. Adherence to instructions as required by the examining bodies was strictly followed. Adequate supervision was provided to avoid cheating.

The exercises were carried out in the first term, to ensure that the sampled schools were covered before they sat for their exams. The duly completed mathematics multiple-choice items totalling 8408 were collected and marked. Each correct response attracted 1 mark, incorrect response attracted 0 mark.

## Method of Data Analysis

Research question 1 was analyzed using factor analysis module of SPSS version 21. According to Lord (1980), when the eigenvalue of the first factor is substantially greater than the second then the data can be assumed to be approximately unidimensional.

Research questions 2 and 3 were analyzed using the tetrachoric correlation module of LISREL version 8.80 and frequency count. Assessment of item local
independence is usually achieved by outright tetrachoric/polychoric correlation among items response on a test (Ubi, 2006 cited in Ubi, Joshua, \& Umoinyang, 2011). According to Lord (1978) in Ubi et al (2011), an item is considered locally independent, if the tetrachoric/polychoric correlations among the items are not significantly different from zero.

Research question 4 was analyzed using Reduction in Uncertainty index ( RiU ). This index is expressed by the relation:
$\operatorname{RiU}=1-\sqrt{1-\mathrm{r}^{2}}$

Where, $\mathrm{r}=$ correlation coefficient between the two sets of test scores.

Dorans (2000) stated that when the index of RiU estimated for two test scores is greater than $50 \%$, it is concluded that the tests are close enough to measure the same ability. In order to achieve this feat, three stages of analysis of the data were advanced. They were; (1), calibration of examinees' ability scores (using BILOG MG), (2), correlation of examinees' ability scores on NECO and WAEC tests (using Pearson' moment correlation with SPSS version 21), and thereafter (3), estimation of RiU index.

When the need arise to estimate the ability estimates of examinees there are three models through which the item parameters and ability scores on multiplechoice test could be estimated: These are $1-\mathrm{PL}, 2-\mathrm{PL}$, and $3-\mathrm{PL}$. In order to test whether or not it is reasonable to model the tests items of NECO and WAEC Mathematics tests according to the 1-, 2- or 3- parameter logistic model, the tests
were subjected to the calibration module of BILOG MG using the 1PL, 2PL, and 3 PL respectively.

When the 1-PL was used across the NECO Mathematics tests of years 2011, 2012, 2013, and 2014 it was observed that $30,28,23$, and 19 items respectively were deleted by the model and for WAEC Mathematics tests it was observed that $25,21,26$ and 24 items respectively were deleted by the model. When the 2-PL was used across the NECO Mathematics tests of years 2011, 2012, 2013, and 2014 it was observed that $23,21,19$, and 16 items respectively were deleted by the model and for WAEC Mathematics tests it was observed that $21,19,20$ and 19 items respectively were deleted.

When the 3-PL was used across the NECO Mathematics tests of years 2011, 2012, 2013, and 2014 it was observed that $11,2,8$, and 2 items respectively were deleted by the model and for WAEC Mathematics tests it was observed that 1 , nil, 2 and nil items respectively were deleted by the model. The implication of these results was that the 3-PL model was the most appropriate for the calibration of the tests items. Therefore the two tests were calibrated using the 3-PL model.

Research question 5 was analysed using empirical reliability of the calibration module of BILOG MG version 3.0. According to Cohen and Swerdlik (2009); Kline (2000) and Kline (2005), this value of reliability should not drop below 0.7.

Research question 6 was analysed using Linear equating transformation equations and scatter plots. According to Kolen and Brennan (2004; 2014), the functions are symmetrical when for example, an ability score of +2 on test 1 converts to ability score of +3 on test 2 scale, then an ability score of +3 on test 2 must convert to ability score of +2 on test 1 scale. The procedure is presented as follows: The ability scores of examinees in NECO were placed unto the same metric scale with the ability scores of examinees on WAEC test and the ability scores of the examinees on WAEC test were placed on the same metric scale with the ability scores on NECO Mathematics test using;
$x_{N}=\frac{\sigma_{N}}{\sigma_{W}} x_{W}+\left(\mu_{N}-\frac{\sigma_{N}}{\sigma_{W}} \mu_{W}\right)--\cdots-\cdots-----$ eqn $b$ and
$x_{W}=\frac{\sigma_{W}}{\sigma_{N}} x_{N}+\left(\mu_{W}-\frac{\sigma_{W}}{\sigma_{N}} \mu_{N}\right)$------------eqn c respectively.
Where; $x_{N}$ is the WAEC equivalent of an examinee's ability score on NECO mathematics test,
$x_{W}$ is the NECO equivalent of an examinee's ability score WAEC Mathematics test,
$\mu_{N}$ is the mean of the ability scores of examinees on NECO test, $\mu_{W}$ is the mean of the ability scores of examinees on WAEC Mathematics test, $\sigma_{N}$ is the standard deviation of the ability scores of examinees on NECO Mathematics test, and
$\sigma_{W}$ is the standard deviation of the ability scores of examinees on WAEC Mathematics test,

Equation (b) presents the general equation for placing the ability scores on NECO Mathematics test unto the scale of ability scores of WAEC mathematics test and equation c presents the general equation for placing ability scores of WAEC Mathematics test unto the scale of ability scores of NECO mathematics test.

Hypothesis one was tested using Z-transformation of correlation coefficient statistics. The statistic is given by

$$
Z=\sqrt[r]{\frac{N-2}{1-r^{2}}}(\mathrm{df}=\mathrm{n}-2)
$$

Where r is the correlation coefficient and N is the sample size.

In order to achieve this feat, the responses of examinees to the NECO and WAEC Mathematics multiple choice test were subjected to factor analysis respectively and their factor scores were obtained. Thereafter, the resulting factor scores were correlated to assess the equivalence of the unidimensionality of the tests (Gorsuch, 1983). Furthermore, the significance of the correlation coefficients was tested using Z-transformation.

Hypothesis two was tested at 0.05 level of significance using dependent alpha formula. According to Feldt, Woodruff \& Salih (1987), it is expressed as $Z=\frac{\left(\alpha_{1}-\alpha_{2}\right)(N-2)^{1 / 2}}{\left[4\left(1-\alpha_{1}\right)\left(1-\alpha_{2}\right)\left(1-P^{2}\right)\right]^{1 / 2}}-\cdots-\cdots---$ eqn 1 ( $\mathrm{DF}=\mathrm{N}-2$ )

Where $\alpha_{1}=$ reliability estimate of the first test $\alpha_{2}=$ reliability estimate of the second test, $P^{2}=$ the squared Pearson's correlation coefficient between the total test scores obtained by examinees in the two tests and $N$ is the sample size.

## CHAPTER FOUR

## PRESENTATION AND ANALYSIS OF DATA

This chapter presents the results of data analysis. The results are hereby presented in the order in which the research questions and hypotheses were stated.

Research question 1: To what extent are the SSCE mathematics multiplechoice test items of NECO and WAEC undimensional?

Table 5: Analysis of Total Variance Explained by Major Components of 2011 NECO Mathematics Test

| Component | Initial Eigenvalue |  |  | Extraction Sums of Squared Loadings |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of | Cumulative | Total | $\begin{array}{c}\text { \% of } \\ \text { Variance }\end{array}$ | \% | $\left.\begin{array}{c}\text { Cumulati }\end{array}\right)$

From the table, the highest eigenvalue is 6.98 and is for component 1 and it explains $11.6 \%$ of the variance. Component 2 has 3.00 which accounts for $5.0 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 2.33 to 1 (2.33:1). Thus the test is unidimensional

Table 6: Analysis of Total Variance Explained by the Major Components of 2012 NECO Mathematics Test

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of <br> Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ | Total | $\%$ of Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ |
| 1 | 11.355 | 18.925 | 18.925 | 11.355 | 18.925 | 18.925 |
| 2 | 3.155 | 5.259 | 24.183 |  |  |  |
| 3 | 2.649 | 4.415 | 28.598 |  |  |  |
| 4 | 2.452 | 4.086 | 32.684 |  |  |  |
| 5 | 2.091 | 3.485 | 36.170 |  |  |  |
| 6 | 1.847 | 3.078 | 39.248 |  |  |  |
| 7 | 1.659 | 2.765 | 42.013 |  |  |  |
| 8 | 1.457 | 2.429 | 44.441 |  |  |  |
| 9 | 1.413 | 2.354 | 46.796 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 48 | . 397 | . 662 | 93.585 |  |  |  |
| 49 | . 381 | . 635 | 94.220 |  |  |  |
| 50 | . 365 | . 608 | 94.829 |  |  |  |
| 51 | . 363 | . 604 | 95.433 |  |  |  |
| 52 | . 349 | . 582 | 96.015 |  |  |  |
| 53 | . 345 | . 575 | 96.590 |  |  |  |
| 54 | . 325 | . 542 | 97.132 |  |  |  |
| 55 | . 322 | . 536 | 97.668 |  |  |  |
| 56 | . 310 | . 516 | 98.184 |  |  |  |
| 57 | . 302 | . 504 | 98.688 |  |  |  |
| 58 | . 275 | . 459 | 99.147 |  |  |  |
| 59 | . 272 | . 454 | 99.600 |  |  |  |
| 60 | . 240 | . 400 | 100.000 |  |  |  |

From Table 6, the highest eigenvalue is 11.36 and is for component 1. This eigenvalue explains $18.9 \%$ of the variance. The second component has an eigenvalue of 3.26 which explains $5.5 \%$ of the variance. The ratio of the
eigenvalue of the first component to the second component is 3.5 to $1(3.5: 1)$. The test therefore is unidimensional

Table 7: Analysis of Total Variance Explained by the Major Components of 2013 NECO Mathematics Test

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ | Total | \% of <br> Variance | $\begin{gathered} \text { Cumulative } \\ \% \end{gathered}$ |
| 1 | 6.093 | 10.650 | 10.650 | 6.093 | 10.155 | 10.155 |
| 2 | 3.007 | 5.000 | 15.650 |  |  |  |
| 3 | 2.686 | 4.476 | 20.126 |  |  |  |
| 4 | 2.255 | 3.758 | 23.884 |  |  |  |
| 5 | 1.896 | 3.160 | 27.044 |  |  |  |
| 6 | 1.793 | 2.988 | 30.032 |  |  |  |
| 7 | 1.685 | 2.809 | 32.840 |  |  |  |
| 8 | 1.639 | 2.732 | 35.572 |  |  |  |
| 9 | 1.551 | 2.584 | 38.156 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 44 | . 530 | . 883 | 88.953 |  |  |  |
| 45 | . 524 | . 873 | 89.826 |  |  |  |
| 46 | . 500 | . 833 | 90.659 |  |  |  |
| 47 | . 497 | . 828 | 91.486 |  |  |  |
| 48 | . 476 | . 794 | 92.280 |  |  |  |
| 49 | . 472 | . 787 | 93.067 |  |  |  |
| 50 | . 456 | . 761 | 93.827 |  |  |  |
| 51 | . 430 | . 717 | 94.545 |  |  |  |
| 52 | . 422 | . 704 | 95.249 |  |  |  |
| 53 | . 409 | . 682 | 95.930 |  |  |  |
| 54 | . 406 | . 677 | 96.607 |  |  |  |
| 55 | . 391 | . 651 | 97.259 |  |  |  |
| 56 | . 366 | . 611 | 97.869 |  |  |  |
| 57 | . 355 | . 592 | 98.461 |  |  |  |
| 58 | . 330 | . 550 | 99.011 |  |  |  |
| 59 | . 321 | . 534 | 99.545 |  |  |  |
| 60 | . 273 | . 455 | 100.000 |  |  |  |

From Table 7, the highest eigenvalue is 6.09 and is for component 1. This eigenvalue explains $10.65 \%$ of the variance. The second component has an eigenvalue of 3.01 which explains $5.0 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 2.0 to $1(2.0: 1)$. Thus the test is unidimensional.

Table 8: Analysis of Total Variance Explained by the Major Components of 2014 NECO Mathematics Test

| Component | Initial Eigenvalue |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of <br> Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ | Total | \% of <br> Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ |
| 1 | 13.452 | 22.420 | 22.420 | 13.452 | 22.420 | 22.420 |
| 2 | 3.914 | 6.524 | 28.944 |  |  |  |
| 3 | 3.136 | 5.226 | 34.170 |  |  |  |
| 4 | 2.568 | 4.281 | 38.451 |  |  |  |
| 5 | 2.083 | 3.472 | 41.923 |  |  |  |
| 6 | 1.968 | 3.280 | 45.202 |  |  |  |
| 7 | 1.547 | 2.578 | 47.781 |  |  |  |
| 8 | 1.484 | 2.474 | 50.255 |  |  |  |
| 9 | 1.290 | 2.151 | 52.405 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 44 | . 375 | . 624 | 92.064 |  |  |  |
| 45 | . 367 | . 612 | 92.677 |  |  |  |
| 46 | . 362 | . 603 | 93.280 |  |  |  |
| 47 | . 353 | . 588 | 93.868 |  |  |  |
| 48 | . 342 | . 571 | 94.438 |  |  |  |
| 49 | . 340 | . 567 | 95.005 |  |  |  |
| 50 | . 327 | . 545 | 95.550 |  |  |  |
| 51 | . 319 | . 532 | 96.082 |  |  |  |
| 52 | . 306 | . 510 | 96.592 |  |  |  |
| 53 | . 293 | . 488 | 97.080 |  |  |  |
| 54 | . 281 | . 468 | 97.548 |  |  |  |
| 55 | . 274 | . 456 | 98.004 |  |  |  |
| 56 | . 267 | . 444 | 98.448 |  |  |  |
| 57 | . 252 | . 420 | 98.868 |  |  |  |
| 58 | . 242 | . 403 | 99.272 |  |  |  |
| 59 | . 220 | . 367 | 99.639 |  |  |  |
| 60 | . 217 | . 361 | 100.000 |  |  |  |

From Table 8, the highest eigenvalue is 13.45 and is for component 1. This eigenvalue explains $22.42 \%$ of the variance. The second component has an eigenvalue of 3.91 which explains $6.52 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 3.4 to $1(3.4: 1)$. The test is therefore unidimensional

Table 9: Analysis of Total Variance Explained by the Major Components of 2011 WAEC Mathematics Test

| Component | Initial Eigenvalue |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ | Total | $\begin{gathered} \text { \% of } \\ \text { Variance } \end{gathered}$ | $\begin{gathered} \text { Cumulative } \\ \% \end{gathered}$ |
| 1 | 11.914 | 23.828 | 23.828 | 11.914 | 23.828 | 23.828 |
| 2 | 2.651 | 5.302 | 29.131 |  |  |  |
| 3 | 2.345 | 4.691 | 33.822 |  |  |  |
| 4 | 1.975 | 3.950 | 37.771 |  |  |  |
| 5 | 1.638 | 3.276 | 41.047 |  |  |  |
| 6 | 1.503 | 3.006 | 44.054 |  |  |  |
| 7 | 1.414 | 2.828 | 46.882 |  |  |  |
| 8 | 1.376 | 2.752 | 49.634 |  |  |  |
| 9 | 1.272 | 2.543 | 52.177 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 35 | . 448 | . 896 | 89.434 |  |  |  |
| 36 | . 446 | . 893 | 90.327 |  |  |  |
| 37 | . 422 | . 844 | 91.171 |  |  |  |
| 38 | . 416 | . 831 | 92.002 |  |  |  |
| 39 | . 408 | . 816 | 92.819 |  |  |  |
| 40 | . 389 | . 779 | 93.598 |  |  |  |
| 41 | . 375 | . 749 | 94.347 |  |  |  |
| 42 | . 354 | . 709 | 95.056 |  |  |  |
| 43 | . 347 | . 695 | 95.750 |  |  |  |
| 44 | . 338 | . 676 | 96.426 |  |  |  |
| 45 | . 321 | . 641 | 97.068 |  |  |  |
| 46 | . 317 | . 633 | 97.701 |  |  |  |
| 47 | . 308 | . 616 | 98.317 |  |  |  |
| 48 | . 293 | . 585 | 98.902 |  |  |  |
| 49 | . 279 | . 559 | 99.461 |  |  |  |
| 50 | . 269 | . 539 | 100.000 |  |  |  |

From Table 9, the highest eigenvalue is 11.91 and is for component 1 . This eigenvalue explains $23.8 \%$ of the variance. The second component has an eigenvalue of 2.65 which explains $5.3 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 4.5 to $1(4.5: 1)$. Thus the test is unidimensional.

Table 10: Analysis of Total Variance Explained by the Major Components
of 2012 WAEC Mathematics test

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ | Total | \% of <br> Variance | $\begin{gathered} \text { Cumulative } \\ \% \\ \hline \end{gathered}$ |
| 1 | 10.307 | 20.615 | 20.615 | 10.307 | 20.615 | 20.615 |
| 2 | 2.981 | 5.962 | 26.577 |  |  |  |
| 3 | 2.706 | 5.413 | 31.990 |  |  |  |
| 4 | 2.291 | 4.582 | 36.572 |  |  |  |
| 5 | 1.993 | 3.985 | 40.557 |  |  |  |
| 6 | 1.655 | 3.311 | 43.868 |  |  |  |
| 7 | 1.393 | 2.785 | 46.653 |  |  |  |
| 8 | 1.327 | 2.654 | 49.307 |  |  |  |
| 9 | 1.212 | 2.424 | 51.731 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 35 | . 475 | . 951 | 89.212 |  |  |  |
| 36 | . 440 | . 880 | 90.092 |  |  |  |
| 37 | . 438 | . 875 | 90.967 |  |  |  |
| 38 | . 424 | . 848 | 91.816 |  |  |  |
| 39 | . 408 | . 816 | 92.632 |  |  |  |
| 40 | . 404 | . 808 | 93.439 |  |  |  |
| 41 | . 388 | . 775 | 94.215 |  |  |  |
| 42 | . 364 | . 729 | 94.944 |  |  |  |
| 43 | . 363 | . 727 | 95.671 |  |  |  |
| 44 | . 351 | . 702 | 96.373 |  |  |  |
| 45 | . 327 | . 654 | 97.026 |  |  |  |
| 46 | . 318 | . 637 | 97.663 |  |  |  |
| 47 | . 308 | . 615 | 98.278 |  |  |  |
| 48 | . 306 | . 612 | 98.891 |  |  |  |
| 49 | . 294 | . 588 | 99.479 |  |  |  |
| 50 | . 261 | . 521 | 100.000 |  |  |  |

From Table 10, the highest eigenvalue is 10.31 and is for component 1 . This eigenvalue explains $20.6 \%$ of the variance. The second component has an eigenvalue of 2.98 which explains about $6.0 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 3.5 to $1(3.5: 1)$. The test is therefore unidimensional.

Table 11: Analysis of Total Variance Explained by the Major Components of 2013 WAEC Mathematics Test

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of <br> Variance | $\begin{gathered} \text { Cumulative } \\ \% \end{gathered}$ | Total | $\%$ of <br> Variance | Cumulative \% |
| 1 | 10.289 | 20.577 | 20.577 | 10.289 | 20.577 | 20.577 |
| 2 | 3.125 | 6.249 | 26.827 |  |  |  |
| 3 | 2.816 | 5.632 | 32.459 |  |  |  |
| 4 | 2.327 | 4.654 | 37.112 |  |  |  |
| 5 | 1.843 | 3.686 | 40.798 |  |  |  |
| 6 | 1.451 | 2.901 | 43.700 |  |  |  |
| 7 | 1.351 | 2.702 | 46.402 |  |  |  |
| 8 | 1.303 | 2.606 | 49.007 |  |  |  |
| 9 | 1.242 | 2.484 | 51.492 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 34 | . 471 | . 943 | 88.403 |  |  |  |
| 35 | . 454 | . 908 | 89.311 |  |  |  |
| 36 | . 453 | . 906 | 90.216 |  |  |  |
| 37 | . 429 | . 859 | 91.075 |  |  |  |
| 38 | . 421 | . 842 | 91.916 |  |  |  |
| 39 | . 401 | . 802 | 92.719 |  |  |  |
| 40 | . 385 | . 771 | 93.489 |  |  |  |
| 41 | . 382 | . 765 | 94.254 |  |  |  |
| 42 | . 367 | . 734 | 94.988 |  |  |  |
| 43 | . 365 | . 729 | 95.718 |  |  |  |
| 44 | . 341 | . 683 | 96.400 |  |  |  |
| 45 | . 337 | . 675 | 97.075 |  |  |  |
| 46 | . 328 | . 657 | 97.732 |  |  |  |
| 47 | . 316 | . 631 | 98.363 |  |  |  |
| 48 | . 287 | . 574 | 98.937 |  |  |  |
| 49 | . 268 | . 536 | 99.473 |  |  |  |
| 50 | . 264 | . 527 | 100.000 |  |  |  |

From Table 11, the highest eigenvalue is 10.24 and is for component 1 . This eigenvalue explains 20.6 \% of the variance. The second component has an eigenvalue of 3.13 which explains $6.3 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 3.3 to $1(3.5: 1)$. The test is therefore unidimensional.

Table 12: analysis of total variance explained by the major components of 2014 WAEC mathematics test

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | $\%$ of Variance | Cumulative \% |
| 1 | 10.415 | 20.830 | 20.830 | 10.415 | 20.830 | 20.830 |
| 2 | 2.911 | 5.821 | 26.652 |  |  |  |
| 3 | 2.532 | 5.065 | 31.717 |  |  |  |
| 4 | 1.981 | 3.962 | 35.678 |  |  |  |
| 5 | 1.812 | 3.623 | 39.302 |  |  |  |
| 6 | 1.609 | 3.218 | 42.519 |  |  |  |
| 7 | 1.550 | 3.100 | 45.619 |  |  |  |
| 8 | 1.407 | 2.814 | 48.433 |  |  |  |
| 9 | 1.389 | 2.778 | 51.211 |  |  |  |
| + | + | + | + |  |  |  |
| + | + | + | + |  |  |  |
| 34 | . 465 | . 929 | 88.900 |  |  |  |
| 35 | . 450 | . 901 | 89.800 |  |  |  |
| 36 | . 429 | . 859 | 90.659 |  |  |  |
| 37 | . 416 | . 832 | 91.491 |  |  |  |
| 38 | . 413 | . 825 | 92.316 |  |  |  |
| 39 | . 403 | . 806 | 93.122 |  |  |  |
| 40 | . 381 | . 762 | 93.885 |  |  |  |
| 41 | . 377 | . 754 | 94.639 |  |  |  |
| 42 | . 354 | . 708 | 95.346 |  |  |  |
| 43 | . 342 | . 683 | 96.029 |  |  |  |
| 44 | . 324 | . 649 | 96.678 |  |  |  |
| 45 | . 316 | . 632 | 97.310 |  |  |  |
| 46 | . 304 | . 609 | 97.918 |  |  |  |
| 47 | . 284 | . 567 | 98.486 |  |  |  |
| 48 | . 269 | . 537 | 99.023 |  |  |  |
| 49 | . 255 | . 510 | 99.533 |  |  |  |
| 50 | . 233 | . 467 | 100.000 |  |  |  |

From Table 12, the highest eigenvalue is 10.42 and is for component 1. This eigenvalue explains $20.8 \%$ of the variance. The second component has an eigenvalue of 2.91 which explains $5.8 \%$ of the variance. The ratio of the eigenvalue of the first component to the second component is 3.6 to $1(3.6: 1)$. Thus the test is unidimensional.

Research question 2: How locally independent are SSCE mathematics multiple choice test items of NECO and WAEC?

To answer this research question, tetrachoric correlation analysis was conducted using LISREL Version 8.8. Table 14 through table 21 presents statistics of the tetrachoric correlations (abridged, see appendix 6 for complete version) among the NECO and WAEC tests items.

Table 13: Inter-correlation Matrix of the Tetrachoric Correlation among 2011 NECO Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 56 | 57 | 58 | 59 | 60 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.604 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | -0.221 | -0.637 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.140 | 0.625 | -0.574 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.366 | 0.705 | -0.608 | 0.700 | 1.000 |  |  |  |  |  |  |
| + |  |  |  |  |  | + |  |  |  |  |  |
| 56 | 0.155 | -0.061 | -0.135 | 0.024 | 0.099 |  | 1.000 |  |  |  |  |
| 57 | -0.334 | -0.253 | 0.449 | -0.119 | -0.128 |  | -0.403 | 1.000 |  |  |  |
| 58 | 0.311 | 0.002 | -0.109 | -0.233 | -0.247 |  | -0.046 | -0.301 | 1.000 |  |  |
| 59 | 0.077 | 0.162 | -0.380 | 0.139 | 0.159 |  | 0.099 | -0.451 | 0.104 | 1.000 |  |
| 60 | -0.404 | 0.559 | 0.518 | -0.337 | -0.449 | 0.001 | 0.456 | -0.279 | -0.073 | 1.000 |  |

Table 13 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 60 items of 2011 NECO mathematics test.

Table 14: Summary of Tetrachoric Correlation Coefficients Among the 60Item 2011 NECO Mathematics Test

| Item | $\leq 0.099$ | $0.100-$ | $0.200-$ |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 0.199 | 0.299 | $0.300-$ | 0.400 | $\geq 0.500$ | Total count |
| $1-60$ | 41 | 7 | 5 | 4 | 3 | 2 | 59 |
| $2-60$ | 37 | 6 | 6 | 4 | 5 | 4 | 58 |
| $3-60$ | 39 | 2 | 3 | 4 | 4 | 3 | 57 |
| + | + | + | + | + | + | + | + |
| $57-60$ | 2 | - | - | - | 1 | - | 3 |
| $58-60$ | 1 | 1 | - | - | - | - | 2 |
| $59-60$ | 1 | - | - | - | - | - | 1 |
| Total | 1252 | 195 | 134 | 102 | 47 | 32 | 1770 |
|  | $70.7 \%$ | $11.0 \%$ | $7.6 \%$ | $5.8 \%$ | $2.7 \%$ | $1.8 \%$ | $100 \%$ |

The result shows that for item 1, for example, there were 59 correlation coefficients in that it correlates with item 2, item 3, item 4, item 5, and so on up to item 60. Among the 59 correlation coefficients, 41 have values that are less than or equal to $0.099 ; 7$ have values that fall within 0.100 to 0.199 , and 5 have values within 0.200 to 0.299 ; 3 have values within 0.400 to 0.499 ; and 2 have values greater than or equal to 0.500 . The table also presents the tetrachoric correlations between item 2 and items 3, 4, 5 ... 60 (giving 58 correlation coefficients). The table entire table therefore presents the correlation between each of the 60 items and the other remaining items. This gives 1770 correlation coefficients on the whole.

Out of the 1770 correlations among the 60 items, 1252 (representing $70.7 \%$ ) have tetrachoric correlation coefficients equal to or less than $0.099 ; 195$ (representing 11.0\%) have tetrachoric correlation coefficients within the range
0.100 to $0.199 ; 134$ (representing about $7.6 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 102$ (representing $5.8 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 47$ ( representing $2.7 \%$ ) have correlation coefficients within the range 0.400 to 0.499 ; and 32 (representing 1.8\%) have correlation coefficient greater than or equal to 0.500. In all the larger percentage (about 90\%) of the observed tetrachoric correlation coefficients, among the 60-item NECO Mathematics test of year 2011 are less than or equal to 0.299 which is the minimum yardstick for determining level of local independence of test items. Therefore the items for 2011 NECO test are locally independent.

Table 15: Inter-correlation Matrix of the Tetrachoric Correlation among 2012 NECO Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 56 | 57 | 58 | 59 | 60 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.604 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | 0.688 | 0.677 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.584 | 0.481 | 0.527 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.584 | 0.558 | 0.768 | 0.632 | 1.000 |  |  |  |  |  |  |
| + |  |  |  |  |  | + |  |  |  |  |  |
| 56 | 0.427 | 0.418 | 0.306 | 0.183 | 0.236 |  | 1.000 |  |  |  |  |
| 57 | 0.499 | 0.376 | 0.271 | 0.264 | 0.207 |  | 0.775 | 1.000 |  |  |  |
| 58 | 0.618 | 0.439 | 0.419 | 0.355 | 0.384 |  | 0.656 | 0.752 | 1.000 |  |  |
| 59 | 0.540 | 0.339 | 0.273 | 0.390 | 0.275 |  | 0.469 | 0.564 | 0.535 | 1.000 |  |
| 60 | 0.466 | 0.429 | 0.312 | 0.231 | 0.384 |  | 0.508 | 0.573 | 0.670 | 0.532 | 1.000 |

Table 15 presents the summary of frequencies of the observed polychoric/tetrachoric correlation coefficients among the 60 items of 2012 NECO mathematics test.

Table 16: Summary of Tetrachoric Correlation Coefficients among the 60Item 2012 NECO mathematics test

| Item | $\leq 0.099$ | $\begin{aligned} & 0.100- \\ & 0.199 \end{aligned}$ | $\begin{aligned} & 0.200- \\ & 0.299 \end{aligned}$ | $\begin{aligned} & 0.300- \\ & 0.399 \end{aligned}$ | $\begin{aligned} & 0.400 \\ & -0.499 \\ & \hline \end{aligned}$ | $\geq 0.500$ | Total count |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1-60 | 4 | 4 | 10 | 13 | 14 | 14 | 59 |
| 2-60 | 11 | 10 | 13 | 13 | 8 | 3 | 58 |
| 3-60 | 8 | 6 | 16 | 15 | 8 | 4 | 57 |
| + | + | + | + | + | + | + | + |
| 56-60 | - | - | - | - | 1 | 3 | 4 |
| 57-60 | - | - | - | - | - | 3 | 3 |
| 58-60 | - | - | - | - | - | 2 | 2 |
| 59-60 | - | - | - | - | - | 1 | 1 |
| Total | 330 | 290 | 410 | 401 | 211 | 128 | 1770 |
|  | 18.6\% | 16.4\% | 23.16\% | 22.7\% | 11.9\% | 7.2\% | 100\% |

The result shows that for item1, For example, there were 59 correlation coefficients in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 60. Among the 59 correlation coefficients, 4 have values that are less than or equal to 0.099 ; 4 have values within 0.100 to 0.199 , 10 have values within 0.200 to $0.299 ; 13$ have values within 0.300 to $0.399 ; 14$ have values within 0.400 to 0.499 ; and 14 have values greater than or equal to 0.500 . Similarly the table also presents the correlations between item 2 and each of the other subsequent items, giving rise to 58 correlation coefficients. On the whole therefore, the table presents the correlation coefficients between each of the 60 items and all the other remaining items. This gave rise to 1770 correlation coefficients.

Out of the 1770 correlations among the 60 items, 330 (representing $18.6 \%$ ) have tetrachoric correlation coefficients equal to or less than $0.099 ; 290$ (representing 16.4\%) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 410$ (representing about $23.2 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 401$ (representing $22.7 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 211$ ( representing $11.9 \%$ ) have correlation coefficients within the range 0.400 to 0.499 ; and 128 (representing $1.8 \%$ ) have correlation coefficients greater than or equal to 0.500 . In all a larger percentage (about $58.2 \%$ ) of the observed tetrachoric correlation coefficients, among the 60-item NECO Mathematics test of year 2012 are less than or equal to 0.299 which is the minimum yardstick for determining level of local independence of test items. Therefore, the items for 2012 NECO test are locally independent.

Table 17: Inter-correlation Matrix of the Tetrachoric Correlation among 2013 NECO Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 56 | 57 | 58 | 59 | 60 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.105 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | -0.515 | -0.097 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.060 | 0.461 | -0.312 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.076 | 0.226 | -0.335 | 0.614 | 1.000 |  |  |  |  |  |  |
| + | + | + | + | + | + | + |  |  |  |  |  |
| 56 | -0.260 | 0.157 | 0.060 | -0.067 | 0.119 |  | 1.000 |  |  |  |  |
| 57 | 0.133 | -0.163 | 0.112 | -0.018 | -0.168 |  | 0.053 | 1.000 |  |  |  |
| 58 | 0.544 | 0.213 | -0.278 | 0.076 | -0.080 |  | 0.169 | 0.034 | 1.000 |  |  |
| 59 | 0.094 | 0.054 | -0.118 | 0.336 | 0.171 |  | 0.081 | -0.362 | 0.065 | 1.000 |  |
| 60 | -0.026 | -0.118 | 0.289 | 0.029 | 0.194 |  | -0.185 | 0.474 | -0.169 | -0.224 | 1.000 |

Table 17 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 60 items of 2013 NECO mathematics test.

Table 18: Summary of Tetrachoric Correlation Coefficients among the 60-
Item 2013 NECO Mathematics Test

| Item | $\leq 0.099$ | $\begin{aligned} & 0.100- \\ & 0.199 \end{aligned}$ | $\begin{aligned} & 0.200- \\ & 0.299 \end{aligned}$ | $\begin{aligned} & 0.300- \\ & 0.399 \end{aligned}$ | $\begin{aligned} & 0.400 \\ & -0.499 \\ & \hline \end{aligned}$ | $\geq 0.500$ | Total count |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1-60 | 39 | 8 | 2 | 2 | 1 | 7 | 59 |
| 2-60 | 34 | 13 | 10 | - | 1 | - | 58 |
| 3-60 | 40 | 2 | 9 | 3 | 3 | - | 57 |
| + | + | + | + | + | + | + | + |
| 56-60 | 3 | 1 | - | - | - | - | 4 |
| 57-60 | 2 | - | - | - | 1 | - | 3 |
| 58-60 | 2 | - | - | - | - | - | 2 |
| 59-60 | 1 | - | - | - | - | - | 1 |
| Total | 1188 | 237 | 159 | 102 | 57 | 27 | 1770 |
|  | 67.1\% | 13.4\% | 9.0\% | 5.8\% | 3.2\% | 1.5\% | 100\% |

The result shows that for item1, For example, for item 1, there were 59 correlation coefficients, in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 60 . Among the 59 correlation coefficients, 39 have values that are less than or equal to $0.099 ; 8$ have values within 0.100 to 0.199 , and 2 have values within 0.200 to 0.299 ; 2 have values within 0.300 to 0.399 ; and 1 have values within 0.400 to 0.499 ; and 7 have values greater than or equal to 0.500. Similarly, the table presents the correlations between item 2 and the other subsequent items, giving rise to 58 correlation coefficients. On the whole therefore, the table presents the correlation coefficients between each of the 60 items and all the other remaining items. This gave rise to 1770 correlation coefficients.

Out of the 1770 correlations among the 60 items, 1188 (representing $67.1 \%$ ) have tetrachoric correlation coefficients equal to or less than $0.099 ; 237$ (representing 13.4\%) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 159$ (representing about $9.0 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 102$ (representing $5.8 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 57$ ( representing $3.2 \%$ ) have correlation coefficients within the range 0.400 to 0.499 ; and 27 (representing $1.5 \%$ ) have correlation coefficient greater than or equal to 0.500. In all, about $90 \%$ of the observed tetrachoric correlation coefficients, among the 60-item NECO Mathematics test of year 2013 are less than or equal to 0.299 which is the minimum yardstick for determining level of local
independence of test items. Therefore the items for 2013 NECO test are locally independent.

Table 19: Inter-correlation Matrix of the Tetrachoric Correlation among 2014 NECO Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 56 | 57 | 58 | 59 | 60 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.710 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | 0.491 | 0.731 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.582 | 0.710 | 0.593 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.555 | 0.692 | 0.563 | 0.792 | 1.000 |  |  |  |  |  |  |
| + | + | + | + | + | + | + |  |  |  |  |  |
| 56 | 0.311 | 0.380 | 0.218 | 0.343 | 0.339 |  | 1.000 |  |  |  |  |
| 57 | 0.116 | 0.260 | 0.159 | 0.365 | 0.373 |  | 0.446 | 1.000 |  |  |  |
| 58 | -0.453 | -0.388 | -0.520 | -0.396 | 0.302 |  | -0.340 | 0.177 | 1.000 |  |  |
| 59 | 0.726 | 0.703 | 0.506 | 0.626 | 0.559 |  | 0.538 | 0.500 | -0.726 | 1.000 |  |
| 60 | 0.752 | 0.708 | 0.415 | 0.522 | 0.547 |  | 0.351 | 0.255 | -0.532 | 0.795 | 1.000 |

Table 19 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 60 items of 2014 NECO mathematics test.

Table 20: Summary of Tetrachoric Correlation Coefficients among the 60Item 2014 NECO Mathematics Test

| Item | $\leq 0.099$ | $0.100-$ | $0.200-$ |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 0.199 | 0.299 | $0.300-$ | 0.400 | $\geq 0.500$ | Total count |
|  | 8 | 5 | 23 | 7 | 16 | - | 59 |
| $1-60$ | 4 | 2 | 26 | 7 | 9 | 10 | 58 |
| $2-60$ | 7 | 9 | 4 | 17 | 11 | 9 | 57 |
| $3-60$ | + | + | + | + | + | + | + |
| + | 1 | - | - | 1 | 1 | 1 | 4 |
| $56-60$ | - | 1 | 1 | - | - | 1 | 3 |
| $57-60$ | 2 | - | - | - | - | - | 2 |
| $58-60$ | - | - | - | - | - | 1 | 1 |
| $59-60$ | 295 | 253 | 404 | 348 | 263 | 207 | 1770 |
| Total | $26.7 \%$ | $14.3 \%$ | $22.8 \%$ | $19.7 \%$ | $14.0 \%$ | $11.7 \%$ | $100 \%$ |

The result shows that for item1, For example, there were 59 correlation coefficients, in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 60. Among the 59 correlation coefficients, 8 have values that are less than or equal to 0.099 ; have values within 0.100 to 0.199 , and 5 have values within 0.200 to $0.299 ; 23$ have values within 0.300 to 0.399 ; and 16 have values within 0.400 to 0.499 . The table also presents the tetrachoric correlations between item 2 and items 3, 4, 5 ... 60 (giving 58 correlation coefficients). Similarly, the table presents the correlations between item 2 and the other subsequent items, giving rise to 58 correlation coefficients. On the whole therefore, the table presents the correlation coefficients between each of the 60 items and all the other remaining items. This gave rise to 1770 correlation coefficients.

Out of the 1770 correlations among the 60 items, 295 (representing $16.7 \%$ ) have tetrachoric correlation coefficients equal to or less than $0.099 ; 253$ (representing $14.3 \%$ ) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 404$ (representing about $22.8 \%$ ) have tetrachoric correlation coefficient within the range 0.200 to $0.299 ; 348$ (representing 19.7\%) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 263$ ( representing $14.0 \%$ ) have correlation coefficients within the range 0.400 to 0.499; and 207(representing 11.7\%) have correlation coefficients greater than or equal to 0.500 . In all a larger percentage (about $54 \%$ ) of the observed tetrachoric correlation coefficients, among the 60-item NECO Mathematics test of year 2014 are less than or equal to 0.299 which is the minimum yardstick for determining level of local independence of test items. Therefore the items for 2014 NECO test are locally independent.

Table 21: Inter-correlation Matrix of the Tetrachoric Correlation among 2011 WAEC Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 46 | 47 | 48 | 49 | 50 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.645 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | 0.638 | 0.655 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.597 | 0.713 | 0.596 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.293 | 0.444 | 0.482 | 0.561 | 1.000 |  |  |  |  |  |  |
| + | + | + | + | + | + | + |  |  |  |  |  |
| 46 | 0.425 | 0.410 | 0.486 | 0.502 | 0.438 |  | 1.000 |  |  |  |  |
| 47 | 0.311 | 0.358 | 0.486 | 0.514 | 0.580 |  | 0.723 | 1.000 |  |  |  |
| 48 | 0.468 | 0.611 | 0.477 | 0.593 | 0.407 |  | 0.522 | 0.671 | 1.000 |  |  |
| 49 | 0.468 | 0.242 | 0.408 | 0.333 | 0.165 |  | 0.318 | 0.359 | 0.194 | 1.000 |  |
| 50 | 0.371 | 0.365 | 0.344 | 0.346 | 0.525 |  | 0.514 | 0.530 | 0.397 | 0.423 | 1.000 |

Table 21 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 50 items of 2011 WAEC mathematics test.

Table 22: Summary of Tetrachoric Correlation Coefficients among the 60-

## Item 2011 WAEC Mathematics Test

\(\left.\begin{array}{llllllll}\hline Item \& \leq 0.099 \& 0.100- \& 0.200- \& 0.300- \& 0.400 \& \geq 0.500 \& Total count <br>

\& \& 0.199 \& 0.299 \& 0.399 \& -0.499\end{array}\right]\)|  |
| :--- |
| $1-50$ |
| $2-50$ |

The result shows that for item1, For example, there were 49 correlation coefficients in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 50. Among the 49 correlation coefficients, 5 have values that are less than or equal to $0.099 ; 8$ have values within 0.100 to 0.199 , 18 have values within 0.200 to $0.299 ; 1$ have value within 0.300 to $0.399 ; 13$ have values within 0.400 to 0.499 ; and 4 have values greater than or equal to 0.500 tetrachoric correlations between item 2 and items 3, 4, $5 \ldots 60$ (giving 48 correlation coefficients). The entire table therefore presents the correlation between each of the 50 items and the other remaining items. This gives rise to 1225 correlation coefficients on the whole.

Out of the 1225 correlations among the 50 items, 108 (representing 8.8\%) have tetrachoric correlation coefficients equal to or less than $0.099 ; 170$ (representing 13.9\%) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 447$ (representing about $36.5 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 108$ (representing $8.8 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 240$ ( representing $19.6 \%$ ) have correlation coefficients within the range 0.400 to 0.499; and 152(representing 12.4\%) have correlation coefficients greater than or equal to 0.500 . In all, about $60 \%$ of the observed tetrachoric correlation coefficients, among the 50-item WAEC Mathematics test for year 2011 are less than or equal to 0.299 which is the minimum yardstick for determining level of
local independence of test items. Therefore the items for 2011 WAEC test are locally independent.

Table 23: Inter-correlation Matrix of the Tetrachoric Correlation among 2012 WAEC Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 46 | 47 | 48 | 49 | 50 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.610 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | 0.410 | 0.556 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.421 | 0.566 | 0.802 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.111 | 0.322 | 0.614 | 0.435 | 1.000 |  |  |  |  |  |  |
| + | + | + | + | + | + | + |  |  |  |  |  |
| 46 | 0.333 | 0.292 | 0.351 | 0.300 | -0.016 |  | 1.000 |  |  |  |  |
| 47 | 0.308 | 0.300 | 0.524 | 0.510 | 0.191 |  | 0.557 | 1.000 |  |  |  |
| 48 | 0.544 | 0.540 | 0.407 | 0.407 | 0.095 |  | 0.626 | 0.595 | 1.000 |  |  |
| 49 | 0.342 | 0.500 | 0.552 | 0.496 | 0.392 |  | 0.488 | 0.635 | 0.554 | 1.000 |  |
| 50 | 0.398 | 0.469 | 0.575 | 0.508 | 0.391 |  | 0.411 | 0.636 | 0.492 | 0.699 | 1.000 |

Table 23 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 50 items of 2012 WAEC mathematics test.

Table 24: Summary of Tetrachoric Correlation Coefficients among the 50Item 2012WAEC Mathematics Test

| Item | $\leq 0.099$ | $0.100-$ | $0.200-$ | $0.300-$ | 0.400 | $\geq 0.500$ | Total count |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 0.199 | 0.299 | 0.399 | -0.499 |  |  |
| $1-50$ | 4 | 10 | 20 | 6 | 6 | 3 | 49 |
| $2-50$ | - | 3 | 20 | 15 | 5 | 5 | 48 |
| $3-50$ | - | 2 | 9 | 14 | 16 | 6 | 47 |
| + | + | + | + | + | + | + | + |
| $46-50$ | - | - | - | - | 2 | 2 | 4 |
| $47-50$ | - | - | - | - | - | 3 | 3 |
| $48-50$ | - | - | - | - | 1 | 1 | 2 |
| $49-50$ | - | - | - | - | - | 1 | 1 |
| Total | 94 | 192 | 347 | 302 | 174 | 110 | 1225 |
|  | $7.7 \%$ | $15.7 \%$ | $28.6 \%$ | $24.4 \%$ | $14.5 \%$ | $9.0 \%$ | $100 \%$ |

The result shows that for item1, for example, there were 49 correlation coefficients, in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 50. Among the 49 correlation coefficients, 4 have values that are less than or equal to $0.099 ; 10$ have values within 0.100 to $0.199,20$ have values within 0.200 to $0.299 ; 6$ have values within 0.300 to $0.399 ; 6$ have values within 0.400 to 0.499 ; and 3 have values greater than or equal to 0.500 . Similarly the table also presents the correlations between item 2 and each of the other subsequent items, giving rise to 48 correlation coefficients. On the Whole therefore, the table presents the correlation coefficients between each of the 50 items and all the other remaining items. This gave rise to 1225 correlation coefficients

Out of the 1225 correlations among the 50 items, 94 (representing 7.7\%) have tetrachoric correlation coefficients equal to or less than $0.099 ; 192$ (representing $15.7 \%$ ) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 347$ (representing about $28.6 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 302$ (representing $24.4 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 174$ ( representing $14.5 \%$ ) have correlation coefficients within the range 0.400 to 0.499 ; and 110 (representing $9.0 \%$ ) have correlation coefficients greater than or equal to 0.500 . In all a large percentage (52 \%) of the observed tetrachoric correlation coefficients, among the 50-item WAEC Mathematics test of year 2012 are less than or equal to 0.299 which is the minimum yardstick for
determining level of local independence of test items. Therefore the items for 2012 WAEC test are locally independent.

Table 25: Inter-correlation Matrix of the Tetrachoric Correlation among 2013 WAEC Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 46 | 47 | 48 | 49 | 50 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.723 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | 0.697 | 0.718 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.724 | 0.603 | 0.499 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.711 | 0.436 | 0.477 | 0.545 | 1.000 |  |  |  |  |  |  |
| + | + | + | + | + | + | + |  |  |  |  |  |
| 46 | 0.312 | 0.451 | 0.362 | 0.307 | 0.117 |  | 1.000 |  |  |  |  |
| 47 | 0.162 | 0.426 | 0.339 | 0.219 | 0.193 |  | 0.509 | 1.000 |  |  |  |
| 48 | 0.310 | 0.518 | 0.417 | 0.342 | 0.252 |  | 0.631 | 0.369 | 1.000 |  |  |
| 49 | 0.512 | 0.694 | 0.587 | 0.405 | 0.282 |  | 0.531 | 0.480 | 0.712 | 1.000 |  |
| 50 | 0.490 | 0.389 | 0.288 | 0.390 | 0.314 |  | 0.357 | 0.271 | 0.247 | 0.467 | 1.000 |

Table 25 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 50 items of 2013 WAEC mathematics test.

Table 26: Summary of Tetrachoric Correlation Coefficients among the 50Item 2013 WAEC Mathematics Test

| Item | $\leq 0.099$ | $0.100-$ | $0.200-$ | $0.300-$ | 0.400 | $\geq 0.500$ | Total count |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 0.199 | 0.299 | 0.399 | -0.499 |  |  |
| $1-50$ | 5 | 3 | 9 | 14 | 6 | 12 | 49 |
| $2-50$ | 4 | 1 | 10 | 7 | 16 | 10 | 48 |
| $3-50$ | 4 | 2 | 9 | 8 | 16 | 8 | 47 |
| + | + | + | + | + | + | + | + |
| $46-50$ | - | - | - | 1 |  | 3 | 4 |
| $47-50$ | - | - | 1 | 1 | 1 | - | 3 |
| $48-50$ | - | - | 1 | - | - | 1 | 2 |
| $49-50$ | - | - | - | - | 1 | - | 1 |
| Total | 242 | 187 | 279 | 246 | 157 | 114 | 1225 |
|  | $19.8 \%$ | $15.3 \%$ | $22.8 \%$ | $20.0 \%$ | $12.8 \%$ | $9.3 \%$ | $100 \%$ |

The result shows that for item1, For example, there were 49 correlation coefficients in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 50. Among the 49 correlation coefficients, 5 have values that are less than or equal to $0.099 ; 3$ have values within 0.100 to $0.199,9$ have values within 0.200 to $0.299 ; 14$ have values within 0.300 to $0.399 ; 6$ have values within 0.400 to 0.499 ; and 12 have values greater than or equal to 0.500 . Similarly, the table presents the correlations between item 2 and each of the other subsequent items, giving rise to 48 correlation coefficients. On the whole therefore, the table presents the correlation coefficients between each of the 50 items and all the other remaining items. This gave rise to 1225 correlation coefficients.

Out of the 1225 correlations among the 50 items, 242 (representing $19.8 \%$ ) have tetrachoric correlation coefficients equal to or less than $0.099 ; 187$ (representing 15.3\%) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 279$ (representing about $22.8 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 246$ (representing $20.0 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 157$ ( representing $12.8 \%$ ) have correlation coefficients within the range 0.400 to 0.499; and 114(representing 9.3\%) have correlation coefficients greater than or equal to 0.500 . In all a larger percentage (about $58 \%$ ) of the observed tetrachoric correlation coefficients, among the 50-item WAEC Mathematics test of year 2013 are less than or equal to 0.299 which is the minimum yardstick for
determining level of local independence of test items. Therefore the items for 2013 WAEC are locally independent.

Table 27: Inter-correlation Matrix of the Tetrachoric Correlation among 2014 WAEC Mathematics Test Items

| Items | 1 | 2 | 3 | 4 | 5 | + | 46 | 47 | 48 | 49 | 50 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.000 |  |  |  |  |  |  |  |  |  |  |
| 2 | 0.476 | 1.000 |  |  |  |  |  |  |  |  |  |
| 3 | 0.303 | 0.712 | 1.000 |  |  |  |  |  |  |  |  |
| 4 | 0.375 | 0.675 | 0.477 | 1.000 |  |  |  |  |  |  |  |
| 5 | 0.382 | 0.687 | 0.586 | 0.682 | 1.000 |  |  |  |  |  |  |
| + |  |  |  |  |  | + |  |  |  |  |  |
| 46 | 0.186 | 0.253 | 0.321 | 0.186 | 0.203 |  | 1.000 |  |  |  |  |
| 47 | 0.185 | 0.267 | 0.327 | 0.306 | 0.260 |  | 0.381 | 1.000 |  |  |  |
| 48 | 0.187 | 0.257 | 0.371 | 0.232 | 0.070 |  | 0.370 | 0.302 | 1.000 |  |  |
| 49 | 0.335 | 0.555 | 0.474 | 0.460 | 0.449 |  | 0.557 | 0.467 | 0.324 | 1.000 |  |
| 50 | 0.265 | 0.387 | 0.541 | 0.225 | 0.246 |  | 0.471 | 0.437 | 0.491 | 0.542 | 1.000 |

Table 27 presents the summary of frequencies of the observed tetrachoric correlation coefficients among the 50 items of 2014 WAEC mathematics test.

Table 28: Summary of Tetrachoric Correlation Coefficients among the 50-
Item 2014 WAEC Mathematics Test

| Item | $\leq 0.099$ | $0.100-$ | $0.200-$ |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 0.199 | 0.299 | $0.300-$ | 0.400 | $\geq 0.500$ | Total count |  |
| $1-50$ | 3 | 14 | 14 | 12 | 6 | - | 49 |
| $2-50$ | 3 | 7 | 12 | 9 | 8 | 9 | 48 |
| $3-50$ | - | 3 | 11 | 9 | 14 | 10 | 47 |
| + | + | + | + | + | + | + | + |
| $47-50$ | - | - | - | 1 | 2 | - | 3 |
| $48-50$ | - | - | - | 1 | 1 | - | 2 |
| $49-50$ | - | - | - | - | - | 1 | 1 |
| Total | 202 | 215 | 271 | 249 | 176 | 112 | 1225 |
|  | $16.5 \%$ | $17.6 \%$ | $22.1 \%$ | $20.3 \%$ | $14.4 \%$ | $9.1 \%$ | 100 |

The result shows that for item1, for example there were 49 correlation coefficients, in that it correlates with item 2 , item 3 , item 4 , item 5 , and so on up to item 50. Among the 49 correlation coefficients, 3 have values that are less than or equal to $0.099 ; 14$ have values within 0.100 to 0.199 , 14 have values within 0.200 to $0.299 ; 12$ have values within 0.300 to $0.399 ; 6$ have values within 0.400 to 0.499 ; and none fall within values greater than or equal to 0.500 . Similarly, the table presents the correlations between item 2 and each of the other subsequent items, giving rise to 48 correlation coefficients. On the whole therefore, the table presents the correlation coefficients between each of the 50 items and all the other remaining items. This gave rise to 1225 correlation coefficients.

Out of the 1225 correlations among the 50 items, 202 (representing $16.5 \%$ ) have tetrachoric correlation coefficients equal to or less than $0.099 ; 215$
(representing 17.6\%) have tetrachoric correlation coefficients within the range 0.100 to $0.199 ; 271$ (representing about $22.1 \%$ ) have tetrachoric correlation coefficients within the range 0.200 to $0.299 ; 219$ (representing $20.3 \%$ ) have tetrachoric correlation coefficients within the range 0.300 to $0.399 ; 176$ ( representing $14.4 \%$ ) have correlation coefficients within the range 0.400 to 0.499 ; and 112(representing $9.1 \%$ ) have correlation coefficients greater than or equal to 0.500 . In all a larger percentage (about $56 \%$ ) of the observed tetrachoric correlation coefficients, among the 50-item WAEC Mathematics test of year 2013 are less than or equal to 0.299 which is the minimum yardstick for determining level of local independence of test items. Therefore the items for 2013 WAEC test are locally independent.

Research question 3: How equivalent are the local independence of SSCE mathematics multiple choice test items of NECO and WAEC?

In order to answer this research question, the summary of the tetrachoric correlation coefficients among NECO and WAEC tests items' that fall within the ranges of $\leq 0.299$ and $\geq 0.300$ reported in Table 15 through 29) were compared. Table 30 presents the distribution of the correlation coefficients of pairs of the NECO and WAEC items' falling within ranges $\leq 0.299$ and $\geq 0.300$.

Table 29: Distribution of the correlation coefficients of pairs of the NECO and WAEC items' falling within ranges $\leq \mathbf{0 . 2 9 9}$ and $\geq \mathbf{0 . 3 0 0}$

| Correlation Coefficient |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
| Year | Test | $\leq 0.299$ | $\geq 0.300$ | Total count |
| 2011 | NECO | $1581(89.3 \%)$ | $189(10.7 \%)$ | $1770(100 \%)$ |
|  | WAEC | $725(59.2 \%)$ | $500(40.8 \%)$ | $1225(100 \%)$ |
| 2012 | NECO | $1030(58.2 \%)$ | $740(41.8 \%)$ | $1770(100 \%)$ |
|  | WAEC | $633(51.7 \%)$ | $592(48.3 \%)$ | $1225(100 \%)$ |
| 2013 | NECO | $1584(89.5 \%)$ | $186(10.5 \%)$ | $1770(100 \%)$ |
|  | WAEC | $708(57.8 \%)$ | $517(42.2 \%)$ | $1225(100 \%)$ |
| 2014 | NECO | $1005(53.8 \%)$ | $765(47.2 \%)$ | $1770(100 \%)$ |
|  | WAEC | $688(56.2 \%)$ | $537(43.8 \%)$ | $1225(100 \%)$ |
|  |  |  |  |  |

Table 29 shows that in 2011, $89.3 \%$ of item inter-correlation coefficients for NECO had values equal to or less than 0.299 , while for WAEC it was $59.2 \%$. This shows that in 2011, the NECO items were more locally independent than those of WAEC. In 2012, $58.2 \%$ of item inter-correlation coefficients for NECO had values equal to or less than 0.299 , while for WAEC it was $51.7 \%$. This shows that in 2012, the NECO items were more locally independent than those of WAEC. In 2013, 89.5\% of item inter-correlation coefficients for NECO had values equal to or less than 0.299 , while for WAEC it was $57.8 \%$. This shows that in 2013, the NECO items were more locally independent than those of WAEC. Finally in 2014, 53.8\% of item inter-correlation coefficients for NECO had values equal to or less than 0.299 , while for WAEC it was $56.2 \%$. This shows that in 2014, the WAEC items were more locally independent than those of NECO.

Research question 4: To what extent do SSCE mathematics multiple choice tests of NECO and WAEC measure the same ability (i.e., mathematics proficiency)?

The results are presented as follows:
Stage one, ability score estimation
Table 30: Examinees' Ability scores, Mean and Standard Deviationon of NECO and WAEC

|  | 2011 |  | 2012 |  | 2013 |  | 2014 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| S/N | NECO | WAEC | NECO | WAEC | NECO | WAEC | NECO | WAEC |
| 1 | 0.66 | -0.46 | 0.43 | 0.9 | 0.59 | 0.46 | 0.91 | 0.43 |
| 2 | 0.89 | -0.93 | 0.58 | 1.08 | 0.64 | 0.43 | 1.1 | 1.68 |
| 3 | -0.41 | 0.27 | 0.94 | 0.59 | 0.56 | 0.62 | 0.5 | 1.29 |
| 4 | 0.38 | 0.31 | 0.38 | 0.81 | 0.91 | 0.42 | 0.86 | 1.16 |
| 5 | 0.28 | 0.22 | 0.77 | -0.09 | 0.76 | 0.06 | 0.87 | 0.39 |
| 6 | 0.38 | -0.03 | 0.65 | 0.91 | 0.97 | 0.63 | -1.56 | -0.01 |
| 7 | 0.99 | 0.32 | 0.63 | 0.91 | 0.6 | 0.35 | 1.16 | 1.01 |
| 8 | 0.62 | 0.7 | 0.75 | 1.05 | 0.45 | 0.58 | 1.47 | 1.54 |
| 9 | -0.13 | 0.31 | 0.46 | -0.39 | -1.25 | 0.68 | 0.89 | 0.99 |
| + | + | + | + | + | + | + | + | + |
| + | + | + | + | + | + | + | + | + |
| 1042 | -1.05 | -0.66 | 0.29 | 0.75 | -1.82 | -1.26 | -0.13 | -0.05 |
| 1043 | -1.32 | -0.93 | -0.5 | 0.54 | -0.75 | -0.19 | -0.54 | -1.25 |
| 1044 | -0.3 | -0.49 | 0.19 | -0.18 | -0.15 | -1.32 | -0.02 | -1.3 |
| 1045 | -0.85 | -0.85 | -0.1 | 0.7 | -0.42 | -1.17 | 0.15 | -1.4 |
| 1046 | -0.36 | -1.2 | 0.24 | 1.31 | -0.54 | -1.02 | -0.34 | -0.79 |
| 1047 | -0.27 | -0.1 | -0.23 | 0.25 | -0.32 | -1.26 | -0.16 | -2.08 |
| 1048 | -0.82 | -0.27 | -0.74 | 0.39 | 0.27 | -1.28 | -0.36 | -0.81 |
| 1049 | -0.46 | -0.79 | -0.41 | 0.02 | -0.2 | -2.47 | -1.18 | -0.71 |
| 1050 | -0.94 | -0.98 | -0.42 | 0.8 | 0.15 | -1.58 | -0.16 | -0.8 |
| 1051 | -0.8 | -0.57 | 0.78 | 0.29 | -1.61 | -2.05 | -0.56 | -0.75 |
| MEAN | -0.01 | 0.02 | 0.03 | -0.01 | 0.02 | -0.01 | 0.05 | 0.01 |
| STD | 1.065 | 1.127 | 1.103 | 1.072 | 0.973 | 1.113 | 0.988 | 1.043 |

Table 30 shows abridged ability scores of examinees on NECO and WAEC Mathematics tests of years 2011, 2012, 2013, and 2014 obtained from phase 3 of BILOG MG calibration module, and their respective means and standard deviations. The obtained ability scores were correlated.

## Stage two, correlation of NECO and WAEC mathematics ability scores

Table 31 presents the correlation coefficients of NECO and WAEC mathematics tests of years 2011, 2012, 2013, and 2014.

Table 31: Correlation coefficient of NECO and WAEC mathematics tests

|  | WAEC 2011 | WAEC 2012 | WAEC 2013 | WAEC 2014 |
| :--- | :---: | :---: | :---: | :---: |
| NECO 2011 | 0.39 |  |  |  |
| NECO 2012 |  | 0.47 |  |  |
| NECO 2013 |  |  | 0.45 |  |
| NECO 2014 |  |  |  | 0.25 |

Table 31 shows the correlation coefficient of NECO and WAEC Mathematics tests. The correlation coefficients showed that NECO and WAEC Mathematics tests of years 2011, 2012, and 2013 were moderately related $(r=0.39, r=0.47$, and $r=0.45$ respectively). Furthermore, the table showed that the relationship between NECO and WAEC Mathematics tests of year 2014 was low $(r=0.25)$.

## Stage three, RiU index estimation

In order to evaluate the extent of equivalence of the two tests in the measurement of Mathematics proficiency of students, the RiU index for the two tests were estimated. Table 33 presents the estimated RiU index for NECO and WAEC Mathematics tests.

Table 32: Reduction in Uncertainty Index of NECO and WAEC Mathematics Tests

| Year | RiU index | $\%$ |
| :---: | :---: | :---: |
| 2011 | 0.079 | 7.9 |
| 2012 | 0.117 | 11.7 |
| 2013 | 0.107 | 10.7 |
| 2014 | 0.032 | 3.2 |

Table 32 shows that the estimated Reduction in Uncertainty for NECO and WAEC Mathematics tests of Year 2011, 2012, 2013, and 2014 were $7.9 \%$, $11.7 \%, 10.7 \%$, and $3.2 \%$ respectively. These estimated RiU indices are lower than $50 \%$, the acceptable minimum standard for flagging two tests equivalent in the measurement of the same ability. This result showed that the NECO and WAEC tests of 2011, 2012, 2013, and 2014 are not equivalent in the measurement of mathematical ability.

Research question 5: How reliable are NECO and WAEC mathematics multiple choice tests?

In order to answer this research question, the empirical reliability output obtained from the calibration module of BILOG MG phase 3 for NECO and WAEC Mathematics tests were compared. Table 32 presents the comparison of the empirical reliability of NECO and WAEC Mathematics tests.

Table 33: Reliability coefficients of NECO and WAEC Mathematics tests

|  | Test |  |
| :---: | :---: | :---: |
| Year | NECO | WAEC |
| 2011 | 0.84 | 0.93 |
| 2012 | 0.95 | 0.96 |
| 2013 | 0.87 | 0.97 |
| 2014 | 0.96 | 0.96 |

Table 33 shows that the reliability coefficients of NECO mathematics test of years 2011, 2012, 2013, and 2014 are $0.84,0.95,0.87$, and 0.96 respectively. The table further reveals that the reliability of WAEC Mathematics test of years 2011, 2012, 2013, and 2014 was $0.93,0.96,0.97$, and 0.96 respectively. Thus, the NECO and WAEC Mathematics multiple choice test of 2011, 2012, 2013, and 2014 were highly reliable.

Research question 6: Are the equating functions for placing the ability estimates of examinees in mathematics test of WAEC on the scale mathematics test of NECO and vice versa symmetrical?

The results are presented as follows:
Table 34 presents the transformed scores of the examinees from NECO scale to WAEC scale and from WAEC scale to NECO scale (see Appendix 10 for the computations used in the transformation).

Table 34: Distribution of ability scores of examinees on NECO Mathematics test transformed to WAEC scale and
ability scores of examinees on WAEC Mathematics test transformed to the scale of NECO test

| Year |  | 2011 |  |  | 2012 |  |  |  | 2013 |  |  |  |  | 2014 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | NECO | $\operatorname{Tr} x_{N}$ | $\operatorname{Tr} x_{W}$ |  | $\operatorname{Tr} x_{N}$ |  | $\operatorname{Tr} x_{W}$ |  | $\operatorname{Tr} x_{N}$ |  | $\operatorname{Tr} x_{W}$ |  | $\operatorname{Tr} x_{N}$ |  | $\operatorname{Tr} x_{W}$ |  |
| S/N |  |  | WAEC |  | NECO |  | WAEC |  | NECO |  | WAEC |  | NECO |  | WAEC |  |
| 1 | 0.66 | -0.46 | -0.46 | 0.71 | 0.43 | 0.97 | 0.9 | 0.38 | 0.59 | 0.41 | 0.46 | 0.64 | 0.91 | 0.45 | 0.43 | 1.93 |
| 2 | 0.89 | -0.91 | -0.93 | 0.95 | 0.58 | 1.15 | 1.08 | 0.52 | 0.64 | 0.39 | 0.43 | 0.7 | 1.1 | 1.63 | 1.68 | 2.12 |
| 3 | -0.41 | 0.23 | 0.27 | -0.42 | 0.94 | 0.65 | 0.59 | 0.87 | 0.56 | 0.55 | 0.62 | 0.61 | 0.5 | 1.26 | 1.29 | 1.52 |
| 4 | 0.38 | 0.26 | 0.31 | 0.41 | 0.38 | 0.87 | 0.81 | 0.33 | 0.91 | 0.38 | 0.42 | 1.01 | 0.86 | 1.14 | 1.16 | 1.88 |
| 5 | 0.28 | 0.18 | 0.22 | 0.31 | 0.77 | -0.05 | -0.09 | 0.71 | 0.76 | 0.06 | 0.06 | 0.84 | 0.87 | 0.41 | 0.39 | 1.89 |
| 6 | 0.38 | -0.06 | -0.03 | 0.41 | 0.65 | 0.98 | 0.91 | 0.59 | 0.97 | 0.56 | 0.63 | 1.08 | -1.56 | 0.03 | -0.01 | -0.54 |
| 7 | 0.99 | 0.27 | 0.32 | 1.06 | 0.63 | 0.98 | 0.91 | 0.57 | 0.6 | 0.32 | 0.35 | 0.66 | 1.16 | 1 | 1.01 | 2.18 |
| 8 | 0.62 | 0.63 | 0.7 | 0.67 | 0.75 | 1.12 | 1.05 | 0.69 | 0.45 | 0.52 | 0.58 | 0.48 | 1.47 | 1.5 | 1.54 | 2.49 |
| 9 | -0.13 | 0.26 | 0.31 | -0.13 | 0.46 | -0.36 | -0.39 | 0.41 | -1.25 | 0.6 | 0.68 | -1.46 | 0.89 | 0.98 | 0.99 | 1.91 |
| + | + | + | + | + | + | + | + | + | + | + | + | + | + | + | + | + |
| 1037 | -1.87 | -0.98 | -1 | -1.97 | -0.5 | -0.41 | -0.44 | -0.53 | -0.87 | -0.89 | -1.03 | -1.03 | -0.87 | -1.3 | -1.41 | 0.15 |
| 1038 | -0.48 | -1.19 | -1.23 | -0.5 | -0.56 | -1.17 | -1.18 | -0.58 | 0.3 | -1.08 | -1.25 | 0.31 | -1.36 | -0.22 | -0.27 | -0.34 |
| 1039 | -1.89 | -0.7 | -0.71 | -1.99 | -2.73 | -1.39 | -1.39 | -2.69 | -1.56 | -0.89 | -1.03 | -1.81 | -0.86 | -0.57 | -0.64 | 0.16 |
| 1040 | -1.82 | -0.59 | -0.59 | -1.92 | -0.71 | -1.42 | -1.42 | -0.73 | -1.84 | -0.73 | -0.85 | -2.13 | -0.21 | -1.38 | -1.5 | 0.81 |
| 1041 | -0.55 | -0.71 | -0.72 | -0.57 | -0.43 | 0.3 | 0.25 | -0.46 | -0.87 | -0.09 | -0.12 | -1.03 | -0.7 | -1.14 | -1.25 | 0.32 |
| 1042 | -1.05 | -0.65 | -0.66 | -1.1 | 0.29 | 0.81 | 0.75 | 0.24 | -1.82 | -1.09 | -1.26 | -2.11 | -0.13 | -0.01 | -0.05 | 0.89 |


| 1043 | -1.32 | -0.91 | -0.93 | -1.39 | -0.5 | 0.6 | 0.54 | -0.53 | -0.75 | -0.16 | -0.19 | -0.89 | -0.54 | -1.14 | -1.25 | 0.48 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1044 | -0.30 | -0.49 | -0.49 | -0.31 | 0.19 | -0.15 | -0.18 | 0.14 | -0.15 | -1.14 | -1.32 | -0.2 | -0.02 | -1.19 | -1.3 | 1 |
| 1045 | -0.85 | -0.83 | -0.85 | -0.89 | -0.1 | 0.76 | 0.7 | -0.14 | -0.42 | -1.01 | -1.17 | -0.51 | 0.15 | -1.29 | -1.4 | 1.17 |
| 1046 | -0.36 | -1.16 | -1.2 | -0.37 | 0.24 | 1.39 | 1.31 | 0.19 | -0.54 | -0.88 | -1.02 | -0.65 | -0.34 | -0.71 | -0.79 | 0.68 |
| 1047 | -0.27 | -0.12 | -0.1 | -0.28 | -0.23 | 0.3 | 0.25 | -0.26 | -0.32 | -1.09 | -1.26 | -0.4 | -0.16 | -1.93 | -2.08 | 0.86 |
| 1048 | -0.82 | -0.29 | -0.27 | -0.86 | -0.74 | 0.44 | 0.39 | -0.76 | 0.27 | -1.11 | -1.28 | 0.28 | -0.36 | -0.73 | -0.81 | 0.66 |
| 1049 | -0.46 | -0.78 | -0.79 | -0.48 | -0.41 | 0.06 | 0.02 | -0.44 | -0.2 | -2.15 | -2.47 | -0.26 | -1.18 | -0.63 | -0.71 | -0.16 |
| 1050 | -0.94 | -0.96 | -0.98 | -0.98 | -0.42 | 0.86 | 0.8 | -0.45 | 0.15 | -1.37 | -1.58 | 0.14 | -0.16 | -0.72 | -0.8 | 0.86 |
| 1051 | -0.8 | -0.57 | -0.57 | -0.84 | 0.78 | 0.34 | 0.29 | 0.72 | -1.61 | -1.78 | -2.05 | -1.87 | -0.56 | -0.67 | -0.75 | 0.46 |

Table 34 shows the abridged (see Appendix 8 for the complete version) ability scores of examinees on NECO mathematics test and the corresponding scores on the WAEC test scale, and the ability scores of examinees on WAEC Mathematics test and the corresponding ability scores on NECO test scale. The scatter plots of these functions are presented as follows;

Year 2011


Figure 4.1(a): scatter plot of examinees' ability scores on NECO mathematics test of year 2011 and the examinees ability scores on the scale of WAEC mathematics test of 2011.


Figure 4.1(b): scatter plot of examinees' ability scores on WAEC mathematics test of year 2011 and the examinees ability scores on the scale of NECO mathematics test of 2011.

Figures 4.1(a) and 4.1(b) present the scatter plots of examinees' ability scores on NECO Mathematics test of year 2011 and the examinees ability scores when converted to the WAEC Mathematics test of 2011 scale, and the examinees' ability scores on WAEC Mathematics test of year 2011 and the examinees ability scores when converted to the NECO Mathematics test of 2011 scale. The slant lines on the graphs are the lines of best fit of the data.

Figure 4.1 (a) shows that an ability of +2 on NECO Mathematics test converts to ability score of +0.8 on the WAEC Mathematics test scale. While figure 4.1 (b) shows that ability score of +0.8 on WAEC Mathematics converts to ability score of +0.4 on NECO mathematics test scale. The results showed that the functions used in transforming the NECO abilities scores to WAEC scale and vice versa is not symmetrical

For year 2012.


Figure 4.2(a): scatter plot of examinees' ability scores on NECO mathematics test of year 2012 and the examinees ability scores on the scale of WAEC mathematics test of 2012.


Figure 4.2(b): scatter plot of examinees' ability scores on WAEC mathematics test of year 2012 and the examinees ability scores on the scale of NECO mathematics test of 2012.

Figures $4.2(\mathrm{a})$ and $4.2(\mathrm{~b})$ present the scatter plots of examinees' ability scores on NECO Mathematics test of year 2012 and the examinees ability scores when converted to the WAEC Mathematics test of 2012 scale, and the examinees' ability scores on WAEC Mathematics test of year 2012 and the examinees ability scores when converted to the NECO Mathematics test of 2012 scale. The slant lines on the graphs are the line of best fit of the data.

Figure 4.2 (a) shows that an ability of +3 on NECO Mathematics test converts to ability score of +1.3 on the WAEC Mathematics test scale. While
figure 4.2 (b) shows that ability score of +1.3 on WAEC Mathematics converts to ability score of +0.6 on NECO mathematics test scale. The results revealed that the functions used in transforming the NECO abilities scores to WAEC scale and vice versa is not symmetrical


Figure 4.3(a): scatter plot of examinees' ability scores on NECO mathematics test of year 2013 and the examinees ability scores on the scale of WAEC mathematics test of 2013.


Figure 4.3(b): scatter plot of examinees' ability scores on WAEC mathematics test of year 2013 and the examinees ability scores on the scale of NECO mathematics test of 2013.

Figures 4.3(a) and 4.3(b) present the scatter plots of examinees' ability scores on NECO Mathematics test of year 2013 and the examinees ability scores when converted to the WAEC Mathematics test of 2013 scale, and the examinees' ability scores on WAEC Mathematics test of year 2013 and the examinees ability scores when converted to the NECO Mathematics test of 2013 scale. The slant lines on the graphs are the lines of best fit of the data.

Figure 4.3 (a) shows that an ability of +1 on NECO Mathematics test converts to ability score of +0.4 on the WAEC Mathematics test scale. While figure $4.3(\mathrm{~b})$ shows that ability score of +0.4 on WAEC Mathematics converts to ability score of +0.2 on NECO mathematics test scale. The results showed that the functions used in transforming the NECO abilities scores to WAEC scale and vice versa is not symmetrical.

For Year 2014


Figure 4.4(a): scatter plot of examinees' ability scores on NECO mathematics test of year 2014 and the examinees ability scores on the scale of WAEC mathematics test of 2014.


Figure 4.4(b): scatter plot of examinees' ability scores on WAEC mathematics test of year 2014 and the examinees ability scores on the scale of NECO mathematics test of 2014.

Figures 4.4(a) and 4.4(b) present the scatter plots of examinees' ability scores on NECO Mathematics test of year 2014 and the examinees ability scores when converted to the WAEC Mathematics test of 2014 scale, and the examinees' ability scores on WAEC Mathematics test of year 2014 and the examinees ability scores when converted to the NECO Mathematics test of 2014 scale. The slant lines on the graphs are the line of best fit of the data.

Figure 4.4(a) shows that an ability of +1.5 on NECO Mathematics test converts to ability score of +0.7 on the WAEC Mathematics test scale, while figure 4.4(b) shows that ability score of +0.7 on WAEC Mathematics converts to ability score of +1.5 on NECO mathematics test scale.

## HYPOTHESES

1. There is no significant relationship between the unidimensionality of NECO and WAEC Mathematics multiple choice tests.

The results are presented as follows:

Table 35: Z-transformation of correlation coefficient between NECO and WAEC Mathematics multiple choice test unidimensionality

|  | N | R | $\mathrm{Z}_{\text {cal }}$ | $\mathrm{Z}_{\text {tab }}$ | Decision |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2011 | 1051 | 0.487 | 18.06 | 1.96 | Reject, Ho |
| 2012 | 1051 | 0.556 | 21.665 | 1.96 | Reject, Ho |
| 2013 | 1051 | 0.607 | 24.738 | 1.96 | Reject, Ho |
| 2014 | 1051 | 0.369 | 12.859 | 1.96 | Reject, Ho |

Table 35 presents the Z-transformation of correlation coefficient of the factor scores of NECO and WAEC Mathematics multiple choice tests showing the extent of similarity of the unidimensionality of the tests. The table shows that the unidimensionality of NECO and WAEC Mathematics multiple choice test were moderately related across 2011, 2012, 2013 and 2014 respectively ( $\mathrm{r}=$ $0.487,0.556,0.607$ and 0.369$)$.For $2011,2012,2013$, and 2014, the calculated Zvalue was greater than the critical value (1.96) at 0.05 level of significance.Therefore the null hypothesis was rejected. Hence there is significant relationship between the unidimensionality of NECO and WAEC Mathematics multiple choice tests from 2011 to 2014.
2. There is no significant difference between the reliability estimates of NECO and WAEC Mathematics multiple choice tests.

The results are presented as follows:

Table 36: Dependent alpha of significance test of reliability estimates of NECO and WAEC Mathematics multiple choice tests

| Year |  | Reliability | N | P | $\mathrm{P}^{2}$ | $\mathrm{Z}_{\text {Cal }}$ | $\mathrm{Z}_{\text {Tab }}$ | Decision |
| :---: | :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- |
| 2011 | NECO | 0.84 | 1051 | 0.39 | 0.15 | -14.94 | 1.96 | Reject, Ho |
|  | WAEC | 0.93 |  |  |  |  |  |  |
| 2012 | NECO | 0.95 | 1051 | 0.47 | 0.22 | -4.10 | 1.96 | Reject, Ho |
|  | WAEC | 0.96 |  |  |  |  |  |  |
| 2013 | NECO | 0.87 | 1051 | 0.45 | 0.20 | -28.99 | 1.96 | Reject, Ho |
|  | WAEC | 0.97 |  |  |  |  |  |  |
| 2014 | NECO | 0.96 | 1051 | 0.25 | 0.06 | 0.00 | 1.96 | do not Reject Ho |
|  | WAEC | 0.96 |  |  |  |  |  |  |

Table 36 presents the dependent alpha of significance test of reliability estimates of 2011 to 2014 NECO and WAEC Mathematics multiple choice tesst showing the extent of equivalence of the reliability estimates of the tests. For 2011, 2012, and 2013 respectively, the calculated Z-value was greater than the critical value (1.96) at 0.05 level of significance. As a result, for 2011, 2012, and 2013 respectively, the null hypothesis was rejected. Thus, there was significant difference between the reliability estimates of NECO and WAEC Mathematics multiple choice tests of 2011, 2012 and 2013. Furthermore, the table revealed that for 2014, the critical value was greater than the calculated Zvalue. Therefore for 2014, the null hypothesis was not rejected. This implies
that there was no significant difference between the reliability estimates of NECO and WAEC Mathematics multiple choice tests of 2014.

## Summary of the major Findings

From the analyses presented in this chapter, the major findings that emerged from the study revealed that;

1. NECO and WAEC Mathematics tests of 2011, 2012, 2013, and 2014 were all unidimensioal.
2. There is significant relationship between the unidimensionality of NECO and WAEC Mathematics multiple choice tests from 2011 to 2014.
3. NECO and WAEC mathematics test items of years 2011, 2012, 2013, and 2014 were locally independent.
4. NECO Mathematics test items of years 2011, 2012, and 2013 were more locally independent than the WAEC test items of the respective years, while in 2014 WAEC test items were more locally independent than those of NECO.
5. The certainty of equivalence of NECO and WAEC Mathematics tests in the measurement of same ability is very low across the years (i.e., years; 2011, 2012, 2013, and 2014).
6. NECO and WAEC Mathematics test for 2011, 2012, and 2013 had high but different reliability estimates, while in 2014 NECO and WAEC Mathematics tests had high and equal reliability estimates.
7. There was significant difference between the reliability estimates of NECO and WAEC Mathematics multiple choice tests of 2011, 2012 and 2013. For 2014 there was no significant different between the estimated reliabilities of NECO and WAEC.
8. The functions used in transforming ability scores of examinees on NECO mathematics test to the scale of WAEC test and the function used in transforming ability scores on WAEC test to the scale of NECO test for year 2011, 2012, and 2013 were far from being symmetrical, while in 2014 the functions were approximately symmetrical.
9. NECO and WAEC mathematics tests of 2011, 2012, 2013, and 2014 did not fulfil all the conditions that are required for test scores obtained from two tests designed to measure the same ability of examinees to be used interchangeably, therefore they are not equivalent.

## CHAPTER FIVE <br> DISCUSSION OF RESULTS, CONCLUSION AND <br> RECOMMENDATIONS

This chapter is divided into the following sub-headings: Discussion of results, Conclusion(s), Implications of the study, Recommendation(s), Limitations of the study, Suggestions for further study.

## Discussion of Results

The discussion of result was done under the following subheadings: unidimensionality, comparability of local independence, Equivalence of WAEC and NECO Mathematics tests in the measurement of the same ability, reliability estimates of WAEC and NECO Mathematics test, symmetry of functions for transforming the ability score of WAEC to the NECO scale and NECO to the scale of WAEC.

## Unidimensionality of NECO and WAEC Mathematics tests

The results revealed that the NECO and WAEC mathematics tests of year 2011, 2012, 2013, and 2014 respectively as principal component analysis produced eigenvalues whose ratio of the first component to that of the second component were equal to or greater than two to one $(2: 1)$ which is the minimum standard for flagging a test unidemensional. These findings suggest that NECO and

WAEC mathematics tests were unidimensional. This implies that NECO and Mathematics tests measure approximately, one dominant trait (i.e., Mathematics proficiency) of examinees.

On the comparability of the unidimensionality of the WAEC and NECO Mathematics tests, the results revealed that there was no significant difference between the unidimensionality of NECO and WAEC mathematics multiple choice tests of 2011, 2012, 2013, and 2014. The findings suggest that WAEC and NECO tests measure mathematics proficiency in similar manner.

The criteria which enforced the results is in line with the set condition for assessing unidimensionality by Hambleton (2004), Orlando, Sherbonve and Thissen (2001). According to Orlando, Sherbonve and Thissen (2001), and Hambleton (2004), a dichotomized test items are considered unidimensional when the first eigenvalue is substantially greater than the next. The findings in term of unidimensionality of the Mathematics tests of the two examining bodies agrees with that of Metibemu (2016), who found that tests with eigenvalue ratio of first component to the second component is $2: 1$ or more were unidimensional.

## Comparability of Local independence of NECO and WAEC mathematics

## tests items

The results revealed that the NECO and WAEC Mathematics test items of years; 2011, 2012, 2013, and 2014 were locally independent as a larger percentage of the correlation coefficients of pairs of NECO and WAEC Mathematics tests' items that fall within the range $\leq 0.299$ (the minimum standard for adjudging locally independent, individual items that is paired together on tetrachoric correlation analysis) were close to zero. This finding agrees with that of Ojerinde (2013) who found that test items set by larger percentage of pairs of items correlation coefficients falling within the range $\leq$ 0.299 .

More importantly, the finding of this study validates the findings of the study of Olabode (2014), in which the results revealed that both NECO and WAEC Mathematics tests items of year 2012 were locally independent when the local independence of NECO and WAEC Mathematics test items of year 2012 were assessed and compared using responses of 500 senior secondary school year III students of Ogun state to WAEC and NECO Mathematics test.

Furthermore, the results revealed that NECO Mathematics test items for years 2011, 2012, and 2013 were more locally independent than the WAEC test items of the respective years as the percentage of inter- item correlation coefficients for NECO Mathematics questions whose values were below 0.300 were greater than that of WAEC.

However, WAEC Mathematics test items for 2014 were more locally independent than NECO Mathematics test items. The findings of the current study as regards the Mathematics test items of the two examining bodies for years 2011, 2012, and 2013 is in line with the findings of the study of Olabode (2014) on the comparison of local independence of NECO and WAEC Mathematics tests items for year 2012. In this regard Olabode's study submitted that NECO Mathematics test items were more locally independent than the WAEC's counterparts.

## Equivalence of NECO and WAEC Mathematics tests in the measurement of the same ability

The results revealed that the certainty of equivalence of NECO and WAEC Mathematics test in the measurement of same ability is very low across the years (i.e., years; 2011, 2012, 2013, and 2014) as the estimated Reduction in Uncertainty for NECO and WAEC Mathematics tests for Years 2011, 2012, 2013, and 2014 were $7.9 \%, 11.7 \%, 10.7 \%$, and $3.2 \%$ respectively. These estimated RiU indices are lower than $50 \%$, the acceptable minimum standard for flagging two tests equivalent in the measurement of the same ability as stated by Dorans (2000; 2004).

## Reliability estimates of NECO and WAEC Mathematics test

The results revealed that the reliability coefficients of NECO and WAEC mathematics tests for years 2011, 2012, 2013, and 2014 respectively were high. The results further revealed that there was significant reliabilities of NECO and WAEC Mathematics tests for years 2011, 2012, and 2013. But for 2014 there was no significant difference between the reliability estimates of NECO and WAEC Mathematics test. The findings suggest that WAEC and NECO Mathematics tests are highly reliable. However, the findings further suggest that the WAEC and NECO mathematics tests reliabilities are not equivalent for 2011, 2012, and 2013. The findings of the current study regarding the reliabilities estimates of NECO and WAEC Mathematics tests agree with those of the studies of Bamidele and Adewale (2013) and Kolawole (2007). These studies found that NECO and WAEC Mathematics tests are highly reliable. On comparison of the Mathematics tests of the two examining bodies for years 2011, 2012, and 2013, the findings of the current study disagree with the submission of the studies of Bamidele and Adewale (2013) and Kolawole (2007) which found that NECO and WAEC mathematics tests had equal reliability estimates. However, the finding of the current study regarding the comparison of reliability estimates of NECO and WAEC Mathematics tests for year 2014 agrees with the findings of Bamidele and Adewale (2013) and

Kolawole (2007) which submitted that NECO and WAEC mathematics tests had equal reliability estimates.

## Symmetry of the functions for transforming ability scores of examinees on

 NECO Mathematics test to the scale of WAEC test and the functions for transforming ability scores of examinees on WAEC Mathematics test to the scale of NECO test.The results revealed that the functions for transforming ability scores of examinees on NECO Mathematics test to the scale of WAEC test and the functions for transforming ability scores of examinees on WAEC Mathematics test to the scale of NECO test for years; 2011, 2012, and2013 were not symmetrical as the scores on NECO scale converted to WAEC scale was not the same with the score on WAEC scale when converted to NECO scale. However, the functions for transforming ability scores of examinees on NECO Mathematics test to the scale of WAEC test and the functions for transforming ability scores of examinees on WAEC Mathematics test to the scale of NECO test for year 2014 were symmetrical as the scores on NECO scale converted to WAEC scale was the same with the score on WAEC scale when converted to NECO scale. The condition which informed the findings is predicated on the assessment of symmetry of equating functions established by Lord (1980) and validated by Kolen and Brennan (2004; 2014). The condition states that for the equating functions used to transform scores on one test form to the other and
vice versa to be symmetrical, the function used to transform test scores on test 1 to the scale of test 2 must be the inverse of the function used to transform test scores on test 2 to the scale of test 1 . That is, if the scatter grams of the equating functions are plotted, an ability score of +2 on test 1 converts to any ability score say +3 on test 2 scale, then an ability score of +3 on test 2 must convert to ability score of +2 on test 1 scale.

## Conclusion

From the findings of the study, the conclusion was that NECO and WAEC mathematics multiple -choice tests for 2011, 2012, 2013 and 2014 are unidimensional and the test items are locally independent. Furthermore, the mathematics multiple-choice tests of the two examining bodies were highly reliable. However, the tests do not fulfil the conditions that are required for test scores obtained from two tests designed to measure the same ability of examinees to be used interchangeably. Therefore, NECO and WAEC mathematics multiple-choice tests are not equivalent in the measurement of examinees proficiency in mathematics. Consequently, NECO and WAEC mathematics test' scores should not be used interchangeably.

## Implications of the study

The study revealed that NECO and WAEC mathematics tests of 2011, 2012, 2013 and 2014 are valid and reliable. Furthermore, in comparative terms, the
two tests possess different validity and reliability estimates. However, the tests are not equivalent, since the tests did not fulfil all the conditions that are required for test scores obtained from two tests designed to measure the same ability of examinees to be used interchangeably.

## Recommendations

Based on the findings of the study, the following recommendations were made:

1. Education authorities should review their stands on the equivalence by government fiat placed on the two examinations conducted by WAEC and NECO mathematics test-items.
2. Government should establish a central body that will be saddled with the responsibility of ensuring that WAEC and NECO test are as equivalent as possible.
3. Government should be motivated at all levels to create a psychometric department/unit in their ministry so as to enhance a more practical drive towards ensuring valid and reliable assessment of students.

## Limitations of the study

Some of the students used for the first administration of the instrument were unable to complete the subsequent instrument, either because they were absent during the administration or intentionally did not want to continue. This affected the number of students whose scores were used to compute the data.

During the time of this study, it was difficult sourcing and procuring the software, like BILOG MG and LISREL used for the study.

## Suggestions for further studies

The following are suggested topics for further investigations

1. Study may be replicated using another population
2. The study may be replicated using another subject offered at SSCE level especially English language.
3. The equivalence of WAEC and NECO mathematics tests of different years may be assessed.
4. The study may be replicated using NECO, WAEC and NABTEB mathematics test.

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## APPENDIX 1: Statistics Of Entry And Results Of WAEC And NECO From 2004 To 2014 For

 Credit Pass.| WAEC |  | NECO |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Year | Total Sat | A1-C6 | $\%$ | Total Sat | A1-C6 | $\%$ |
| 2002 | 908,235 | 309,409 | 34.06 | - | 408,145 | 66.16 |
| 2003 | 926,212 | 341,928 | 36.91 | - | 390,962 | 61.50 |
| 2004 | $1,019,524$ | 346,410 | 33.97 | 973,611 | 381.029 | 50.35 |
| 2005 | $1,054,853$ | 402,982 | 38.20 | 127,1351 | 514,342 | 51.59 |
| 2006 | $1,149,277$ | 472,674 | 41.12 | 897,791 | 434,807 | 48.43 |
| 2007 | $1,249,028$ | 584,024 | 46.75 | 961,455 | 524325 | 54.53 |
| 2008 | 1340,907 | 767,396 | 57.24 | $1,092,215$ | 776,745 | 71.11 |
| 2009 | 1345,052 | 609,849 | 45.34 | $1,219,888$ | 45,827 | 20.84 |
| 2010 | 1306335 | 548,066 | 41.94 | 234,959 | 45,686 | 19.44 |
| 2011 | 1308,965 | 608,866 | 40.35 | $1,318,597$ | 22,9878 | 24.81 |
| 2012 | $1,658.357$ | 838379 | 50.58 | $1,645,577$ | 586,892 | 35.66 |
| 2013 | $1,540,902$ | 864,273 | 56.09 | 958,444 | 291,457 | 31.98 |
| 2014 | $1,658,304$ | 970,921 | 58.54 | $1,646,150$ | 587,044 | 35.66 |

APPENDIX 2: Population of Senior Secondary school Students in Imo State for 2014/2015 academic Session

| ZONES | No. <br> Schools | Of Local Govt. | SSSI | SSSII | SSSIII | TOTAL |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- |
| OKIGWE | 63 | Isiala Mbano | 1138 | 1191 | 2396 | 4725 |
|  |  | Okigwe | 747 | 700 | 1001 | 2448 |
|  |  | Ehime Mbano | 977 | 935 | 1991 | 3903 |
|  |  | Ihitte Uboma | 662 | 628 | 1362 | 2652 |
| ORLU | 98 | Obowo | 690 | 619 | 1079 | 2388 |
|  |  | Ideato North | 679 | 573 | 763 | 2015 |
|  |  | Ideato South | 453 | 486 | 833 | 1772 |
|  |  | Njaba | 694 | 852 | 941 | 2487 |
|  |  | Nkwerre | 736 | 368 | 624 | 1728 |
|  |  | Nwangele | 416 | 414 | 337 | 1167 |
|  |  | Orlu | 1934 | 1868 | 2462 | 12467 |
|  |  | Orsu | 734 | 624 | 643 | 2010 |
|  |  | Oguta | 1146 | 1015 | 1299 | 3460 |
|  |  | Ohaji | 1079 | 1033 | 2251 | 4363 |
|  |  | Oru-East | 1012 | 983 | 1505 | 3500 |
|  |  | Oru-West | 783 | 609 | 890 | 2282 |
|  |  | Ikeduru | 1279 | 1166 | 1995 | 4440 |
|  |  | Mbaitoli | 1911 | 1615 | 1912 | 5438 |
|  |  | Owerri Municipal | 4043 | 2963 | 1404 | 8410 |
|  |  | Owerri North | 1791 | 1724 | 2271 | 5786 |
|  |  | Owerri West | 1897 | 1551 | 2152 | 5600 |
|  |  | Aboh Mbaise | 1307 | 1254 | 1366 | 3927 |
|  |  | Ahiazu Mbaise | 1603 | 1324 | 1335 | 4262 |
|  |  | Ezinihitte Mbaise | 990 | 1038 | 1350 | 3378 |
|  |  | Ngor Okpala | 1304 | 1190 | 1985 | 4479 |
|  |  |  |  |  |  |  |

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Source: Statistics Unit of the State Ministry of Education, Owerri for 2014/2015 academic session.
Appendix 3: SECONDARY SCHOOLS ENROLMENTSSS
OKIGWE ZONE ..... M
F Total M F ..... Total
Isiala Mbano LGA
1 Amaraku Comm Sec. Sch. ..... $74 \quad 106$ ..... 180
2 Amauzari Comp. Sec.Sch. ..... $70 \quad 48$ ..... 118
3 Anara Comm. Sec. Sch ..... 160140 ..... 300
4 Comm.. Sec. Sch. Osuachara ..... 46 ..... 41 ..... 87
5 Comp. Sec. Sch. Mbeke-osu ..... 126 ..... 120246
6 Eziama Sec. Sch. Osu-Ama ..... $72 \quad 78$ ..... 150
7 Ezihe Comm Comm Sec. ..... 8698 ..... 184
Sch
8 Obollo Sec. Tech Sch. ..... 10080 ..... 180
9 Ogbor-Ugiri Comm Sec. Sch ..... $46 \quad 42$ ..... 88
10 Ohohia Sec. Sch ..... 4250 ..... 92
11 Okohia-Osu Tech Colege ..... 83 ..... 83
12 Umuduru-Osu Comm Sec. ..... 115 ..... 159274
Sch
13 Umuneke-Ugiri Sec Sch ..... 8680 ..... 166
14 Umunkwo Girls Sec. Sch ..... 25 ..... 25
15 Umuozu Sec. Sch Ugiri ..... 9591 ..... 25
Total 120111158 ..... 2359
OKIGWE L.G.A
1 Agbobu Comm Sec; Sch. ..... 125118 ..... 243
2 Aku Comm Sec. Sch. ..... 30 ..... $50 \quad 80$
3 Comm.. Sec. Sch. Okigwe ..... 6250 ..... 114
4 Ezinachi Comm. Sec Sch. ..... 70 ..... $75 \quad 145$
5 Girls Sec. Sch. Ihube ..... 32 ..... 68100
6 Umulolo Boys Sec. School ..... 16 ..... 16
7 Umulolo High School ..... 0 ..... 4141
8 Umuowa-Ibu Sec Tech. Sch ..... 9 ..... 13 ..... 22
9 Urban Sec. Sch. Ubaha- ..... 94 ..... 66 ..... 160
Okigwe
Total ..... 438 ..... 481919
Onuimo L.G.A
1 Comm. Sec Sch. Okwe ..... 4450 ..... 94
2 Okigwe National Grammar ..... 6133 ..... 94
School
3 Okwelle Sec. Sch. Okwelle ..... 10 ..... 11 ..... 21
4 Umucheke Okwe Compr ..... $42 \quad 57$ ..... 99
Sec.
5 Umuduru Egbeaguru Sec. ..... 16 ..... 14 ..... 30
Sch338

## SECONDARY SCHOOLS ENROLMENT - IMO STATE

## 2014/2015 <br> OKIGWE ZONE

Ehime Mbano LGA
1 Agbaghara Nsu Comm. Sec. Sch.
2 Agbaia Sec Tech Sch
3 Model Sec. Sch. Umualumaku/Umuihim
4 Comm.. Sec. Sch. Umunumo
5 Compr. Sec. Sch. Umunakanu
6 Dioka Nzerem Comm. Sec. Sch.
7 Ezeoke High Sch, Nsu
8 Ibeafor Sec Sch Umunumo
9 Nsu Compr. High Sch. Umuanunu
10 Umuduru-Nsu Boys Sec. Sch
11 Umueleke/Umueze Sec. Commerical S
12 Umueze 1 Sec. Tech Sch.
13 Umueze 11 Sec. Sch.
14 Umukabia Sec. Sch
15 Umuezeala Sec. Sch
16 Union Compr Sec Sch
17 Umuezeala Ogwara Sec. Sch
Tota
IhitteUboma L.G.A Sch..
1 Abueke Coom Sec. $\quad 78 \quad 68 \quad 146$
2 Amainyi High School
$\begin{array}{lll}23 & 32 & 55\end{array}$
3 Amainyinta Comm Sec. Sch.
$\begin{array}{lll}49 & 39 & 88\end{array}$
4 Amakohia Sec. Sch.
$18 \quad 47 \quad 65$
5 Boys High Sch. Amauzu-Ihitte $\quad 47 \quad 40 \quad 87$
6 Madonna See. Sch Etiti
$\begin{array}{lll}39 & 40 & 79\end{array}$
7 Nwaeruru Mbakwe Compr Sec Sch. Umuihi
8 Okata Compr. Sec Sch.
9 Ohoma Sec Sch Ikperejere
$48 \quad 23 \quad 61$

10 Umuezegwu Sec. Tech Sch.
Total
$65 \quad 40 \quad 105$
$40 \quad 50 \quad 90$

Total
$\begin{array}{lll}52 & 65 & 117\end{array}$
459494953
Obowo LGA
Achara Sec. Commer ${ }_{3}$ Sch. $\quad 82$ 68 $\begin{array}{lll} & 150\end{array}$
Amanze Compr Sec. Sch $\quad \begin{array}{cccc}50 & 80 & 130\end{array}$
$\begin{array}{llll}\text { Avutu Sec Tech Sch } & 30 & 50 & 80\end{array}$
Ehunachi Compr Sec. Sch.
$\begin{array}{lll}64 & 47 & 111\end{array}$
Okenalogho Sec. Tech Sch. :
Okwuohia Comm Sec. Sch.
Umuariam Sec. Tech Sch
$\begin{array}{lll}40 & 31 & 71\end{array}$
$100 \quad 153 \quad 253$
$112 \quad 88 \quad 200$

## SECONDARY SCHOOLS ENROLMENT - IMO STATE

2014/2015<br>ORLU ZONE Ideato North LGA

1 Akokwa Nigh School
2 Akokwa Sec. Tech. Sch.
3 Akpulu Sec. Sch
4 Commercial Sec. Sch. Osina
5 Compr Sec. Sch. Aniche Obinetiti
6 Compr Sec. Sch. Uruala
7 Iheme Mem S.S. Arondizuogu
8 National High School Arondizuogu
9 Obodoukwu Sec. Tech Sch.
Ideato South LGA
1 Amanator Comm Sec. Sch Ogboza
Total
$54 \quad 67$
$\begin{array}{ll}6 & 22 \quad 48\end{array}$
434689
1635
51
4036
76
293160

2 Isiekenesi High Sch.
3 National Sec. Sch. Ntueke
$39 \quad 70 \quad 109$
293338631

4 Sec. Tech. Sch. Dikenafai
$\begin{array}{lll}60 & 36 & 96\end{array}$

5 Ugbele Comm Sec. Sch
$\begin{array}{lll}61 & 77 & 138\end{array}$

6 Umueshi Sec. Sch.
$58 \quad 50 \quad 108$

7 Umuma Isiaku Comp. Sec. Sch.
$59 \quad 24$ 83

8 Umuobom Comm Se. Sch
4654
100
$\begin{array}{ll}22 & 16 \\ 38\end{array}$
$58 \quad 68 \quad 126$

ISU LGA
1 Compr Sec. Sch. Amurie Omanze $\quad \begin{array}{llll}65 & 86 & 151\end{array}$
2 Ebenator Ekwe Sec. Sch.
3 Ekwe Sec. Sch. 150149299
4 Isuniaba Compr Sec. Sch 85 - 85
5 Isunjaba High Sch - 7373
6 Sec. Etch Sch. Amandugba - 128128
NJABA LGA
1 Amucha Sec. Tech Sch.
2 Comp. Sec. Sch Nkume-Isu
3 Girls Sec. Tech. Sch. Umuaka
4 Sec. Compr Sch. Atta
5 Sec. Tech Sch. Okwudor
NWKERE LGA
1 Comm Sec. Sch Amanknara
$\begin{array}{lll}33 & 29 & 62\end{array}$

Total
461464
925
$19 \quad 2241$
$45 \quad 54 \quad 99$
$-\quad-\quad 141 \quad 141$
$46 \quad 98 \quad 144$
$29 \quad 60 \quad 89$
Total 393385778

Total
$461464-925$
$\begin{array}{lll}64 & 86 & 150\end{array}$
Total $174401 \quad 575$
2 Compr Sec Sch. Eziama Obaire ..... 1210 ..... 22
3 Owerre-Kwoii Sec. Sch. ..... 8448 ..... 132
4 Nkwere Hioh Sch ..... 32776
Total ..... 188136324
SECONDARY SCHOOLS ENROLMENT - IMO STATE
2014/2015
NWANGELE LGA
1 Comm. Sec. school Abba2 Comm. Sec Sth Agbaja3 Dick Tiger Mem Sec SchAmaigbo
4 Isu Girls Sec Sch.- $\quad 24 \quad 24$
5 Isu High School
6 King Jaja High Sch AmaigboTotal183198381Orlu LGA1 Comm. Sec. Sch Umuna$114 \quad 84 \quad 198$
2 Comm. Sec Sch Mgbee ..... 180 ..... 105285
3 Com Sec. Sch. Obinugwu ..... $28 \quad 21$ ..... 49
4 Compr Sec Sch. Umuzike ..... 8268 ..... 150
5 Eziachi Sec. Sch. Orlu ..... $70 \quad 65$ ..... 135
6 Girls Uzor Compr Coll ..... 200200
Diioma
7 Green Uzor Compr Coll ..... $66 \quad 75 \quad 141$
Ihioma
8 Ihitte Owerre Comm. Sec. Sch ..... $36 \quad 48$ ..... 84
9 Ogberuru Sec. Sch. ..... $62 \quad 59$ ..... 121
10 Ojike Memorial Sec. Sch ..... $63 \quad 37$ ..... 100
11 OkporoSec ..... 5084 ..... 512
12 Owerre-Ebeiri Comm Sec Sch ..... $42 \quad 90 \quad 132$
13 Sec Tech Sch Umuowa ..... $63 \quad 70 \quad 133$
14 Township Compr Sec. Sch ..... 31 ..... 5081
Amaifeke
15 Umueze Comm Sec. Sch ..... $90 \quad 902992$
Amaike
16 Umutanze Comm Sec. Sch. ..... $83 \quad 90 \quad 173$
Total ..... 15181172635
Orsu LGA
1 Amanachi Commercial Sec. ..... $68 \quad 100 \quad 168$
Sch
2 Comm. Sec Sch Awo-idemili ..... $50 \quad 31 \quad 81$
3 Eziawa Compr. Sec Sch ..... $40 \quad 31 \quad 71$
4 Girls Sec, Sch Awo-idemili ..... $74 \quad 74$
5 Ihittenansa Sec. Sch. ..... 4981 ..... 130
6 Orsu-Ihitteukwa Sec. Sch ..... $90 \quad 62 \quad 152$
7 Umuhu Okabia Sec Sch' ..... 45 ..... $62 \quad 107$
Total ..... 342 ..... 441783
SECONDARY SCHOOLS ENROLMENT - IMO STATE
2014/2015 ..... SSSORUL ZONE $\quad$ M F Total $\mathrm{M} \quad \mathrm{F}$ TotalOguta LGA
1 Agwa Sec. Sch. ..... $144 \quad 106 \quad 250$
2 Comm.. sec Sch Awa ..... 111107 ..... 218
3 Egbuoma Sec. Sch. ..... 58102 ..... 160
4 Ejemekwuru/Akabor Sec. Sch. ..... $60 \quad 60$ ..... 120
5 Eziorsu Sec. Sch. ..... 11190 ..... 201
6 Izombe Sec. Commercial Sch. ..... 3535 ..... 70
7 Priscillia Mem Se ..... 9895 ..... 193
8 St Micheals Sec Sch Orsuobodo ..... $80 \quad 82$ ..... 162
9 Trinity High Sch. ..... $60-60$
10 Umunwamma Girls Sec. Sch Izombe
Total ..... $751 \quad 677 \quad 1434$
Ohaji/Egbema LGA
1 Abacheke Compr Sec Sch Egbema
2 Comm.. Sec Sch Awara ..... $20 \quad 76 \quad 96$
3 Commercial Sec. Sch Assa ..... $46 \quad 68$ ..... 114
4 Mmahu Sec Sch Egbema ..... $\begin{array}{lll}66 & 71 & 137\end{array}$
5 Obosima Sec. Tech School ..... $\begin{array}{lll}72 & 74 & 146\end{array}$
6 Ohuoba Compr Sec. Ohaji ..... $87 \quad 60 \quad 147$
7 Umuapu Sec. Sch. Ohaji ..... 00171 ..... 72
8 Umudike Compr Sec. Sch. ..... $\begin{array}{lll}73 & 77 & 150\end{array}$
9 Umuokanne Compr Sec. Sch. ..... $148 \quad 97 \quad 245$
10 Egbema Sec. Sch ..... $50 \quad 37$ ..... 87
11 Umunwaku Sec. Sch. ..... $\begin{array}{lll}60 & 86 & 146\end{array}$
Total ..... 823718 ..... 1541
Oru East LGA
1 Akatta Sec. Sch. ..... $\begin{array}{lll}65 & 129 & 194\end{array}$
2 Akuma Sec. Sch ..... $\begin{array}{lll}77 & 99 & 176\end{array}$
3 Amiri Girls Sec. Sch ..... $73 \quad 63$ ..... 136
4 Amiri Girls Sec Sch. ..... $120 \quad 120$
5 Compr Sec Sch Awo-Omamma ..... $\begin{array}{lll}52 & 90 & 142\end{array}$
6 Omuma Sec. Tech Sch. ..... $58 \quad 40 \quad 98$
7 Sec Tech. Sch Awo-Omamma ..... 10495 ..... 199
8 Ubogwu Sec. Comm Sch. Awo- ..... $51 \quad 74$ ..... 125
Omamma
Total 480712 ..... 1192
Oru West LGA
1 Comp Sec. Sch Ozara ..... $47 \quad 51 \quad 98$
2 Comp Sec. Sch Ubulu ..... $90 \quad 82$ ..... 172
3 Compr Sec Sch Ibiasegbe ..... $92 \quad 52$ ..... 144
4 Mgbidi Sec Sch ..... 66
5 Nempi Sec. Sch ..... $72 \quad 152$
6 Ohakpu Sec Compr School ..... $44 \quad 83$
7 Otulu Sec. Comml Sch. ..... $30 \quad 18$ ..... 48
8 Umuorji Girls Sec. Sch Mgbidi ..... 113113
Total ..... 374432 ..... 806
SECONDARY SCHOOLS ENROLMENT - IMO STATE
2014/2015
OWERRI ZONEIkeduru
1 Amaimo Comm Girls Sec SchoolSSS
M F Total M F Total
2 Amatta Comm Sec. Sch ..... 5141 ..... 92
3 Atta Boys Sec. Sch ..... $66 \quad 69$ ..... 135
4 Comm. Sec. Sch Ugirike ..... $50 \quad 100$ ..... 150
5 Comm. Sec Sch Inyishi ..... 5134 ..... 85
6 Compr Sec. Sch Avuvu ..... 5248 ..... 100
7 Iho-Dimeze Compr Sec. Sch ..... 10582 ..... 187
8 Ngugo Compr Sec. Sch Ikeduru ..... $37 \quad 46$ ..... 83
9 Owu-Amakohia Sec. Sch ..... $40 \quad 50 \quad 90$
10 Se. Commer Sch. Eziama ..... $62 \quad 75$ ..... 177
11 Umudimsec Sch. ..... $46 \quad 74$ ..... 120
12 Umuoziri Sec. Tech Sch. ..... $48 \quad 72$ ..... 120
13 Uzoagba Sec. Sch. ..... $90 \quad 55 \quad 145$
Total 751677 ..... 1434
Mbaitoli LGA
1 Afara Sec Sch ..... $43 \quad 48$ ..... 91
2 Comm. Sec. Sch Eziama Obiato ..... $46 \quad 54 \quad 100$
3 Comm.. Sec. Sch Mbieri ..... $87 \quad 37$ ..... 124
4 Comm. Sec. Umuonyeali ..... 4632 ..... 78
5 Compr Sec. Sch Ogwa ..... $69 \quad 61 \quad 130$
6 Compr Sec Sch Ubommiri ..... 78 ..... 78
7 Girls Sec Sch Ifakala ..... 37 ..... 37
8 Girls Sec. Sch Ubommiri ..... 98 ..... 98
9 Ifakala Comm Sec Sch ..... $56 \quad 60$ ..... 116
10 IMS Deaf and Dumb Orodo ..... $35 \quad 25$ ..... 60
11 Mbieri Sec. Tech Sch ..... 3418 ..... 52
12 Obazu Girls Sec Sch ..... 211 ..... 211
13 Ogbaku Girls Sec, Sch. ..... - 98 ..... 98
14 Orodo Sec Tech Sch ..... $46 \quad 65$ ..... 111
15 Umueze Ogwa Sec. Sch. Ogwa ..... 1049 ..... 59
16 Umunoha Ogwa Sec. Sch. Ogwa ..... 8179 ..... 166
17 Umunoha Sec Sch ..... 6553 ..... 118
18 Umuobom Comm. Sec. Sch.
MbieriTotal$781 \quad 10551842$
Owerri Municipal
1 Boys Sec. Sch New Owerri ..... 238 ..... 238
2 Comp. Dev. Sec. Sch. Douglas Rd. ..... 174224 ..... 394
Owerri
3 Emmanuel College Owerri ..... 54 ..... 54
4 Govt Sec. Sch. College Owerri ..... 571 - 571
5 Govt. Technical College Owerri ..... 211 ..... 211
6 Ikenegbu Girls Sec. Sch. Owerri ..... 240
7 Urban Dev. Sec. Sch. ..... 58164 ..... 222
8 Owerri City College
$9 \quad$ Young Scientist College Owerri

## SECONDARY SCHOOLS ENROLMENT - IMO STATE

2014/2015 SSS

Owerri North LGA<br>M F Total M F<br>Total Oguta LGA

1 Agbala Sec. Sch. ..... $\begin{array}{lll}60 & 80 & 140\end{array}$
2 Akwakuma Girls Sec. Sch. ..... 203 ..... 203
3 Cassita Maria Sec. Sch. Emekuku ..... $122 \quad 122$
4 Comm. Sec. Sch. Obibiezena ..... $86 \quad 91 \quad 177$
5 Comm. Sec. Sch. Emekuku ..... 4933 ..... 82
6 Compr Sec. Sch. Amakohia ..... 166106 ..... 272
7 Compr Sec. Sch. Emekuku ..... 4238 ..... 80
8 Compr Sec. Sch. Orji
9 Development Sec Sch Mbaoma ..... $43 \quad 30$ ..... 73
Emii
10 Egbu Comp. Sec Sch. ..... $63 \quad 63$ ..... 126
11 Emekuku High School ..... 40
12 Emii Sec. Tech Sch. ..... $83 \quad 75$ ..... 158
13 Naze Sec. Sch. ..... 119121 ..... 140
14 Ogbeke Obibi Sec. Sch. ..... $80 \quad 120$ ..... 200
15 Uratta Sec. Sch. ..... $61 \quad 47$ ..... 108
16 Obube Compr Sec. Sch. Egbele
Total ..... 8921132 ..... 2024
Owerri West LGA
1 Amakohia Ubi Sec. Sch. ..... $43 \quad 47$ ..... 90
2 Ara Sec. Sch ..... 117130 ..... 247
3 Army Day Sec. Sch. Obinze ..... $83 \quad 67$ ..... 150
4 Comp. Sec. Sch. Emeabiam ..... $30 \quad 60$ ..... 90
5 Compr Sec. Sch Avu ..... 3850 ..... 88
6 Eziobodo Sec. Tech Sch ..... $70 \quad 82$ ..... 152
7 Ihiagwa Sec. Sch. ..... 7179 ..... 150
8 Ndegwu Sec. Sch. ..... 144146 ..... 290
9 Nekede Sec. Sch. ..... $59 \quad 61 \quad 120$
10 Oforola Comm Sec Sch ..... $70 \quad 82$ ..... 152
11 Orogwe Comm Sec School ..... 100104 ..... 204
12 Sec. Tech Sch Irete ..... 5342 ..... 95
Total 878950 ..... 1828
SECONDARY SCHOOLS ENROLMENT - IMO STATE
2014/2015OWERRI ZONE
Oboh Mbaise LGA
1 Comm. Sec. School Lagwa ..... $26 \quad 24$ ..... 50
2 Compr Sec. School Amuzu ..... $28 \quad 44 \quad 72$
3 Lorji COmm Sec Sch ..... $45 \quad 50 \quad 95$
4 Mbaise Sec. School ..... 87 ..... 87
5 Mbutu Sec. Commercial School ..... $85 \quad 65 \quad 150$
6 Nguru Sec. Tech School ..... 8565 ..... 150
7 Nguru Sec. Tech School ..... $65 \quad 75$ ..... 140
8 Oke-Ovoro Sec. Sch ..... $22 \quad 87$ ..... 109
9 Okwuato Sec. School ..... 83128 ..... 211
10 Uzunorji Comm Sec. School ..... $74 \quad 51$ ..... 125
Total 622570 ..... 1192
Ahiazu LGA
1 Ahiara Techical College ..... 1811 ..... 182
2 Ahiazu Sec. Sch. Afor-Oru ..... $\begin{array}{lll}34 & 33 & 67\end{array}$
3 Comm. Sec Sch. Obodo Ahiara ..... $\begin{array}{lll}68 & 102 & 170\end{array}$
4 Comm. Sec Sch. Amuzi ..... $44 \quad 81 \quad 125$
5 Ihenworie Sec Sch. ..... $\begin{array}{lll}50 & 60 & 110\end{array}$
6 Okirika Nweke Compr Sec. Sch ..... 3135 ..... 66
7 Oparanadim Comp. Sec. Sch. ..... $\begin{array}{lll}66 & 59 & 125\end{array}$
8 Sec. Comnl Sch.Otulu Ahiara ..... $\begin{array}{lll}35 & 48 & 83\end{array}$
9 Sec. Tech. Sch Obohia ..... $57 \quad 93$ ..... 150
10 Umuokirika Sec. Tech Sch. ..... $\begin{array}{lll}14 & 12 & 26\end{array}$Total$580 \quad 5241104$
Ezinihitte LGA
1 Chokoneze Sec. Tech. School ..... $\begin{array}{lll}30 & 40 & 70\end{array}$
2 Comm. Sec. Sch. Itu ..... $\begin{array}{lll}46 & 38 & 84\end{array}$
3 Comm. Sec. Sch. Obizi ..... $\begin{array}{lll}18 & 32 & 50\end{array}$
4 Eziagbogu Sec Sch. ..... $31 \quad 51$ ..... 82
5 Eziudo Girls High Sch. ..... $91 \quad 91$
6 Eziudo Sec. Tech. Sch. ..... 70 - 70
7 Thitte Ezinihittte Sec. Sch. ..... $\begin{array}{lll}70 & 67 & 137\end{array}$
8 Ime-Onicha Compr Sec Sch. ..... $31 \quad 43 \quad 74$
9 Obizi-High Sch. ..... 85 - 85
10 Oboamma Sec. Sch ..... $30 \quad 45 \quad 75$
11 Okpofe Sec. Sch. ..... $49 \quad 56 \quad 105$
12 Onicha Sec. Sch.
13 Udo Sec. Tech Sch. ..... $10 \quad 10 \quad 20$145
Total ..... 471477948

## SECONDARY SCHOOLS ENROLMENT - IMO STATE

2014/2015SSS
OWERRIZONE ..... M F Total M F Tot
Ngor Okpala ..... al
1 Amaka-Ntu Sec. Sch. ..... 147116 ..... 263
2 Comm. Sec. Sch Mbato ..... 4961 ..... 110
3 Compr. Sec. Sch. Umuekwune ..... $58 \quad 70$ ..... 128
4 Imerienwe Girls Sec. Ch. Ngor Okpala ..... $80 \quad 80$
5 Logara Sec. Sch. ..... 5258 ..... 110
6 Ngor-Okpala High Sch. ..... 11090 ..... 200
7 Nguru-Umaro Sec. Sch. ..... 5641 ..... 97
8 Obiangwu Sec. Comm. Sch. ..... 4347 ..... 90
9 Orisheze Comm. Sec. Sch. Orisheze ..... $120 \quad 112$ ..... 231
10 Owari Gramma Sch. Imerienwe ..... 17089 ..... 259
11 Umuhu Compr. Sec. Sch. ..... 4955 ..... 104
12 Umuhiagu Sec. Sch. Total 899864 ..... 176
Appendix 4: Sampled schools and their population

| s/no | Sample Schools | Number of SS 3 students |
| :---: | :---: | :---: |
| 1 | Technical College Orlu |  |
| 2 | Ebeneator Ekwe Secondary School Okwuato Orlu | 30 |
| 3 | Comprehensive Secondary School, Okwuato Orlu | 19 |
| 4 | Our Saviour Secondary School, Orlu | 17 |
| 5 | Township Comprehensive Secondary School, Amaifeke | 25 |
| 6 | Girls Secondary School, Awo - Idemili | 21 |
| 7 | Obodoukwu Secondary Technical School | 25 |
| 8 | Umueshi Secondary School | 20 |
| 9 | Isu Girls Secondary Abba | 26 |
| 10 | Nkwere High School | 39 |
| 11 | Ahiazu Secondary School, Lude | 38 |
| 12 | Emmanuel College, Owerr | 47 |
| 13 | Ikenegbu Girls Secondary School | 59 |
| 14 | Government College, Owerri | 70 |
| 15 | Girls Secondary School, Ubomiri | 40 |
| 16 | Amakohia Comprehensive Secondary School | 37 |
| 17 | Afara Secondary School | 32 |
| 18 | Ihagwa Secondary School | 48 |
| 19 | Eziudo Girls High School | 25 |
| 20 | Umuohiagu Secondary School | 27 |
| 21 | Umukwo Girls Secondary School | 30 |
| 22 | Anara Secondary School | 32 |

23 Osu Technical College ..... 61
24 Amaraku Secondary School ..... 42
25 Madonna Secondary School, Etiti ..... 52
26 Abueke Comprehensive Secondary School ..... 65
27 Girls Secondary School, Ihube ..... 27
28 Umuariam Secondary Technical School ..... 33
29 Comprehensive Secondary School, Umunakanu ..... 24
30 Umulolo High School ..... 16
Total ..... 1051

Appendix 5: Keys of WAEC Question papers from 2011 to 2014

| S/NO | 2011 | 2012 | 2013 | 2014 |
| :---: | :---: | :---: | :---: | :---: |
| 1 | B | A | C | B |
| 2 | D | D | B | A |
| 3 | D | B | D | C |
| 4 | C | C | D | C |
| 5 | A | C | C | B |
| 6 | B | A | D | A |
| 7 | A | D | B | D |
| 8 | A | A | B | B |
| 9 | A | C | C | C |
| 10 | C | B | D | A |
| 11 | B | D | C | C |
| 12 | B | B | A | D |
| 13 | A | D | A | D |
| 14 | A | C | D | A |
| 15 | B | A | A | B |
| 16 | D | B | B | B |
| 17 | D | A | A | D |
| 18 | C | C | B | A |
| 19 | A | C | C | A |
| 20 | D | B | D | C |
| 21 | C | A | D | B |
| 22 | A | C | D | B |
| 23 | B | C | B | B |
| 24 | B | D | C | B |
| 25 | B | B | C | A |


| 26 | C | B | B | D |
| :---: | :---: | :---: | :---: | :---: |
| 27 | A | C | D | C |
| 28 | B | A | C | C |
| 29 | A | C | C | A |
| 30 | D | C | A | C |
| 31 | B | A | C | C |
| 32 | C | D | B | D |
| 33 | A | C | A | C |
| 34 | A | D | D | C |
| 35 | B | D | D | C |
| 36 | C | B | A | D |
| 37 | A | D | A | A |
| 38 | D | A | B | C |
| 39 | A | A | D | A |
| 40 | C | B | C | C |
| 41 | D | A | A | B |
| 42 | A | C | B | C |
| 43 | D | B | B | C |
| 44 | A | D | A | A |
| 45 | C | C | B | B |
| 46 | B | A | B | C |
| 47 | D | D | C | D |
| 48 | B | B | C | D |
| 49 | B | D | D | C |
| 50 | C | A | A | A |

Appendix 6: Keys of NECO Question papers from 2011 to 2014

| S/NO | 2011 | 2012 | 2013 | 2014 |
| :---: | :---: | :---: | :---: | :---: |
| 1 | B | D | D | B |
| 2 | C | C | A | E |
| 3 | A | B | A | B |
| 4 | B | B | E | D |
| 5 | C | C | E | B |
| 6 | A | C | D | E |
| 7 | D | C | E | D |
| 8 | D | B | B | E |
| 9 | A | E | C | D |
| 10 | A | C | B | C |
| 11 | D | C | B | E |
| 12 | D | C | D | C |
| 13 | A | C | C | D |
| 14 | A | D | E | E |
| 15 | C | B | D | B |
| 16 | D | E | A | B |
| 17 | B | E |  | D |
| 18 | A | E | A | B |
| 19 | C | D | D | D |
| 20 | A | E | C | B |
| 21 | A | E | C | A |
| 22 | B | D | B | B |
| 23 | A | A | E | A |
| 24 | A | A | A | D |
| 25 | C | A | D | C |
| 26 | E | A | B | E |
| 27 | C | A | A | B |
| 28 | A | A | B | A |
| 29 | D | B | B | E |
| 30 | B | D | B | D |


| 31 | B | C | D | A |
| :---: | :---: | :---: | :---: | :---: |
| 32 | C | C | B | E |
| 33 | A | A | B | B |
| 34 | E | D | C | C |
| 35 | B | D | C | D |
| 36 | E | A | B | A |
| 37 | B | D | E | E |
| 38 | B | A | B | D |
| 39 | C | C | B | A |
| 40 | D | D | D | B |
| 41 | E | B | E | D |
| 42 | D | B | B | A |
| 43 | E | D | D | B |
| 44 | D | B | B | B |
| 45 | D | B | D | E |
| 46 | C | B | D | E |
| 47 | D | E | C | B |
| 48 | D | B | D | B |
| 49 | B | B | B | C |
| 50 | A | C | B | C |
| 51 | C | A | C | C |
| 52 | A | D | C | B |
| 53 | D | C | E | B |
| 54 | C | E | E | A |
| 55 | E | B | B | E |
| 56 | D | A | E | B |
| 57 | D | D | D | A |
| 58 | E | C | C | B |
| 59 | A | C | B | E |
| 60 | B | A | B | E |

## APPENDIX 7: Factor analysis of NECO Mathematics test of year 2011

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 6.978 | 11.630 | 11.630 | 6.978 | 11.630 | 11.630 |
| 2 | 3.003 | 5.005 | 16.635 |  |  |  |
| 3 | 2.394 | 3.990 | 20.624 |  |  |  |
| 4 | 2.153 | 3.588 | 24.213 |  |  |  |
| 5 | 1.907 | 3.178 | 27.391 |  |  |  |
| 6 | 1.840 | 3.067 | 30.458 |  |  |  |
| 7 | 1.683 | 2.804 | 33.262 |  |  |  |
| 8 | 1.595 | 2.658 | 35.920 |  |  |  |
| 9 | 1.557 | 2.595 | 38.515 |  |  |  |
| 10 | 1.414 | 2.357 | 40.873 |  |  |  |
| 11 | 1.355 | 2.259 | 43.131 |  |  |  |
| 12 | 1.297 | 2.162 | 45.293 |  |  |  |
| 13 | 1.230 | 2.050 | 47.343 |  |  |  |
| 14 | 1.189 | 1.982 | 49.325 |  |  |  |
| 15 | 1.179 | 1.965 | 51.290 |  |  |  |
| 16 | 1.131 | 1.885 | 53.174 |  |  |  |
| 17 | 1.105 | 1.842 | 55.017 |  |  |  |
| 18 | 1.056 | 1.760 | 56.777 |  |  |  |
| 19 | 1.005 | 1.675 | 58.451 |  |  |  |
| 20 | . 982 | 1.637 | 60.089 |  |  |  |
| 21 | . 962 | 1.603 | 61.692 |  |  |  |
| 22 | . 928 | 1.547 | 63.239 |  |  |  |
| 23 | . 900 | 1.500 | 64.739 |  |  |  |
| 24 | . 882 | 1.470 | 66.209 |  |  |  |
| 25 | . 866 | 1.443 | 67.653 |  |  |  |
| 26 | . 853 | 1.422 | 69.075 |  |  |  |
| 27 | . 817 | 1.362 | 70.436 |  |  |  |
| 28 | . 796 | 1.327 | 71.763 |  |  |  |
| 29 | . 778 | 1.297 | 73.060 |  |  |  |
| 30 | . 771 | 1.285 | 74.346 |  |  |  |
| 31 | . 726 | 1.210 | 75.556 |  |  |  |
| 32 | . 696 | 1.161 | 76.716 |  |  |  |
| 33 | . 690 | 1.150 | 77.866 |  |  |  |
| 34 | . 684 | 1.140 | 79.006 |  |  |  |



Extraction Method: Principal Component Analysis.

Factor analysis of NECO Mathematics test of year 2012

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 11.355 | 18.925 | 18.925 | 11.355 | 18.925 | 18.925 |
| 2 | 3.155 | 5.259 | 24.183 |  |  |  |
| 3 | 2.649 | 4.415 | 28.598 |  |  |  |
| 4 | 2.452 | 4.086 | 32.684 |  |  |  |
| 5 | 2.091 | 3.485 | 36.170 |  |  |  |
| 6 | 1.847 | 3.078 | 39.248 |  |  |  |
| 7 | 1.659 | 2.765 | 42.013 |  |  |  |
| 8 | 1.457 | 2.429 | 44.441 |  |  |  |
| 9 | 1.413 | 2.354 | 46.796 |  |  |  |
| 10 | 1.331 | 2.219 | 49.014 |  |  |  |
| 11 | 1.296 | 2.159 | 51.174 |  |  |  |
| 12 | 1.253 | 2.089 | 53.262 |  |  |  |
| 13 | 1.228 | 2.046 | 55.308 |  |  |  |
| 14 | 1.135 | 1.892 | 57.200 |  |  |  |
| 15 | 1.099 | 1.832 | 59.032 |  |  |  |
| 16 | . 957 | 1.596 | 60.628 |  |  |  |
| 17 | . 938 | 1.563 | 62.190 |  |  |  |
| 18 | . 915 | 1.525 | 63.715 |  |  |  |
| 19 | . 898 | 1.497 | 65.212 |  |  |  |
| 20 | . 877 | 1.462 | 66.674 |  |  |  |
| 21 | . 835 | 1.392 | 68.066 |  |  |  |
| 22 | . 817 | 1.362 | 69.428 |  |  |  |
| 23 | . 778 | 1.296 | 70.724 |  |  |  |
| 24 | . 762 | 1.270 | 71.994 |  |  |  |
| 25 | . 715 | 1.192 | 73.186 |  |  |  |
| 26 | . 693 | 1.155 | 74.341 |  |  |  |
| 27 | . 685 | 1.142 | 75.483 |  |  |  |
| 28 | . 674 | 1.124 | 76.606 |  |  |  |
| 29 | . 656 | 1.093 | 77.700 |  |  |  |
| 30 | . 619 | 1.031 | 78.731 |  |  |  |
| 31 | . 606 | 1.010 | 79.741 |  |  |  |
| 32 | . 594 | . 990 | 80.731 |  |  |  |
| 33 | . 569 | . 948 | 81.680 |  |  |  |
| 34 | . 566 | . 944 | 82.623 |  |  |  |



Extraction Method: Principal Component Analysis.

Factor analysis of NECO Mathematics test of year 2013
Total Variance Explained

| Component | Initial Eigenvalues |  |  |  | Extraction Sums of Squared Loadings |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |  |
| 1 | 6.093 | 10.155 | 10.155 | 6.093 | 10.155 | 10.155 |  |
| 2 | 3.297 | 5.495 | 15.650 |  |  |  |  |
| 3 | 2.686 | 4.476 | 20.126 |  |  |  |  |
| 4 | 2.255 | 3.758 | 23.884 |  |  |  |  |
| 5 | 1.896 | 3.160 | 27.044 |  |  |  |  |
| 6 | 1.793 | 2.988 | 30.032 |  |  |  |  |
| 7 | 1.685 | 2.809 | 32.840 |  |  |  |  |
| 8 | 1.639 | 2.732 | 35.572 |  |  |  |  |




Extraction Method: Principal Component Analysis.
Factor analysis of NECO Mathematics test of year 2014

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 13.452 | 22.420 | 22.420 | 13.452 | 22.420 | 22.420 |
| 2 | 3.914 | 6.524 | 28.944 |  |  |  |
| 3 | 3.136 | 5.226 | 34.170 |  |  |  |
| 4 | 2.568 | 4.281 | 38.451 |  |  |  |
| 5 | 2.083 | 3.472 | 41.923 |  |  |  |
| 6 | 1.968 | 3.280 | 45.202 |  |  |  |
| 7 | 1.547 | 2.578 | 47.781 |  |  |  |
| 8 | 1.484 | 2.474 | 50.255 |  |  |  |
| 9 | 1.290 | 2.151 | 52.405 |  |  |  |
| 10 | 1.254 | 2.090 | 54.495 |  |  |  |
| 11 | 1.212 | 2.019 | 56.515 |  |  |  |
| 12 | 1.200 | 2.001 | 58.515 |  |  |  |
| 13 | 1.069 | 1.781 | 60.297 |  |  |  |
| 14 | . 996 | 1.660 | 61.956 |  |  |  |
| 15 | . 984 | 1.641 | 63.597 |  |  |  |
| 16 | . 916 | 1.527 | 65.125 |  |  |  |
| 17 | . 850 | 1.416 | 66.541 |  |  |  |
| 18 | . 831 | 1.385 | 67.926 |  |  |  |
| 19 | . 799 | 1.331 | 69.257 |  |  |  |
| 20 | . 769 | 1.282 | 70.540 |  |  |  |
| 21 | . 758 | 1.264 | 71.804 |  |  |  |



Extraction Method: Principal Component Analysis.

Factor analysis of WAEC Mathematics test of year 2011

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 11.914 | 23.828 | 23.828 | 11.914 | 23.828 | 23.828 |
| 2 | 2.651 | 5.302 | 29.131 |  |  |  |
| 3 | 2.345 | 4.691 | 33.822 |  |  |  |
| 4 | 1.975 | 3.950 | 37.771 |  |  |  |
| 5 | 1.638 | 3.276 | 41.047 |  |  |  |
| 6 | 1.503 | 3.006 | 44.054 |  |  |  |
| 7 | 1.414 | 2.828 | 46.882 |  |  |  |
| 8 | 1.376 | 2.752 | 49.634 |  |  |  |
| 9 | 1.272 | 2.543 | 52.177 |  |  |  |
| 10 | 1.162 | 2.324 | 54.501 |  |  |  |
| 11 | 1.123 | 2.246 | 56.748 |  |  |  |
| 12 | 1.033 | 2.066 | 58.814 |  |  |  |
| 13 | . 988 | 1.977 | 60.790 |  |  |  |
| 14 | . 895 | 1.789 | 62.579 |  |  |  |
| 15 | . 858 | 1.716 | 64.296 |  |  |  |
| 16 | . 847 | 1.693 | 65.989 |  |  |  |
| 17 | . 821 | 1.642 | 67.631 |  |  |  |
| 18 | . 798 | 1.596 | 69.227 |  |  |  |
| 19 | . 788 | 1.576 | 70.803 |  |  |  |
| 20 | . 744 | 1.488 | 72.291 |  |  |  |
| 21 | . 692 | 1.383 | 73.674 |  |  |  |
| 22 | . 682 | 1.364 | 75.038 |  |  |  |
| 23 | . 664 | 1.328 | 76.366 |  |  |  |
| 24 | . 646 | 1.292 | 77.659 |  |  |  |
| 25 | . 623 | 1.246 | 78.905 |  |  |  |
| 26 | . 608 | 1.216 | 80.121 |  |  |  |
| 27 | . 588 | 1.175 | 81.296 |  |  |  |
| 28 | . 577 | 1.154 | 82.450 |  |  |  |
| 29 | . 554 | 1.108 | 83.558 |  |  |  |
| 30 | . 533 | 1.066 | 84.624 |  |  |  |
| 31 | . 527 | 1.053 | 85.677 |  |  |  |
| 32 | . 502 | 1.004 | 86.681 |  |  |  |
| 33 | . 468 | . 936 | 87.617 |  |  |  |
| 34 | . 460 | . 921 | 88.538 |  |  |  |


| 35 | .448 | .896 | 89.434 |
| :--- | :--- | :--- | :--- |
| 36 | .446 | .893 | 90.327 |
| 37 | .422 | .844 | 91.171 |
| 38 | .416 | .831 | 92.002 |
| 39 | .408 | .816 | 92.819 |
| 40 | .389 | .779 | 93.598 |
| 41 | .375 | .749 | 94.347 |
| 42 | .354 | 95.056 |  |
| 43 | .347 | .695 | 95.750 |
| 44 | .338 | .676 | 96.426 |
| 45 | .321 | .641 | 97.068 |
| 46 | .317 | .633 | 97.701 |
| 47 | .308 | .616 | 98.317 |
| 48 | .293 | .585 |  |
| 49 | .279 | 989 |  |
| 50 |  |  |  |

Extraction Method: Principal Component Analysis.

Factor analysis of WAEC Mathematics test of year 2012

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 10.307 | 20.615 | 20.615 | 10.307 | 20.615 | 20.615 |
| 2 | 2.981 | 5.962 | 26.577 |  |  |  |
| 3 | 2.706 | 5.413 | 31.990 |  |  |  |
| 4 | 2.291 | 4.582 | 36.572 |  |  |  |
| 5 | 1.993 | 3.985 | 40.557 |  |  |  |
| 6 | 1.655 | 3.311 | 43.868 |  |  |  |
| 7 | 1.393 | 2.785 | 46.653 |  |  |  |
| 8 | 1.327 | 2.654 | 49.307 |  |  |  |
| 9 | 1.212 | 2.424 | 51.731 |  |  |  |
| 10 | 1.185 | 2.369 | 54.100 |  |  |  |
| 11 | 1.096 | 2.191 | 56.291 |  |  |  |
| 12 | 1.066 | 2.132 | 58.424 |  |  |  |
| 13 | . 948 | 1.896 | 60.320 |  |  |  |
| 14 | . 936 | 1.872 | 62.192 |  |  |  |
| 15 | . 887 | 1.775 | 63.967 |  |  |  |
| 16 | . 835 | 1.670 | 65.637 |  |  |  |
| 17 | . 789 | 1.577 | 67.214 |  |  |  |



Extraction Method: Principal Component Analysis.

Factor analysis of WAEC Mathematics test of year 2013

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | $\%$ of Variance | Cumulative $\%$ | Total | $\%$ of Variance | Cumulative $\%$ |
|  |  |  |  |  |  |  |


| 1 | 10.289 | 20.577 | 20.577 | 10.289 | 20.577 | 20.577 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 3.125 | 6.249 | 26.827 |  |  |  |
| 3 | 2.816 | 5.632 | 32.459 |  |  |  |
| 4 | 2.327 | 4.654 | 37.112 |  |  |  |
| 5 | 1.843 | 3.686 | 40.798 |  |  |  |
| 6 | 1.451 | 2.901 | 43.700 |  |  |  |
| 7 | 1.351 | 2.702 | 46.402 |  |  |  |
| 8 | 1.303 | 2.606 | 49.007 |  |  |  |
| 9 | 1.242 | 2.484 | 51.492 |  |  |  |
| 10 | 1.180 | 2.360 | 53.852 |  |  |  |
| 11 | 1.085 | 2.170 | 56.022 |  |  |  |
| 12 | 1.047 | 2.093 | 58.116 |  |  |  |
| 13 | 1.001 | 2.001 | 60.117 |  |  |  |
| 14 | . 957 | 1.915 | 62.032 |  |  |  |
| 15 | . 919 | 1.838 | 63.869 |  |  |  |
| 16 | . 876 | 1.752 | 65.621 |  |  |  |
| 17 | . 839 | 1.678 | 67.298 |  |  |  |
| 18 | . 800 | 1.600 | 68.899 |  |  |  |
| 19 | . 777 | 1.553 | 70.452 |  |  |  |
| 20 | . 744 | 1.488 | 71.941 |  |  |  |
| 21 | . 716 | 1.432 | 73.372 |  |  |  |
| 22 | . 686 | 1.372 | 74.744 |  |  |  |
| 23 | . 660 | 1.319 | 76.063 |  |  |  |
| 24 | . 637 | 1.274 | 77.337 |  |  |  |
| 25 | . 630 | 1.260 | 78.598 |  |  |  |
| 26 | . 609 | 1.217 | 79.815 |  |  |  |
| 27 | . 599 | 1.197 | 81.012 |  |  |  |
| 28 | . 572 | 1.144 | 82.156 |  |  |  |
| 29 | . 559 | 1.118 | 83.274 |  |  |  |
| 30 | . 550 | 1.100 | 84.374 |  |  |  |
| 31 | . 534 | 1.068 | 85.441 |  |  |  |
| 32 | . 519 | 1.037 | 86.479 |  |  |  |
| 33 | . 491 | . 981 | 87.460 |  |  |  |
| 34 | . 471 | . 943 | 88.403 |  |  |  |
| 35 | . 454 | . 908 | 89.311 |  |  |  |
| 36 | . 453 | . 906 | 90.216 |  |  |  |
| 37 | . 429 | . 859 | 91.075 |  |  |  |
| 38 | . 421 | . 842 | 91.916 |  |  |  |
| 39 | . 401 | . 802 | 92.719 |  |  |  |
| 40 | . 385 | . 771 | 93.489 |  |  |  |


| 41 | .382 |  |  |
| :--- | ---: | ---: | ---: |
| 42 | .367 | .765 | 94.254 |
| 43 | .734 | 94.988 |  |
| 44 | .365 | .729 | 95.718 |
| 45 | .683 | 96.400 |  |
| 46 | .337 | .675 |  |
| 46 | .328 | .657 |  |
| 47 | .316 | .631 | 98.732 |
| 48 | .287 | .574 | 98.937 |
| 49 | .268 | .536 |  |
| 50 | .264 |  |  |

Extraction Method: Principal Component Analysis.

Factor analysis of WAEC Mathematics test of year 2014

Total Variance Explained

| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 10.415 | 20.830 | 20.830 | 10.415 | 20.830 | 20.830 |
| 2 | 2.911 | 5.821 | 26.652 |  |  |  |
| 3 | 2.532 | 5.065 | 31.717 |  |  |  |
| 4 | 1.981 | 3.962 | 35.678 |  |  |  |
| 5 | 1.812 | 3.623 | 39.302 |  |  |  |
| 6 | 1.609 | 3.218 | 42.519 |  |  |  |
| 7 | 1.550 | 3.100 | 45.619 |  |  |  |
| 8 | 1.407 | 2.814 | 48.433 |  |  |  |
| 9 | 1.389 | 2.778 | 51.211 |  |  |  |
| 10 | 1.258 | 2.517 | 53.728 |  |  |  |
| 11 | 1.214 | 2.428 | 56.156 |  |  |  |
| 12 | 1.127 | 2.254 | 58.410 |  |  |  |
| 13 | 1.025 | 2.051 | 60.461 |  |  |  |
| 14 | . 962 | 1.925 | 62.386 |  |  |  |
| 15 | . 941 | 1.881 | 64.267 |  |  |  |
| 16 | . 877 | 1.753 | 66.020 |  |  |  |
| 17 | . 857 | 1.713 | 67.734 |  |  |  |
| 18 | . 829 | 1.659 | 69.392 |  |  |  |
| 19 | . 809 | 1.618 | 71.010 |  |  |  |
| 20 | . 746 | 1.493 | 72.503 |  |  |  |
| 21 | . 724 | 1.447 | 73.951 |  |  |  |
| 22 | . 685 | 1.369 | 75.320 |  |  |  |



Extraction Method: Principal Component Analysis.

## Appendix 8: distribution of abilities scores of examinee on NECO and WAEC Mathematics test

| Year |  | 2011 |  |  | 2012 |  |  |  |  | 2013 |  |  | $\operatorname{Tr} x_{N}$ |  | 2014 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| S/N | NECO | $\operatorname{Tr} x_{N}$ | WAEC | $\operatorname{Tr} x_{W}$ | NECO | $\operatorname{Tr} x_{N}$ | WAEC | $\operatorname{Tr} x_{W}$ | NECO | $\operatorname{Tr} x_{N}$ | WAEC | $\operatorname{Tr} x_{W}$ | NECO |  | WAEC | $\operatorname{Tr} x_{W}$ |
| 1 | 0.66 | -0.46 | -0.46 | 0.71 | 0.43 | 0.97 | 0.9 | 0.38 | 0.59 | 0.41 | 0.46 | 0.64 | 0.91 | 0.45 | 0.43 | 1.93 |
| 2 | 0.89 | -0.91 | -0.93 | 0.95 | 0.58 | 1.15 | 1.08 | 0.52 | 0.64 | 0.39 | 0.43 | 0.7 | 1.1 | 1.63 | 1.68 | 2.12 |
| 3 | -0.41 | 0.23 | 0.27 | -0.42 | 0.94 | 0.65 | 0.59 | 0.87 | 0.56 | 0.55 | 0.62 | 0.61 | 0.5 | 1.26 | 1.29 | 1.52 |
| 4 | 0.38 | 0.26 | 0.31 | 0.41 | 0.38 | 0.87 | 0.81 | 0.33 | 0.91 | 0.38 | 0.42 | 1.01 | 0.86 | 1.14 | 1.16 | 1.88 |
| 5 | 0.28 | 0.18 | 0.22 | 0.31 | 0.77 | -0.05 | -0.09 | 0.71 | 0.76 | 0.06 | 0.06 | 0.84 | 0.87 | 0.41 | 0.39 | 1.89 |
| 6 | 0.38 | -0.06 | -0.03 | 0.41 | 0.65 | 0.98 | 0.91 | 0.59 | 0.97 | 0.56 | 0.63 | 1.08 | -1.56 | 0.03 | -0.01 | -0.54 |
| 7 | 0.99 | 0.27 | 0.32 | 1.06 | 0.63 | 0.98 | 0.91 | 0.57 | 0.6 | 0.32 | 0.35 | 0.66 | 1.16 | 1 | 1.01 | 2.18 |
| 8 | 0.62 | 0.63 | 0.7 | 0.67 | 0.75 | 1.12 | 1.05 | 0.69 | 0.45 | 0.52 | 0.58 | 0.48 | 1.47 | 1.5 | 1.54 | 2.49 |
| 9 | -0.13 | 0.26 | 0.31 | -0.13 | 0.46 | -0.36 | -0.39 | 0.41 | -1.25 | 0.6 | 0.68 | -1.46 | 0.89 | 0.98 | 0.99 | 1.91 |
| 10 | -0.13 | 0.35 | 0.4 | -0.13 | 0.78 | 1.03 | 0.96 | 0.72 | 0.53 | 0.5 | 0.56 | 0.58 | -1.42 | 1.5 | 1.54 | -0.4 |
| 11 | 0.66 | 0.56 | 0.62 | 0.71 | 0.32 | 1.06 | 0.99 | 0.27 | -1.29 | 0.52 | 0.58 | -1.51 | 0.76 | 0.44 | 0.42 | 1.78 |
| 12 | 0.68 | 1.27 | 1.38 | 0.73 | 1.16 | 0.69 | 0.63 | 1.09 | 0.27 | 1.22 | 1.39 | 0.28 | -1.74 | 1.56 | 1.61 | -0.72 |
| 13 | 0.68 | 1.27 | 1.38 | 0.73 | 1.14 | 0.69 | 0.63 | 1.07 | 0.27 | 0.83 | 0.94 | 0.28 | -1.38 | 1.38 | 1.42 | -0.36 |
| 14 | 0.68 | 0.98 | 1.07 | 0.73 | 1.22 | 0.72 | 0.66 | 1.15 | 0.33 | 1.18 | 1.34 | 0.35 | -1.17 | 1.18 | 1.2 | -0.15 |
| 15 | 0.74 | 1.27 | 1.38 | 0.79 | 1.16 | 0.69 | 0.63 | 1.09 | 0.27 | 1.22 | 1.39 | 0.28 | -1.61 | 1.56 | 1.61 | -0.59 |
| 16 | 0.68 | 1.27 | 1.38 | 0.73 | 1.22 | 0.69 | 0.63 | 1.15 | 0.27 | 1.22 | 1.39 | 0.28 | -0.75 | 1.07 | 1.09 | 0.27 |
| 17 | 0.81 | 1.27 | 1.38 | 0.87 | 0.37 | 0.78 | 0.72 | 0.32 | -0.66 | 0.89 | 1.01 | -0.79 | -1.56 | 0.93 | 0.94 | -0.54 |
| 18 | -0.91 | 1.27 | 1.38 | -0.95 | -1.96 | 0.65 | 0.59 | -1.95 | 0.99 | 0.85 | 0.96 | 1.1 | 1.08 | 0.86 | 0.87 | 2.1 |
| 19 | 0.68 | 0.82 | 0.9 | 0.73 | 1.22 | 0.69 | 0.63 | 1.15 | 0.27 | -0.72 | -0.84 | 0.28 | -1.1 | 1.19 | 1.21 | -0.08 |
| 20 | 0.68 | 1.14 | 1.24 | 0.73 | 1.22 | 0.63 | 0.57 | 1.15 | 0.27 | -0.75 | -0.87 | 0.28 | -1.34 | 0.78 | 0.78 | -0.32 |


| 21 | 0.68 | 1.35 | 1.46 | 0.73 | 1.22 | 0.69 | 0.63 | 1.15 | 0.2 | 1.22 | 1.39 | 0.2 | -1.56 | 1.07 | 1.09 | -0.54 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 22 | 0.68 | 1.27 | 1.38 | 0.73 | 1.16 | 0.69 | 0.63 | 1.09 | 0.27 | 1.22 | 1.39 | 0.28 | -1.56 | 1.32 | 1.35 | -0.54 |
| 23 | 0.78 | -0.87 | -0.89 | 0.84 | -1.9 | -0.24 | -0.27 | -1.89 | -1.3 | 0.03 | 0.02 | -1.52 | -1.53 | 1.44 | 1.48 | -0.51 |
| 24 | 1.14 | -0.87 | -0.89 | 1.22 | -2.34 | -0.13 | -0.17 | -2.31 | -1.3 | 0.03 | 0.02 | -1.52 | -1.53 | 1.44 | 1.48 | -0.51 |
|  | 1.71 | -0.74 | -0.75 | 1.82 | -1.91 | -0.04 | -0.08 | -1.9 | -1.3 | 0.09 | 0.09 | -1.52 | -1.44 | 0.83 | 0.83 | -0.42 |
|  | 1.14 | -0.87 | -0.89 | 1.22 | -1.91 | -0.13 | -0.17 | -1.9 | -1.3 | 0.04 | 0.04 | -1.52 | -1.44 | 1.44 | 1.48 | -0.42 |
|  | 1.14 | -0.87 | -0.89 | 1.22 | -0.28 | -0.26 | -0.29 | -0.31 | -1.28 | -0.01 | -0.02 | -1.49 | -1.53 | 1.44 | 1.48 | -0.51 |
|  | 1.14 | -0.87 | -0.89 | 1.22 | -1.88 | -0.08 | -0.12 | -1.87 | -1.3 | 0.03 | 0.02 | -1.52 | -1.56 | 1.44 | 1.48 | -0.54 |
|  | 1.66 | -0.98 | -1.01 | 1.77 | -2.05 | -0.27 | -0.3 | -2.03 | -1.22 | 0.21 | 0.23 | -1.43 | -1.17 | 1.56 | 1.61 | -0.15 |
|  | 1.71 | -0.98 | -1.01 | 1.82 | -1.96 | -0.01 | -0.05 | -1.95 | -1.3 | 0.09 | 0.09 | -1.52 | -1.37 | -1.03 | -1.13 | -0.35 |
|  | 1.34 | 2.09 | 2.24 | 1.43 | -1.96 | -0.04 | -0.08 | -1.95 | -1.3 | 0.07 | 0.07 | -1.52 | -1.39 | 0.96 | 0.97 | -0.37 |
|  | 0.52 | -1.11 | -1.14 | 0.56 | -1.76 | -0.16 | -0.19 | -1.75 | -1.15 | 0.23 | 0.25 | -1.35 | -1.53 | 0.64 | 0.63 | -0.51 |
|  | 1.48 | -0.98 | -1.01 | 1.58 | -1.96 | -0.16 | -0.19 | -1.95 | -1.35 | 0.18 | 0.19 | -1.57 | -1.53 | 0.64 | 0.63 | -0.51 |
|  | -0.03 | -0.87 | -0.89 | -0.02 | -0.87 | -0.16 | -0.19 | -0.89 | -1.25 | 0.03 | 0.02 | -1.46 | -1.17 | -0.68 | -0.76 | -0.15 |
|  | 0.44 | -0.8 | -0.82 | 0.48 | -1.96 | -0.3 | -0.33 | -1.95 | -1.3 | 0.16 | 0.17 | -1.52 | -1.24 | -1.16 | -1.27 | -0.22 |
|  | 0.82 | -0.63 | -0.64 | 0.88 | -1.98 | -0.13 | -0.17 | -1.96 | -1.25 | -0.43 | -0.5 | -1.46 | -0.39 | -1.18 | -1.29 | 0.63 |
|  | -0.69 | -0.82 | -0.84 | -0.72 | -0.24 | 1.21 | 1.14 | -0.27 | -1.25 | 0.75 | 0.85 | -1.46 | -1.53 | 1.14 | 1.16 | -0.51 |
|  | -1.2 | -1.54 | -1.6 | -1.26 | -0.78 | 0.85 | 0.79 | -0.8 | -0.69 | 1.22 | 1.39 | -0.82 | -1.36 | 1.15 | 1.17 | -0.34 |
|  | -0.55 | -0.98 | -1.01 | -0.57 | -1.13 | 0.03 | -0.01 | -1.14 | 1.08 | -0.2 | -0.24 | 1.21 | -1.11 | 0.79 | 0.79 | -0.09 |
|  | 1.14 | -0.87 | -0.89 | 1.22 | -2.08 | -0.04 | -0.08 | -2.06 | -1.3 | 0.03 | 0.02 | -1.52 | -1.53 | 1.44 | 1.48 | -0.51 |
|  | 1.17 | -0.94 | -0.96 | 1.25 | -2.08 | 0.01 | -0.03 | -2.06 | -1.35 | -0.03 | -0.05 | -1.57 | -1.53 | 1.44 | 1.48 | -0.51 |
|  | 0.92 | -1.01 | -1.04 | 0.98 | -2.08 | -0.04 | -0.08 | -2.06 | -1.33 | -0.02 | -0.04 | -1.55 | -1.57 | 1.44 | 1.48 | -0.55 |
|  | 0.83 | -0.98 | -1.01 | 0.89 | -2.02 | -0.27 | -0.3 | -2 | -1.3 | 0.18 | 0.19 | -1.52 | -1.29 | -1.29 | -1.4 | -0.27 |
|  | 0.04 | -0.91 | -0.93 | 0.05 | -2.14 | -0.16 | -0.19 | -2.12 | -1.08 | -0.07 | -0.09 | -1.27 | -1.44 | 1.38 | 1.41 | -0.42 |
|  | 0.78 | 0.39 | 0.44 | 0.84 | 0.74 | 0.87 | 0.81 | 0.68 | 1.09 | 0.65 | 0.73 | 1.22 | -1.53 | 0.51 | 0.5 | -0.51 |


| 0.82 | 0.95 | 1.04 | 0.88 | 0.5 | 0.63 | 0.57 | 0.45 | 0.74 | 0.5 | 0.56 | 0.82 | -1.53 | 0.95 | 0.96 | -0.51 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.44 | 1.19 | 1.29 | 0.48 | 0.96 | 0.72 | 0.66 | 0.89 | 1.11 | 0.66 | 0.74 | 1.24 | -1.53 | 1.2 | 1.23 | -0.51 |
| 0.17 | 1.39 | 1.5 | 0.19 | 1 | 0.69 | 0.63 | 0.93 | 0.82 | 0.72 | 0.81 | 0.91 | -1.48 | 0.36 | 0.34 | -0.46 |
| 0.36 | 1.16 | 1.26 | 0.39 | 0.93 | 0.61 | 0.55 | 0.86 | 0.71 | 0.88 | 1 | 0.78 | -1.32 | 1 | 1.01 | -0.3 |
| 0.93 | -0.98 | -1.01 | 0.99 | -1.61 | 1.47 | 1.39 | -1.6 | -1.3 | -0.53 | -0.62 | -1.52 | -1.09 | -1.15 | -1.26 | -0.07 |
| 0.46 | -0.98 | -1.01 | 0.5 | -1.8 | -0.3 | -0.33 | -1.79 | -1.3 | 0.06 | 0.06 | -1.52 | -1.3 | 1.23 | 1.26 | -0.28 |
| 1.71 | -0.87 | -0.89 | 1.82 | -2.01 | -0.3 | -0.33 | -1.99 | -1.3 | 0.09 | 0.09 | -1.52 | -1.45 | 0.84 | 0.84 | -0.43 |
| -0.35 | -0.93 | -0.95 | -0.36 | -1.05 | -0.28 | -0.31 | -1.06 | -1.31 | -0.11 | -0.14 | -1.53 | -1.03 | -1.15 | -1.26 | -0.01 |
| 0.74 | -0.93 | -0.95 | 0.79 | -2.01 | -0.16 | -0.19 | -1.99 | -1.36 | -0.08 | -0.1 | -1.59 | -1.04 | 1.16 | 1.18 | -0.02 |
| 0.94 | -0.87 | -0.89 | 1 | -2.03 | -0.2 | -0.23 | -2.01 | -1.3 | 0.04 | 0.04 | -1.52 | -1.04 | 1.2 | 1.23 | -0.02 |
| 0.4 | -0.87 | -0.89 | 0.43 | -1.72 | -0.28 | -0.31 | -1.71 | -0.68 | 0.08 | 0.08 | -0.81 | -1.4 | 1.41 | 1.45 | -0.38 |
| 1.26 | 0.56 | 0.62 | 1.34 | 0.71 | 0.94 | 0.87 | 0.65 | 1.39 | 0.66 | 0.74 | 1.56 | 0.72 | 1.15 | 1.17 | 1.74 |
| -0.15 | -0.06 | -0.03 | -0.15 | 0.9 | 0.65 | 0.59 | 0.83 | 0.89 | 0.52 | 0.58 | 0.99 | -1.23 | 1.13 | 1.15 | -0.21 |
| -0.15 | -0.98 | -1.01 | -0.15 | -1.96 | -0.39 | -0.42 | -1.95 | -1.22 | 0.09 | 0.09 | -1.43 | -1.45 | 0.36 | 0.34 | -0.43 |
| 1.61 | -0.98 | -1.01 | 1.71 | -2.01 | -0.39 | -0.42 | -1.99 | -1.3 | -0.12 | -0.15 | -1.52 | -1.17 | 1.2 | 1.23 | -0.15 |
| 0.04 | -0.76 | -0.77 | 0.05 | -2.46 | -0.13 | -0.17 | -2.43 | -1.3 | 0.03 | 0.02 | -1.52 | -1.58 | 1.44 | 1.48 | -0.56 |
| 1.24 | -0.57 | -0.57 | 1.32 | -1.71 | -1.39 | -1.39 | -1.7 | -1.3 | 0.18 | 0.19 | -1.52 | -1.53 | 1.45 | 1.49 | -0.51 |
| -0.04 | 3.75 | 4 | -0.03 | 0.81 | 0.44 | 0.39 | 0.75 | 0.66 | 0.41 | 0.46 | 0.73 | 0.87 | 0.44 | 0.42 | 1.89 |
| 0.46 | 3.75 | 4 | 0.5 | 0.48 | 0.9 | 0.84 | 0.43 | 0.61 | 0.25 | 0.27 | 0.67 | 0.83 | 0.01 | -0.03 | 1.85 |
| 0.72 | 0.06 | 0.1 | 0.77 | 0.67 | 0.4 | 0.35 | 0.61 | 0.72 | 0.46 | 0.52 | 0.79 | 0.84 | 0.13 | 0.09 | 1.86 |
| 0.12 | 0.38 | 0.43 | 0.14 | 1.14 | -0.43 | -0.46 | 1.07 | 0.59 | 1.01 | 1.14 | 0.64 | 0.65 | 1.54 | 1.58 | 1.67 |
| 0.92 | -0.05 | -0.02 | 0.98 | -0.11 | 0.55 | 0.5 | -0.15 | 0.6 | -0.15 | -0.18 | 0.66 | 0.17 | 0.21 | 0.18 | 1.19 |
| 1.21 | -0.29 | -0.27 | 1.29 | 0.16 | -1.35 | -1.35 | 0.12 | 0.82 | 0.36 | 0.4 | 0.91 | 0.69 | 0.02 | -0.02 | 1.71 |
| -0.54 | -0.59 | -0.59 | -0.56 | 4 | 1.27 | 1.2 | 3.85 | -0.06 | 0.5 | 0.56 | -0.1 | 1.07 | -0.11 | -0.16 | 2.09 |
| 1.03 | 0.13 | 0.17 | 1.1 | 2.57 | 0.7 | 0.64 | 2.46 | 1.01 | 0.7 | 0.79 | 1.13 | 1.01 | 0.62 | 0.61 | 2.03 |


| -1.23 | 0.1 | 0.14 | -1.29 | 2.57 | 0.44 | 0.39 | 2.46 | 0.46 | 0.59 | 0.66 | 0.5 | 0.54 | 0.85 | 0.86 | 1.56 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.4 | 1.05 | 1.14 | 0.43 | 0.36 | 0.89 | 0.83 | 0.31 | 0.96 | 0.87 | 0.98 | 1.07 | 1.04 | 0.99 | 1 | 2.06 |
| -0.45 | -0.42 | -0.41 | -0.47 | 0.25 | 2.06 | 1.96 | 0.2 | 0.22 | 0.67 | 0.76 | 0.22 | 1.2 | -0.36 | -0.42 | 2.22 |
| 0.13 | -0.1 | -0.07 | 0.15 | 2.28 | 0.87 | 0.81 | 2.18 | 0.85 | 0.44 | 0.49 | 0.94 | 0.37 | 0.77 | 0.77 | 1.39 |
| 0.47 | 0.19 | 0.23 | 0.51 | 2.57 | 0.7 | 0.64 | 2.46 | -0.71 | 0.26 | 0.29 | -0.84 | 0.88 | 0.74 | 0.74 | 1.9 |
| -0.35 | -0.01 | 0.02 | -0.36 | 4 | 0.53 | 0.48 | 3.85 | 0.25 | 0.14 | 0.15 | 0.26 | 0.17 | 0.83 | 0.83 | 1.19 |
| -0.51 | 0.07 | 0.11 | -0.53 | 0.25 | 0.77 | 0.71 | 0.2 | 0.64 | 3.22 | 3.67 | 0.7 | 0.92 | 0.86 | 0.87 | 1.94 |
| -0.51 | 0.1 | 0.14 | -0.53 | 0.4 | 0.8 | 0.74 | 0.35 | 0.97 | 1.96 | 2.23 | 1.08 | 0.69 | 0.27 | 0.24 | 1.71 |
| -1.13 | -0.53 | -0.53 | -1.19 | -0.56 | -2.95 | -2.91 | -0.58 | -0.21 | -1.13 | -1.31 | -0.27 | -0.1 | -0.85 | -0.94 | 0.92 |
| 0.27 | -0.03 | 0 | 0.3 | 0.59 | 0.63 | 0.57 | 0.53 | -1.85 | 0.6 | 0.68 | -2.15 | 0.4 | 1.97 | 2.04 | 1.42 |
| 0.29 | 0.46 | 0.52 | 0.32 | -0.12 | 0.08 | 0.04 | -0.16 | 0.05 | 0.3 | 0.33 | 0.03 | 0.88 | -1.2 | -1.31 | 1.9 |
| -0.07 | 0.26 | 0.31 | -0.06 | 0.55 | 0.96 | 0.89 | 0.49 | 0.89 | 1.03 | 1.17 | 0.99 | 1.44 | 0.72 | 0.72 | 2.46 |
| 2.14 | 0.2 | 0.24 | 2.27 | 0.83 | 0.58 | 0.52 | 0.77 | 0.63 | -0.44 | -0.51 | 0.69 | 0.75 | 0.39 | 0.37 | 1.77 |
| -0.17 | 0.15 | 0.19 | -0.17 | 0.52 | 0.28 | 0.23 | 0.47 | 0.04 | 0.39 | 0.43 | 0.02 | 0.12 | 0.58 | 0.57 | 1.14 |
| 0.09 | 0 | 0.03 | 0.11 | 0.35 | 0.6 | 0.54 | 0.3 | 0.58 | 0.47 | 0.53 | 0.63 | 0.93 | 3.23 | 3.37 | 1.95 |
| 2.14 | 0.26 | 0.31 | 2.27 | 0.47 | 0.59 | 0.53 | 0.42 | 0.9 | 0.25 | 0.27 | 1 | 1.08 | 0.22 | 0.19 | 2.1 |
| 0.53 | -0.1 | -0.07 | 0.57 | 0.07 | 2.06 | 1.96 | 0.03 | -0.18 | 0.74 | 0.83 | -0.24 | 0.7 | -0.94 | -1.03 | 1.72 |
| -0.38 | -0.2 | -0.18 | -0.39 | 0.23 | 2.51 | 2.4 | 0.18 | 0.07 | 0.67 | 0.76 | 0.05 | 0.63 | 0.34 | 0.32 | 1.65 |
| -0.51 | 0.07 | 0.11 | -0.53 | 0.25 | 0.77 | 0.71 | 0.2 | 0.61 | -0.01 | -0.02 | 0.67 | 0.46 | 0.85 | 0.86 | 1.48 |
| 2.14 | 0.22 | 0.26 | 2.27 | 0.37 | 0.08 | 0.04 | 0.32 | -1.4 | 0.18 | 0.2 | -1.63 | 0.47 | 0.22 | 0.19 | 1.49 |
| 0.17 | 0.66 | 0.73 | 0.19 | 0.66 | 0.67 | 0.61 | 0.6 | 0.38 | 1.36 | 1.54 | 0.4 | 0.25 | 0.24 | 0.21 | 1.27 |
| 0.77 | 0.74 | 0.82 | 0.82 | 0.7 | 0.56 | 0.51 | 0.64 | -1.9 | 0.96 | 1.09 | -2.2 | 1.48 | 0.35 | 0.33 | 2.5 |
| 0.12 | 0.47 | 0.53 | 0.14 | 0.67 | 0.62 | 0.56 | 0.61 | 0.62 | 0.88 | 1 | 0.68 | 1.48 | 0.9 | 0.91 | 2.5 |
| -0.26 | 0.13 | 0.17 | -0.27 | 0.73 | 0.47 | 0.42 | 0.67 | 0.31 | 0.77 | 0.87 | 0.32 | 1.35 | 0.83 | 0.83 | 2.37 |
| 0.7 | 0.54 | 0.6 | 0.75 | 0.76 | 0.91 | 0.85 | 0.7 | 1.24 | 0.76 | 0.86 | 1.39 | 1.37 | 0.71 | 0.71 | 2.39 |


| -0.3 | 0.15 | 0.19 | -0.31 | 0.32 | 0.44 | 0.39 | 0.27 | 0.58 | 0.46 | 0.51 | 0.63 | 1.8 | 0.5 | 0.49 | 2.82 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.14 | 0.25 | 0.3 | -0.14 | 0.54 | 0.41 | 0.36 | 0.48 | 0.67 | 0.46 | 0.51 | 0.74 | 1.05 | 3.23 | 3.37 | 2.07 |
| 0.71 | 0.25 | 0.3 | 0.76 | -0.1 | 0.42 | 0.37 | -0.14 | 0.68 | 3.22 | 3.67 | 0.75 | 0.59 | 0.48 | 0.46 | 1.61 |
| 2.14 | 0.57 | 0.64 | 2.27 | 0.29 | -0.13 | -0.17 | 0.24 | 0.61 | 0.04 | 0.03 | 0.67 | 0.63 | 0.28 | 0.25 | 1.65 |
| 0.28 | 0.05 | 0.08 | 0.31 | 0.45 | 0.3 | 0.25 | 0.4 | 0.58 | 0.29 | 0.32 | 0.63 | 0.8 | 3.23 | 3.37 | 1.82 |
| 0.22 | 0.07 | 0.11 | 0.24 | -0.25 | -0.43 | -0.46 | -0.28 | 0.69 | 0.56 | 0.63 | 0.76 | 0.55 | 0.22 | 0.19 | 1.57 |
| 1.02 | 0.62 | 0.69 | 1.09 | 0.31 | 0.84 | 0.78 | 0.26 | 1.86 | 0.76 | 0.86 | 2.1 | 0.84 | 0.95 | 0.96 | 1.86 |
| 0.68 | 1.27 | 1.38 | 0.73 | 1.22 | 0.69 | 0.63 | 1.15 | 1.86 | 1.22 | 1.39 | 2.1 | -1.53 | 0.97 | 0.98 | -0.51 |
| 1.12 | 3.75 | 4 | 1.19 | 0.47 | 0.7 | 0.64 | 0.42 | 0.87 | 0.87 | 0.98 | 0.97 | 1.02 | 0.5 | 0.49 | 2.04 |
| 0.79 | 1.17 | 1.27 | 0.85 | 1.07 | 1.15 | 1.08 | 1 | 1.04 | 0.83 | 0.94 | 1.16 | 1.02 | 1.03 | 1.05 | 2.04 |
| -0.99 | -1.07 | -1.1 | -1.04 | 0.69 | -1.6 | -1.59 | 0.63 | 0.15 | -1.12 | -1.29 | 0.14 | -0.7 | -3.75 | -4 | 0.32 |
| -1.27 | -1.42 | -1.47 | -1.33 | -0.48 | -2.53 | -2.5 | -0.51 | 0.42 | -0.65 | -0.76 | 0.45 | -0.15 | -2.55 | -2.74 | 0.87 |
| -0.32 | -1.96 | -2.04 | -0.33 | -1.86 | -1.19 | -1.2 | -1.85 | 0.32 | -2.13 | -2.45 | 0.34 | -0.23 | -0.54 | -0.61 | 0.79 |
| -0.05 | -1.39 | -1.44 | -0.04 | -1 | -1.51 | -1.51 | -1.01 | -0.33 | -1.28 | -1.48 | -0.41 | -0.2 | -1.2 | -1.31 | 0.82 |
| -1.06 | -1.44 | -1.49 | -1.11 | -0.7 | -1.37 | -1.37 | -0.72 | 0 | -0.93 | -1.08 | -0.03 | -0.13 | -1.17 | -1.28 | 0.89 |
| 2.24 | 1.95 | 2.09 | 2.38 | 4 | 1.34 | 1.26 | 3.85 | 1.86 | 0.78 | 0.88 | 2.1 | 1.2 | 0.61 | 0.6 | 2.22 |
| 1.08 | 1.05 | 1.14 | 1.15 | 0.48 | -0.2 | -0.23 | 0.43 | 1.86 | 0.6 | 0.68 | 2.1 | 1.17 | 0.43 | 0.41 | 2.19 |
| -0.56 | 1.4 | 1.51 | -0.58 | 1.79 | 0.1 | 0.06 | 1.7 | 1.88 | -0.86 | -1 | 2.12 | 1.67 | 0.74 | 0.74 | 2.69 |
| -0.96 | -0.96 | -0.98 | -1.01 | -0.54 | -1.25 | -1.25 | -0.56 | 0.64 | -0.09 | -0.12 | 0.7 | -0.92 | -0.22 | -0.27 | 0.1 |
| -0.56 | 1.27 | 1.38 | -0.58 | 1.05 | 0.1 | 0.06 | 0.98 | 1.88 | -0.86 | -1 | 2.12 | 1.57 | 0.74 | 0.74 | 2.59 |
| 0.32 | 3.75 | 4 | 0.35 | 0.22 | 0.74 | 0.68 | 0.17 | 0.17 | 0.8 | 0.9 | 0.16 | 0.6 | 0.33 | 0.31 | 1.62 |
| -0.59 | 3.75 | 4 | -0.61 | 0.39 | -1.07 | -1.08 | 0.34 | 0.65 | 0.77 | 0.87 | 0.71 | 0.98 | 0.68 | 0.68 | 2 |
| 0.03 | -0.06 | -0.03 | 0.04 | -0.01 | 2.63 | 2.52 | -0.05 | 0.6 | -0.04 | -0.06 | 0.66 | 0.47 | 0.59 | 0.58 | 1.49 |
| 0.17 | 0.08 | 0.12 | 0.19 | 0.29 | 2.93 | 2.81 | 0.24 | 1.07 | -1.09 | -1.26 | 1.19 | 0.74 | 0.13 | 0.1 | 1.76 |
| -0.19 | 0.52 | 0.58 | -0.19 | 0.33 | 2.93 | 2.81 | 0.28 | 0.7 | 0.37 | 0.41 | 0.77 | 0.13 | -0.13 | -0.18 | 1.15 |


| 0.03 | 2.38 | 2.55 | 0.04 | 0.17 | 0.08 | 0.04 | 0.13 | 0.66 | -0.37 | -0.43 | 0.73 | 0.49 | 0.11 | 0.07 | 1.51 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.12 | 2.69 | 2.88 | -0.12 | 0.24 | 0.49 | 0.44 | 0.19 | 0.48 | 0.44 | 0.49 | 0.52 | 0.84 | 0.38 | 0.36 | 1.86 |
| 0.56 | 0.57 | 0.64 | 0.6 | 0.44 | 0.56 | 0.51 | 0.39 | 1.86 | 0.28 | 0.31 | 2.1 | 0.89 | 0.36 | 0.34 | 1.91 |
| 0.26 | 3.18 | 3.4 | 0.29 | 0.42 | 1.05 | 0.98 | 0.37 | 0.71 | 0.11 | 0.12 | 0.78 | 0.35 | -0.29 | -0.35 | 1.37 |
| -0.2 | 3.75 | 4 | -0.2 | 0.76 | 0.8 | 0.74 | 0.7 | 0.89 | -0.12 | -0.15 | 0.99 | 0.81 | 0.06 | 0.02 | 1.83 |
| 0.47 | 3.75 | 4 | 0.51 | 0.72 | 0.52 | 0.47 | 0.66 | 0.5 | -0.07 | -0.09 | 0.54 | 0.49 | 0.12 | 0.08 | 1.51 |
| 0.92 | 0.92 | 1.01 | 0.98 | 1.25 | 0.67 | 0.61 | 1.18 | 1.11 | 0.95 | 1.08 | 1.24 | 1.17 | 0.86 | 0.87 | 2.19 |
| 2.14 | -0.03 | 0 | 2.27 | 0.81 | 0.46 | 0.41 | 0.75 | 0.49 | 0.22 | 0.24 | 0.53 | 1.01 | 0.06 | 0.02 | 2.03 |
| 2.14 | 0.36 | 0.41 | 2.27 | 0.3 | 0.69 | 0.63 | 0.25 | 1.01 | 0.32 | 0.36 | 1.13 | 0.45 | 0.15 | 0.12 | 1.47 |
| -0.16 | 0.28 | 0.33 | -0.16 | -0.8 | 0.66 | 0.6 | -0.82 | 0.19 | 0.8 | 0.9 | 0.19 | -0.56 | 0.26 | 0.23 | 0.46 |
| 0.56 | 0.44 | 0.5 | 0.6 | 2.18 | 0.92 | 0.86 | 2.08 | 1.87 | 0.88 | 0.99 | 2.11 | -0.34 | 1.55 | 1.59 | 0.68 |
| -1.85 | -0.97 | -0.99 | -1.95 | -0.56 | -1.3 | -1.3 | -0.58 | -0.96 | -1.13 | -1.3 | -1.13 | -0.91 | -1.83 | -1.97 | 0.11 |
| -1.58 | -0.51 | -0.51 | -1.66 | -0.67 | -1.54 | -1.54 | -0.69 | -1.53 | -1.82 | -2.09 | -1.78 | -0.41 | -1.3 | -1.41 | 0.61 |
| -1.21 | -0.92 | -0.94 | -1.27 | -0.11 | -2.51 | -2.48 | -0.15 | -1.74 | -1.19 | -1.37 | -2.02 | -0.53 | -2.07 | -2.23 | 0.49 |
| -1.9 | -0.52 | -0.52 | -2 | -0.81 | -0.47 | -0.5 | -0.83 | -1.58 | -1.13 | -1.31 | -1.84 | -0.28 | -0.81 | -0.9 | 0.74 |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 3.22 | 3.67 | 2.1 | 0.28 | 0.8 | 0.8 | 1.3 |
| 1.01 | 0.53 | 0.59 | 1.08 | 0.5 | 0.95 | 0.88 | 0.45 | 0.73 | 0.51 | 0.57 | 0.81 | 0.56 | 0.21 | 0.18 | 1.58 |
| 0.63 | -1.36 | -1.41 | 0.68 | 1.07 | -1.11 | -1.12 | 1 | 0.89 | -1.08 | -1.25 | 0.99 | 1.69 | -1.75 | -1.89 | 2.71 |
| 2.14 | 3.75 | 4 | 2.27 | 2.16 | 2.93 | 2.81 | 2.06 | 1.86 | 1.85 | 2.11 | 2.1 | 2.37 | 2.36 | 2.45 | 3.39 |
| 2.14 | 3.75 | 4 | 2.27 | 1.09 | 2.93 | 2.81 | 1.02 | 1.67 | 1.25 | 1.42 | 1.88 | 1.19 | 3.4 | 3.55 | 2.21 |
| 1.41 | 0.56 | 0.62 | 1.5 | 0.84 | 0.58 | 0.52 | 0.78 | 0.99 | 1.02 | 1.15 | 1.1 | 0.98 | 0.77 | 0.77 | 2 |
| 1.22 | 0.91 | 0.99 | 1.3 | 0.66 | 1.23 | 1.16 | 0.6 | 1.15 | 0.93 | 1.05 | 1.29 | 0.81 | 0.81 | 0.81 | 1.83 |
| 0.43 | 1.41 | 1.52 | 0.46 | 1 | 2.16 | 2.06 | 0.93 | 1.57 | 0.82 | 0.93 | 1.77 | 0.85 | 0.67 | 0.66 | 1.87 |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 1.13 | 1.28 | 2.1 | 3.73 | 3.4 | 3.55 | 4.75 |
| -0.08 | 3.3 | 3.52 | -0.07 | -0.44 | 2.93 | 2.81 | -0.47 | 1.32 | 3.22 | 3.67 | 1.48 | 1.41 | 3.4 | 3.55 | 2.43 |


| 0.68 | 0.66 | 0.73 | 0.73 | 0.91 | 1.01 | 0.94 | 0.84 | 1 | 0.9 | 1.02 | 1.11 | 0.68 | 0.75 | 0.75 | 1.7 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1.34 | 3.44 | 3.67 | 1.43 | 0.93 | 0.86 | 0.8 | 0.86 | 1.67 | 2.26 | 2.57 | 1.88 | 0.62 | -1.2 | -1.31 | 1.64 |
| 1.15 | 0.23 | 0.27 | 1.23 | 0.27 | 0.76 | 0.7 | 0.22 | -1.37 | 0.67 | 0.76 | -1.6 | 0.94 | 1.27 | 1.3 | 1.96 |
| -0.48 | -0.77 | -0.78 | -0.5 | -0.39 | -1.14 | -1.15 | -0.42 | -0.26 | -2.48 | -2.85 | -0.33 | -0.32 | -1.05 | -1.15 | 0.7 |
| 1.06 | 2.14 | 2.3 | 1.13 | 1.03 | 2.17 | 2.07 | 0.96 | 1.7 | 0.88 | 1 | 1.91 | 0.63 | 1.4 | 1.44 | 1.65 |
| 2.14 | 1.9 | 2.04 | 2.27 | 1.2 | -1.39 | -1.39 | 1.13 | 1.67 | 3.22 | 3.67 | 1.88 | -1.53 | 3.4 | 3.55 | -0.51 |
| 0.48 | -0.1 | -0.07 | 0.52 | 1.06 | 1.42 | 1.34 | 0.99 | 0.9 | -0.25 | -0.3 | 1 | 0.79 | 0.31 | 0.28 | 1.81 |
| 0.09 | -0.91 | -0.93 | 0.11 | -0.68 | -0.03 | -0.07 | -0.7 | 0.4 | 1.01 | 1.14 | 0.43 | 0.98 | 0.7 | 0.7 | 2 |
| -1.87 | -0.65 | -0.66 | -1.97 | -0.43 | -0.88 | -0.89 | -0.46 | -1.76 | -1.12 | -1.29 | -2.04 | -0.69 | -1.3 | -1.41 | 0.33 |
| -1.36 | -0.84 | -0.86 | -1.43 | -0.44 | -1.26 | -1.26 | -0.47 | -1.71 | -1.13 | -1.3 | -1.99 | -0.47 | -0.81 | -0.9 | 0.55 |
| -2.01 | -0.8 | -0.82 | -2.12 | -0.53 | -0.73 | -0.75 | -0.56 | -1.26 | -0.79 | -0.91 | -1.47 | -0.53 | -1.34 | -1.46 | 0.49 |
| -0.02 | 3.15 | 3.37 | -0.01 | 1.7 | 0.99 | 0.92 | 1.61 | 1.98 | 3.22 | 3.67 | 2.24 | 0.37 | 0.41 | 0.39 | 1.39 |
| 1.7 | -0.37 | -0.36 | 1.81 | 0.71 | 0.81 | 0.75 | 0.65 | 1.67 | 0.79 | 0.89 | 1.88 | 3.73 | 0.72 | 0.72 | 4.75 |
| 0 | 0.74 | 0.81 | 0.01 | 0.77 | 0.97 | 0.9 | 0.71 | 1.02 | 0.61 | 0.69 | 1.14 | 1.45 | 0.45 | 0.43 | 2.47 |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.72 | 2.6 | 3.85 | 1.73 | 0.95 | 1.08 | 1.95 | 0.18 | 0.69 | 0.69 | 1.2 |
| -0.05 | 1.45 | 1.57 | -0.04 | 0.74 | 0.92 | 0.86 | 0.68 | -0.85 | 0.52 | 0.58 | -1 | 0.17 | 0.86 | 0.87 | 1.19 |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.76 | 2.64 | 3.85 | 0.34 | 3.22 | 3.67 | 0.36 | 3.73 | 3.4 | 3.55 | 4.75 |
| 1.28 | 2.92 | 3.12 | 1.36 | 2.67 | 0.84 | 0.78 | 2.56 | 1.86 | 3.22 | 3.67 | 2.1 | 1.52 | 1.14 | 1.16 | 2.54 |
| 1.81 | 1.07 | 1.16 | 1.92 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 0.88 | 1 | 2.1 | 1.28 | 1.12 | 1.14 | 2.3 |
| 1.06 | 0.53 | 0.59 | 1.13 | 4 | 2.93 | 2.81 | 3.85 | 1.88 | 1.02 | 1.16 | 2.12 | 3.03 | 3.4 | 3.55 | 4.05 |
| 2.14 | 3.75 | 4 | 2.27 | 1.09 | 2.93 | 2.81 | 1.02 | 1.86 | 3.22 | 3.67 | 2.1 | 1.81 | 0.61 | 0.6 | 2.83 |
| 0.46 | 1.9 | 2.04 | 0.5 | 1.05 | 0.29 | 0.24 | 0.98 | 0.61 | 0.32 | 0.36 | 0.67 | -0.16 | 0.62 | 0.61 | 0.86 |
| -0.34 | -1.1 | -1.13 | -0.35 | 0.1 | 0.28 | 0.23 | 0.06 | 0.48 | -1.2 | -1.39 | 0.52 | 0.98 | 0.67 | 0.66 | 2 |
| 1.37 | 0.11 | 0.15 | 1.46 | 0.22 | 0.26 | 0.21 | 0.17 | 0.22 | 0.39 | 0.43 | 0.22 | 0.62 | 0.02 | -0.02 | 1.64 |
| 0.16 | 0.75 | 0.83 | 0.18 | 0.47 | 0.98 | 0.91 | 0.42 | 0.9 | 0.22 | 0.24 | 1 | 0.81 | 0.68 | 0.68 | 1.83 |


| -1.06 | -1.29 | -1.33 | -1.11 | -0.23 | -1.54 | -1.54 | -0.26 | -0.31 | -0.66 | -0.77 | -0.38 | -0.9 | -1.2 | -1.31 | 0.12 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1.56 | 1.53 | 1.65 | 1.66 | 1.18 | 1.23 | 1.16 | 1.11 | 0.88 | 1.38 | 1.57 | 0.98 | 2.02 | 0.84 | 0.84 | 3.04 |
| 1.15 | 0.24 | 0.29 | 1.23 | 0.55 | 1.22 | 1.15 | 0.49 | 1.12 | 0.55 | 0.62 | 1.25 | 0.6 | 0.84 | 0.85 | 1.62 |
| 1.52 | -0.16 | -0.14 | 1.62 | 0.55 | 0.3 | 0.25 | 0.49 | 1.15 | 0.63 | 0.71 | 1.29 | 1.4 | 0.77 | 0.77 | 2.42 |
| -0.97 | -0.87 | -0.89 | -1.02 | -1.25 | -1.49 | -1.49 | -1.26 | -1.7 | -1.13 | -1.3 | -1.97 | -0.44 | -0.97 | -1.07 | 0.58 |
| 0.96 | 0 | 0.03 | 1.03 | 0.01 | 0.52 | 0.47 | -0.03 | -0.26 | 0.43 | 0.48 | -0.33 | 0.89 | -0.01 | -0.05 | 1.91 |
| 0.31 | 0.82 | 0.9 | 0.34 | 1.15 | 1.68 | 1.59 | 1.08 | 0.54 | 0.87 | 0.98 | 0.59 | 1.18 | 0.59 | 0.58 | 2.2 |
| 1.1 | 1.69 | 1.82 | 1.17 | 1.4 | 0.98 | 0.91 | 1.32 | -0.2 | 0.86 | 0.97 | -0.26 | 0.43 | 1.02 | 1.03 | 1.45 |
| 1.24 | 1.05 | 1.14 | 1.32 | 0.54 | 0.37 | 0.32 | 0.48 | 0.83 | 0.34 | 0.38 | 0.92 | 0.78 | 0.27 | 0.24 | 1.8 |
| 1.07 | 0.77 | 0.85 | 1.14 | -0.01 | 1.74 | 1.65 | -0.05 | 1.92 | 0.89 | 1.01 | 2.17 | 1.15 | 0.55 | 0.54 | 2.17 |
| 0.2 | 0.01 | 0.04 | 0.22 | 0.49 | 0.65 | 0.59 | 0.44 | -1.26 | 0.45 | 0.5 | -1.47 | 0.69 | 0.29 | 0.26 | 1.71 |
| 1.09 | 0.87 | 0.95 | 1.16 | 1.49 | 1.41 | 1.33 | 1.41 | 1.23 | 0.75 | 0.85 | 1.38 | 0.66 | 0.76 | 0.76 | 1.68 |
| 0.73 | 0.89 | 0.97 | 0.78 | 1 | 1.22 | 1.15 | 0.93 | 1.08 | -0.3 | -0.35 | 1.21 | 1.92 | 1.01 | 1.02 | 2.94 |
| 1.54 | 1.2 | 1.3 | 1.64 | 0.9 | 1.38 | 1.3 | 0.83 | 0.6 | 0.69 | 0.78 | 0.66 | 1.6 | 1.07 | 1.09 | 2.62 |
| 0.87 | 1.01 | 1.1 | 0.93 | 0.17 | 0.43 | 0.38 | 0.13 | 0.87 | 0.78 | 0.88 | 0.97 | 0.84 | 0.67 | 0.66 | 1.86 |
| -0.59 | -0.65 | -0.66 | -0.61 | 0.88 | -0.94 | -0.95 | 0.82 | -1.39 | -2.61 | -3 | -1.62 | 0.41 | -1.31 | -1.43 | 1.43 |
| 1.35 | 0.13 | 0.17 | 1.44 | 2.02 | 0.17 | 0.13 | 1.92 | 0.81 | 0.82 | 0.93 | 0.9 | 1.17 | 0.61 | 0.6 | 2.19 |
| -0.55 | -0.58 | -0.58 | -0.57 | -0.76 | -0.89 | -0.9 | -0.78 | 0.05 | -0.98 | -1.13 | 0.03 | -1.33 | -0.43 | -0.5 | -0.31 |
| -0.3 | -0.62 | -0.62 | -0.31 | -0.75 | -1.24 | -1.24 | -0.77 | -1.43 | -1.08 | -1.25 | -1.67 | 0.06 | -0.22 | -0.27 | 1.08 |
| -0.09 | 0.56 | 0.62 | -0.09 | 0.38 | 1.03 | 0.96 | 0.33 | 1.13 | 0.67 | 0.76 | 1.26 | 0.97 | 0.98 | 0.99 | 1.99 |
| 0.16 | 1.64 | 1.77 | 0.18 | 0.11 | -1.3 | -1.3 | 0.07 | 0.48 | -1.35 | -1.56 | 0.52 | 3.73 | -1.93 | -2.08 | 4.75 |
| 0.04 | 0.49 | 0.55 | 0.05 | 0.36 | 0.83 | 0.77 | 0.31 | -1.25 | 0.95 | 1.08 | -1.46 | 0.21 | 1.1 | 1.12 | 1.23 |
| 1.81 | 0.45 | 0.51 | 1.92 | 0.79 | 1.19 | 1.12 | 0.73 | 1.21 | 1.53 | 1.74 | 1.35 | 1.13 | 0.98 | 0.99 | 2.15 |
| 0.51 | 1.25 | 1.35 | 0.55 | 0.88 | 1.41 | 1.33 | 0.82 | 1.66 | 0.69 | 0.78 | 1.87 | 0.74 | 0.9 | 0.91 | 1.76 |
| 1.27 | 0.39 | 0.44 | 1.35 | 0.51 | 1.03 | 0.96 | 0.46 | 0.27 | 0.59 | 0.66 | 0.28 | 0.82 | 0.53 | 0.52 | 1.84 |


| 0.9 | 2.1 | 2.25 | 0.96 | 1.32 | 1.67 | 1.58 | 1.24 | 1.33 | 0.64 | 0.72 | 1.49 | 1.11 | 0.8 | 0.8 | 2.13 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.18 | -0.23 | -0.21 | 0.2 | 0.78 | 1.03 | 0.96 | 0.72 | 0.41 | 0.67 | 0.75 | 0.44 | 0.54 | 1.04 | 1.06 | 1.56 |
| 0.23 | 0.49 | 0.55 | 0.25 | 0.72 | 0.81 | 0.75 | 0.66 | 0.64 | 0.48 | 0.54 | 0.7 | 0.85 | 0.17 | 0.14 | 1.87 |
| 0.44 | 0.46 | 0.52 | 0.48 | 0.92 | 0.91 | 0.85 | 0.85 | 0.49 | 0.73 | 0.82 | 0.53 | 1.05 | 0.7 | 0.7 | 2.07 |
| 2.14 | 2.66 | 2.85 | 2.27 | 0.9 | 2.83 | 2.71 | 0.83 | 1.86 | -1.69 | -1.94 | 2.1 | 2.85 | 0.84 | 0.85 | 3.87 |
| 0.87 | 0.22 | 0.26 | 0.93 | 1.06 | 3 | 2.88 | 0.99 | 1.86 | 1.3 | 1.48 | 2.1 | 2 | 0.56 | 0.55 | 3.02 |
| 2.14 | 2.01 | 2.16 | 2.27 | 4 | 1.43 | 1.35 | 3.85 | 1.86 | 1.25 | 1.42 | 2.1 | 3.73 | 3.4 | 3.55 | 4.75 |
| 0.6 | 0.73 | 0.8 | 0.64 | 0.87 | 2.56 | 2.45 | 0.81 | 0.86 | 1.05 | 1.19 | 0.95 | 0.87 | 0.84 | 0.85 | 1.89 |
| -1.23 | 1.23 | 1.33 | -1.29 | -0.51 | -0.07 | -0.11 | -0.54 | -1.45 | 0.7 | 0.79 | -1.69 | 0.35 | 0.58 | 0.57 | 1.37 |
| -0.04 | -0.57 | -0.57 | -0.03 | 0.53 | 0.92 | 0.86 | 0.48 | 0.21 | -1.05 | -1.21 | 0.21 | 0.22 | 0.63 | 0.62 | 1.24 |
| 0.63 | 0.95 | 1.04 | 0.68 | 0.99 | 0.99 | 0.92 | 0.92 | 1.07 | 0.82 | 0.93 | 1.19 | 0.82 | 0.76 | 0.76 | 1.84 |
| 1.49 | 1.62 | 1.75 | 1.59 | 0.92 | 0.98 | 0.91 | 0.85 | 0.71 | 0.96 | 1.09 | 0.78 | 0.93 | 0.91 | 0.92 | 1.95 |
| 1.3 | 1.11 | 1.21 | 1.39 | 0.95 | 0.99 | 0.92 | 0.88 | 0.69 | 0.61 | 0.69 | 0.76 | 0.68 | 0.28 | 0.25 | 1.7 |
| 1.96 | 0.78 | 0.86 | 2.08 | 1.79 | 0.21 | 0.17 | 1.7 | 0.86 | -0.08 | -0.1 | 0.95 | 2.46 | 1.1 | 1.12 | 3.48 |
| -1.19 | -0.86 | -0.88 | -1.25 | -0.64 | -0.45 | -0.48 | -0.66 | -1.52 | -2.18 | -2.51 | -1.77 | -0.92 | -1.91 | -2.06 | 0.1 |
| 1.19 | 1.4 | 1.51 | 1.27 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 1.25 | 1.42 | 2.1 | 3.73 | 3.4 | 3.55 | 4.75 |
| 0.86 | 0.66 | 0.73 | 0.92 | 1.12 | -1.41 | -1.41 | 1.05 | 1.68 | 1.01 | 1.14 | 1.89 | -0.88 | -0.06 | -0.11 | 0.14 |
| 2.14 | 3.61 | 3.85 | 2.27 | 4 | 2.93 | 2.81 | 3.85 | 0.36 | 3.22 | 3.67 | 0.38 | -1.09 | 0.24 | 0.21 | -0.07 |
| -0.54 | -0.64 | -0.65 | -0.56 | -0.81 | -0.58 | -0.6 | -0.83 | -0.27 | -1.13 | -1.31 | -0.34 | -0.58 | -1.36 | -1.48 | 0.44 |
| 2.14 | 3.75 | 4 | 2.27 | 1.89 | 2.67 | 2.56 | 1.8 | 0.34 | 1.7 | 1.93 | 0.36 | 3.73 | 0.97 | 0.98 | 4.75 |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.72 | 2.6 | 3.85 | 1.73 | 0.95 | 1.08 | 1.95 | -1.46 | 0.69 | 0.69 | -0.44 |
| 1.66 | 0.29 | 0.34 | 1.77 | 0.76 | 0.88 | 0.82 | 0.7 | 0.29 | 0.68 | 0.77 | 0.3 | -1.53 | 0.44 | 0.42 | -0.51 |
| 0.5 | 1 | 1.09 | 0.54 | 0.36 | 0.85 | 0.79 | 0.31 | 1.3 | 0.58 | 0.65 | 1.46 | 0.97 | 0.82 | 0.82 | 1.99 |
| 0.48 | 0.99 | 1.08 | 0.52 | -1.04 | 1.32 | 1.24 | -1.05 | 0.86 | 0.8 | 0.9 | 0.95 | 0.77 | 1.01 | 1.02 | 1.79 |
| 0.22 | 0.67 | 0.74 | 0.24 | 0.88 | -0.71 | -0.73 | 0.82 | 1.14 | 0.68 | 0.77 | 1.27 | 1.19 | 0.99 | 1 | 2.21 |


| 0.07 | 0.38 | 0.43 | 0.08 | 0.76 | 0.5 | 0.45 | 0.7 | 0.94 | 1.03 | 1.17 | 1.05 | 0.23 | 0.8 | 0.8 | 1.25 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.15 | 1.18 | 1.28 | 0.17 | 0.74 | 1.26 | 1.19 | 0.68 | 0.48 | 1.13 | 1.28 | 0.52 | 1.24 | 1.06 | 1.08 | 2.26 |
| 0.44 | 0.64 | 0.71 | 0.48 | 0.89 | 0.97 | 0.9 | 0.83 | 0.64 | 0.73 | 0.82 | 0.7 | 1.05 | 0.67 | 0.67 | 2.07 |
| 1.03 | 0.4 | 0.45 | 1.1 | 0.57 | 0.1 | 0.06 | 0.51 | 0.93 | 0.85 | 0.96 | 1.03 | 0.48 | 0.84 | 0.84 | 1.5 |
| 1.51 | 1.16 | 1.26 | 1.61 | 0.67 | 1.34 | 1.26 | 0.61 | 0.9 | 0.89 | 1.01 | 1 | 0.64 | 0.89 | 0.9 | 1.66 |
| 1.4 | 0.92 | 1.01 | 1.49 | 1.25 | 1.22 | 1.15 | 1.18 | 0.89 | 1.15 | 1.31 | 0.99 | 1.26 | 1.22 | 1.25 | 2.28 |
| 0.51 | 0.66 | 0.73 | 0.55 | 0.52 | 0.78 | 0.72 | 0.47 | 0.78 | 0.78 | 0.88 | 0.86 | 1.07 | 0.93 | 0.94 | 2.09 |
| 0.34 | 0.28 | 0.33 | 0.37 | 1.08 | 0.9 | 0.84 | 1.01 | 0.19 | 0.59 | 0.66 | 0.19 | 1.13 | 0.52 | 0.51 | 2.15 |
| -0.01 | 0.35 | 0.4 | 0 | -0.5 | 0.54 | 0.49 | -0.53 | 0.2 | 0.25 | 0.28 | 0.2 | 0.29 | 0.13 | 0.09 | 1.31 |
| 0.09 | 0.78 | 0.86 | 0.11 | 1.11 | 0.81 | 0.75 | 1.04 | 0.77 | 0.71 | 0.8 | 0.85 | 0.7 | 0.45 | 0.43 | 1.72 |
| 0.38 | 0.64 | 0.71 | 0.41 | 0.07 | 0.94 | 0.87 | 0.03 | 0.6 | 0.25 | 0.28 | 0.66 | 0.72 | 0.52 | 0.51 | 1.74 |
| -0.78 | -0.53 | -0.53 | -0.82 | -0.82 | -0.35 | -0.38 | -0.84 | -0.28 | -1.13 | -1.3 | -0.35 | -0.31 | -0.09 | -0.14 | 0.71 |
| -1.91 | -0.42 | -0.41 | -2.01 | -0.94 | -0.29 | -0.32 | -0.95 | -0.07 | -0.28 | -0.33 | -0.11 | -1.31 | -1.18 | -1.29 | -0.29 |
| -1.51 | -0.72 | -0.73 | -1.59 | -1.6 | -0.87 | -0.88 | -1.6 | 0.34 | -0.69 | -0.8 | 0.36 | -0.62 | -2.12 | -2.28 | 0.4 |
| 0 | 0.74 | 0.81 | 0.01 | 0.83 | 0.97 | 0.9 | 0.77 | 0.78 | 0.6 | 0.67 | 0.86 | 1.26 | 0.24 | 0.21 | 2.28 |
| 1.22 | 0.79 | 0.87 | 1.3 | 0.73 | 1.09 | 1.02 | 0.67 | 1.28 | 0.9 | 1.02 | 1.43 | 0.81 | 0.77 | 0.77 | 1.83 |
| 1.03 | 0.56 | 0.62 | 1.1 | 0.69 | 0.64 | 0.58 | 0.63 | 0.9 | 0.81 | 0.91 | 1 | 0.52 | 0.84 | 0.85 | 1.54 |
| -0.63 | -0.5 | -0.5 | -0.66 | -0.85 | -1.71 | -1.7 | -0.87 | -0.19 | -0.48 | -0.56 | -0.25 | -1.22 | -0.37 | -0.43 | -0.2 |
| -1.24 | 1.23 | 1.33 | -1.3 | -0.76 | 1.09 | 1.02 | -0.78 | -0.14 | 0.91 | 1.03 | -0.19 | -0.73 | -1.12 | -1.22 | 0.29 |
| 0.06 | 0.38 | 0.43 | 0.07 | 0.2 | 1.45 | 1.37 | 0.15 | 0.63 | 0.83 | 0.94 | 0.69 | 0.41 | 1.08 | 1.1 | 1.43 |
| -0.19 | 1.01 | 1.1 | -0.19 | 0.7 | 0.85 | 0.79 | 0.64 | 0.85 | 0.93 | 1.05 | 0.94 | 0.64 | 0.57 | 0.56 | 1.66 |
| -0.02 | 0.73 | 0.8 | -0.01 | 0.88 | 0.78 | 0.72 | 0.82 | 0.85 | 0.6 | 0.67 | 0.94 | 0.73 | 0.33 | 0.31 | 1.75 |
| 0.16 | 0.84 | 0.92 | 0.18 | 1.13 | 0.81 | 0.75 | 1.06 | 1.21 | 0.6 | 0.68 | 1.35 | 1 | 1.3 | 1.33 | 2.02 |
| -0.04 | 0.83 | 0.91 | -0.03 | 0.88 | 1.01 | 0.94 | 0.82 | 0.68 | 0.66 | 0.74 | 0.75 | 0.85 | 0.31 | 0.28 | 1.87 |
| 1.1 | 0.35 | 0.4 | 1.17 | -0.5 | 0.45 | 0.4 | -0.53 | -0.09 | 0.18 | 0.19 | -0.13 | 0.52 | 0 | -0.04 | 1.54 |


| 1.1 | 0.35 | 0.4 | 1.17 | -0.5 | 0.45 | 0.4 | -0.53 | -0.09 | 0.18 | 0.19 | -0.13 | 0.52 | 0 | -0.04 | 1.54 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1.6 | 0.57 | 0.63 | 1.7 | -0.17 | 0.6 | 0.54 | -0.21 | 0.96 | 0.51 | 0.57 | 1.07 | 1.09 | -0.28 | -0.34 | 2.11 |
| 0.27 | 1.1 | 1.2 | 0.3 | 0.61 | 1.39 | 1.31 | 0.55 | 1.64 | 0.95 | 1.08 | 1.85 | 1.47 | 0.9 | 0.91 | 2.49 |
| -0.06 | 0.57 | 0.63 | -0.05 | 1.72 | 0.45 | 0.4 | 1.63 | 0.54 | 1.01 | 1.14 | 0.59 | 1 | 0.87 | 0.88 | 2.02 |
| -0.14 | 0.53 | 0.59 | -0.14 | 0.42 | 1.12 | 1.05 | 0.37 | 1.03 | 0.9 | 1.02 | 1.15 | 0.83 | 0.73 | 0.73 | 1.85 |
| 0.27 | 0.24 | 0.29 | 0.3 | 0.71 | 0.73 | 0.67 | 0.65 | 0.32 | 0.7 | 0.79 | 0.34 | 1.36 | 0.19 | 0.16 | 2.38 |
| 0.78 | 0.37 | 0.42 | 0.84 | 0.62 | 0.66 | 0.6 | 0.56 | 0.84 | 0.71 | 0.8 | 0.93 | 1.1 | 0.49 | 0.48 | 2.12 |
| 0.3 | 0.87 | 0.95 | 0.33 | -0.33 | 0.58 | 0.52 | -0.36 | 0.76 | 0.7 | 0.79 | 0.84 | 1.36 | -0.08 | -0.13 | 2.38 |
| 0.88 | 1.07 | 1.16 | 0.94 | 0.85 | 0.24 | 0.19 | 0.79 | -0.01 | 0.74 | 0.83 | -0.04 | 1.32 | 0.44 | 0.42 | 2.34 |
| 0.03 | 0.47 | 0.53 | 0.04 | 1.43 | 0.73 | 0.67 | 1.35 | 0.54 | 0.96 | 1.09 | 0.59 | 1 | 0.82 | 0.82 | 2.02 |
| 0.21 | 0.73 | 0.8 | 0.23 | 0.62 | 0.82 | 0.76 | 0.56 | 0.84 | 0.72 | 0.81 | 0.93 | 0.5 | 0.05 | 0.01 | 1.52 |
| -0.19 | 0.35 | 0.4 | -0.19 | -0.5 | 0.96 | 0.89 | -0.53 | 1.33 | 0.25 | 0.28 | 1.49 | 0.42 | 0.15 | 0.12 | 1.44 |
| 0.76 | 0.92 | 1.01 | 0.81 | 0.24 | 0.97 | 0.9 | 0.19 | 1.24 | 0.59 | 0.66 | 1.39 | 0.87 | 1.06 | 1.08 | 1.89 |
| 0.11 | 0.33 | 0.38 | 0.13 | 0.45 | 0.72 | 0.66 | 0.4 | 1.05 | 0.68 | 0.77 | 1.17 | 0.74 | 0.76 | 0.76 | 1.76 |
| 1.44 | 0.53 | 0.59 | 1.53 | 1.23 | 0.68 | 0.62 | 1.16 | 1 | 0.93 | 1.05 | 1.11 | 0.79 | 0.83 | 0.83 | 1.81 |
| 0.13 | 0.6 | 0.67 | 0.15 | -0.41 | 0.56 | 0.51 | -0.44 | 0.77 | 0.67 | 0.76 | 0.85 | 0.68 | -0.21 | -0.26 | 1.7 |
| -0.24 | -3.61 | -3.79 | -0.24 | 0.53 | 0.68 | 0.62 | 0.48 | 0.38 | 0.68 | 0.77 | 0.4 | 1.01 | 0.99 | 1 | 2.03 |
| -0.12 | 1.52 | 1.64 | -0.12 | 0.79 | -0.41 | -0.44 | 0.73 | 0.02 | -1.08 | -1.25 | -0.01 | 0.84 | -1.14 | -1.25 | 1.86 |
| 0.29 | 0.67 | 0.74 | 0.32 | 0.3 | 0.68 | 0.62 | 0.25 | 0.57 | 0.73 | 0.82 | 0.62 | 0.95 | 0.99 | 1 | 1.97 |
| 0.6 | 1.56 | 1.68 | 0.64 | 0.56 | 0.73 | 0.67 | 0.5 | 0.36 | 0.67 | 0.75 | 0.38 | 0.81 | 0.99 | 1 | 1.83 |
| -0.52 | 0.67 | 0.74 | -0.54 | -0.72 | 0.68 | 0.62 | -0.74 | 0.27 | 0.67 | 0.75 | 0.28 | -0.22 | 0.95 | 0.96 | 0.8 |
| -1.35 | 0.67 | 0.74 | -1.42 | -0.94 | 0.68 | 0.62 | -0.95 | 0.65 | 0.68 | 0.77 | 0.71 | 0.34 | 0.99 | 1 | 1.36 |
| -0.24 | 0.32 | 0.37 | -0.24 | 0.46 | 1.07 | 1 | 0.41 | 0.74 | 1.09 | 1.24 | 0.82 | 0.53 | 0.66 | 0.65 | 1.55 |
| -0.09 | 0.67 | 0.74 | -0.09 | 0.79 | 0.68 | 0.62 | 0.73 | 0.58 | 0.68 | 0.77 | 0.63 | 0.76 | 0.99 | 1 | 1.78 |
| -0.52 | 0.67 | 0.74 | -0.54 | 0.3 | 0.68 | 0.62 | 0.25 | 0.41 | 0.68 | 0.77 | 0.44 | 0.84 | 0.99 | 1 | 1.86 |


| 1.6 | 0.67 | 0.74 | 1.7 | 1.33 | -0.57 | -0.59 | 1.25 | 0.9 | 0.68 | 0.77 | 1 | 2.41 | 0.99 | 1 | 3.43 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.84 | 0.42 | 0.48 | 0.9 | 0.91 | 0.65 | 0.59 | 0.84 | 1.59 | 0.66 | 0.74 | 1.79 | 1.95 | 0.97 | 0.98 | 2.97 |
| 0.66 | 0.9 | 0.98 | 0.71 | 1.22 | 0.68 | 0.62 | 1.15 | 1.94 | 0.68 | 0.77 | 2.19 | 1.57 | 0.41 | 0.39 | 2.59 |
| 0.44 | -0.15 | -0.13 | 0.48 | 0.04 | 0.68 | 0.62 | 0 | 0.47 | 0.68 | 0.77 | 0.51 | 1.96 | 0.95 | 0.96 | 2.98 |
| 2 | 0.67 | 0.74 | 2.13 | 1.08 | 0.68 | 0.62 | 1.01 | 1.46 | 0.68 | 0.77 | 1.64 | 0.91 | 0.99 | 1 | 1.93 |
| 1.56 | 0.67 | 0.74 | 1.66 | 0.62 | 0.68 | 0.62 | 0.56 | 0.42 | 0.68 | 0.77 | 0.45 | 1.69 | 0.88 | 0.89 | 2.71 |
| 1.76 | 1.87 | 2.01 | 1.87 | 1.26 | 0.9 | 0.84 | 1.18 | 3.58 | 1 | 1.13 | 4.07 | 1.51 | 1.21 | 1.24 | 2.53 |
| 0.66 | 0.92 | 1 | 0.71 | -1.06 | -0.17 | -0.2 | -1.07 | 0.23 | 1.11 | 1.26 | 0.23 | -0.3 | 1.01 | 1.02 | 0.72 |
| 0.9 | 0.93 | 1.02 | 0.96 | 0.93 | 1.17 | 1.1 | 0.86 | 0.55 | 1.52 | 1.73 | 0.6 | 1.01 | 0.34 | 0.32 | 2.03 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.63 | -1.65 | -1.64 | 0.57 | 1.11 | -0.09 | -0.12 | 1.24 | 0.74 | -1.12 | -1.22 | 1.76 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.63 | -1.65 | -1.64 | 0.57 | 1.21 | -0.09 | -0.12 | 1.35 | 0.74 | -1.12 | -1.22 | 1.76 |
| 0.08 | 0.85 | 0.93 | 0.09 | -0.44 | -1.65 | -1.64 | -0.47 | 0.55 | -0.09 | -0.12 | 0.6 | 0.31 | -1.12 | -1.22 | 1.33 |
| 0.03 | 0.85 | 0.93 | 0.04 | 0.38 | -1.65 | -1.64 | 0.33 | 0.55 | -0.09 | -0.12 | 0.6 | 0.56 | -1.12 | -1.22 | 1.58 |
| -0.06 | 0.85 | 0.93 | -0.05 | 0.29 | -1.65 | -1.64 | 0.24 | 0.55 | -0.09 | -0.12 | 0.6 | 0.56 | -1.12 | -1.22 | 1.58 |
| -0.78 | -0.71 | -0.72 | -0.82 | -0.85 | -1.4 | -1.4 | -0.87 | -1.37 | -1.79 | -2.06 | -1.6 | -0.3 | -1.66 | -1.8 | 0.72 |
| -0.78 | -0.71 | -0.72 | -0.82 | -0.85 | -1.4 | -1.4 | -0.87 | -1.37 | -1.79 | -2.06 | -1.6 | -0.3 | -1.66 | -1.8 | 0.72 |
| -0.65 | -0.42 | -0.41 | -0.68 | 0.25 | -0.02 | -0.06 | 0.2 | 0.09 | -1.32 | -1.52 | 0.07 | 0.3 | 0.35 | 0.33 | 1.32 |
| 0.04 | 1.28 | 1.39 | 0.05 | -0.66 | 0.65 | 0.59 | -0.68 | -0.28 | 1.1 | 1.25 | -0.35 | 0.19 | 0.63 | 0.62 | 1.21 |
| -2.76 | -1.15 | -1.18 | -2.91 | -0.31 | -0.82 | -0.84 | -0.34 | -0.55 | -1.95 | -2.24 | -0.66 | -0.52 | -0.93 | -1.02 | 0.5 |
| -0.99 | -0.88 | -0.9 | -1.04 | -0.16 | -1.87 | -1.86 | -0.2 | -0.2 | -1.87 | -2.15 | -0.26 | -1.1 | -0.8 | -0.89 | -0.08 |
| -1.71 | -0.72 | -0.73 | -1.8 | -0.18 | -1.9 | -1.89 | -0.21 | -1.33 | -2.16 | -2.48 | -1.55 | -0.44 | -0.67 | -0.75 | 0.58 |
| -0.85 | -1.06 | -1.09 | -0.89 | -0.25 | -1.09 | -1.1 | -0.28 | -0.85 | -1.12 | -1.29 | -1 | -0.67 | -1.17 | -1.28 | 0.35 |
| -1.67 | -0.91 | -0.93 | -1.76 | -0.54 | -2.44 | -2.41 | -0.56 | -1.17 | -0.74 | -0.86 | -1.37 | -0.68 | -1.21 | -1.32 | 0.34 |
| -1.3 | -0.68 | -0.69 | -1.37 | -0.55 | -1.06 | -1.07 | -0.57 | -1.19 | -1.1 | -1.27 | -1.39 | -0.38 | -1.49 | -1.62 | 0.64 |
| -1.04 | -0.85 | -0.87 | -1.09 | -0.51 | -1.28 | -1.28 | -0.54 | -1.75 | -2.47 | -2.84 | -2.03 | -0.57 | -1.18 | -1.29 | 0.45 |


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| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| -3.01 | -0.72 | -0.73 | -3.17 | -0.4 | -1.96 | -1.94 | -0.43 | -1.47 | -1.48 | -1.71 | -1.71 | -0.32 | -0.98 | -1.08 |  |
| -1.5 | -0.91 | -0.93 | -1.58 | -0.69 | -1.01 | -1.02 | -0.71 | -0.52 | -2.15 | -2.47 | -0.62 | -0.28 | -0.9 | -0.99 | 0.74 |
| -4 | -0.45 | -0.44 | -4.22 | -0.98 | -1.39 | -1.39 | -0.99 | -0.7 | -1.13 | -1.3 | -0.83 | -0.08 | -1.32 | -1.44 | 0.94 |
| 0.87 | -3.61 | -3.79 | 0.93 | 1.04 | 1.03 | 0.96 | 0.97 | 1.59 | 0.67 | 0.76 | 1.79 | 1.18 | 0.85 | 0.86 |  |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 3.22 | 3.67 | 2.1 | 3.73 | 0.71 | 0.71 | 4.75 |
| 0.65 | 0.47 | 0.53 | 0.7 | 0.67 | 0.7 | 0.64 | 0.61 | 0.4 | 0.56 | 0.63 | 0.43 | 0.56 | 0.31 | 0.28 |  |
| 1.67 | 3.75 | 4 | 1.78 | 2.65 | 2.09 | 1.99 | 2.54 | 1.28 | 1.99 | 2.26 | 1.43 | 3.03 | 0.76 | 0.76 |  |
| 2.14 | 3.14 | 3.35 | 2.27 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 1.42 | 1.61 | 2.1 | 2.27 | 1.03 | 1.05 | 3.05 |
| 2.28 | 3.75 | 4 | 2.42 | 4 | -1.33 | -1.33 | 3.85 | 1.86 | 3.22 | 3.67 | 2.1 | 1.23 | 1.16 | 1.18 |  |
| 1 | -1.5 | -1.56 | 1.07 | 0.74 | 0.06 | 0.02 | 0.68 | 0.96 | 0.85 | 0.96 | 1.07 | 1.64 | -0.27 | -0.33 |  |
| 0.55 | -0.14 | -0.12 | 0.59 | 0.27 | 0.14 | 0.1 | 0.22 | -0.75 | 0.94 | 1.06 | -0.89 | 0.03 | 0.31 | 0.29 | 1.05 |
| 1.56 | 3.75 | 4 | 1.66 | 4 | 2.93 | 2.81 | 3.85 | 1.86 | 3.22 | 3.67 | 2.1 | 2.8 | 2.34 | 2.43 | 3.82 |
| 1.62 | 1.57 | 1.69 | 1.72 | 2.12 | 0.89 | 0.83 | 2.02 | -0.09 | -0.97 | -1.12 | -0.13 | 0.18 | 3.17 | 3.3 | 1.2 |
| 0.71 | 3.75 | 4 | 0.76 | -2.14 | -0.24 | -0.27 | -2.12 | 0.27 | 0.38 | 0.42 | 0.28 | 0.19 | 0.6 | 0.59 | 1.21 |
| -1.88 | 0.4 | 0.46 | -1.98 | -0.66 | -0.61 | -0.63 | -0.68 | -0.9 | 3.22 | 3.67 | -1.06 | 0.41 | 0.63 | 0.62 | 1.43 |
| -1.49 | -0.7 | -0.71 | -1.57 | -0.48 | -0.78 | -0.8 | -0.51 | -1.32 | -1.09 | -1.26 | -1.54 | -0.03 | -0.51 | -0.58 | 0.99 |
| -1.99 | -0.99 | -1.02 | -2.1 | -0.57 | -1.35 | -1.35 | -0.59 | -1.42 | -1.13 | -1.31 | -1.65 | -0.26 | -0.3 | -0.36 | 0.76 |
| -0.62 | -1.03 | -1.06 | -0.65 | -0.64 | -4.08 | -4 | -0.66 | -1.32 | -1.13 | -1.3 | -1.54 | -0.35 | -1.17 | -1.28 | 0.67 |
| 2.14 | 3.75 | 4 | 2.27 | 4 | 2.72 | 2.6 | 3.85 | -0.06 | 3.22 | 3.67 | -0.1 | -1.06 | 3.4 | 3.55 | -0.04 |
| -3.29 | -0.69 | -0.7 | -3.47 | -0.34 | -0.88 | -0.89 | -0.37 | -1.09 | -2.18 | -2.51 | -1.28 | -0.67 | -0.76 | -0.84 | 0.35 |
| -0.73 | -0.9 | -0.92 | -0.76 | -0.51 | -2.17 | -2.15 | -0.54 | -0.69 | -1.12 | -1.29 | -0.82 | -0.3 | -2.14 | -2.3 | 0.72 |
| -0.35 | -0.04 | -0.01 | -0.36 | 0.58 | 0.79 | 0.73 | 0.52 | -0.33 | 0.46 | 0.52 | -0.41 | 0.88 | 0.75 | 0.75 | 1.9 |
| -0.15 | 0.81 | 0.89 | -0.15 | 0.03 | 0.9 | 0.84 | -0.01 | 0.8 | 1.17 | 1.33 | 0.89 | 1.1 | 0.01 | -0.03 | 2.12 |
| 0.45 | 0.32 | 0.37 | 0.49 | -0.09 | 0.35 | 0.3 | -0.13 | 0.35 | 0.49 | 0.55 | 0.37 | 0.54 | 0.48 | 0.46 | 1.56 |
| -0.49 | 0.27 | 0.32 | -0.51 | 0.42 | 1.26 | 1.19 | 0.37 | 0.82 | 0.81 | 0.91 | 0.91 | 0.97 | 0.39 | 0.37 | 1.99 |


| -0.49 | 0.27 | 0.32 | -0.51 | 0.42 | 1.26 | 1.19 | 0.37 | 0.54 | 0.81 | 0.91 | 0.59 | 0.97 | 0.38 | 0.36 | 1.99 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.61 | 1.42 | 1.53 | 0.66 | 0.53 | 1.42 | 1.34 | 0.48 | 0.94 | 0.72 | 0.81 | 1.05 | 0.86 | 0.48 | 0.46 | 1.88 |
| 0.5 | -0.12 | -0.09 | 0.54 | 0.93 | 0.74 | 0.68 | 0.86 | 0.58 | 0.82 | 0.93 | 0.63 | 1.25 | 0.21 | 0.18 | 2.27 |
| -0.23 | 0.19 | 0.23 | -0.23 | 1.14 | 1.13 | 1.06 | 1.07 | 0.88 | 0.54 | 0.61 | 0.98 | 0.25 | -1.13 | -1.24 | 1.27 |
| 0.05 | 0.61 | 0.68 | 0.06 | 0.55 | 0.91 | 0.85 | 0.49 | 0.86 | 0.57 | 0.64 | 0.95 | 0.37 | 0.84 | 0.84 | 1.39 |
| 1.08 | 2.23 | 2.39 | 1.15 | -0.44 | -2.67 | -2.63 | -0.47 | 1.5 | 0.72 | 0.81 | 1.69 | -0.48 | 0.96 | 0.97 | 0.54 |
| -0.55 | 0.7 | 0.77 | -0.57 | 0.99 | 0.5 | 0.45 | 0.92 | 0.66 | 1.09 | 1.24 | 0.73 | 1.23 | 0.76 | 0.76 | 2.25 |
| 0.71 | 0.6 | 0.67 | 0.76 | 0.36 | 0.05 | 0.01 | 0.31 | 1.86 | -0.2 | -0.24 | 2.1 | 0.51 | -1.12 | -1.22 | 1.53 |
| 0.14 | 0.42 | 0.48 | 0.16 | 0.6 | 0.54 | 0.49 | 0.54 | 0.96 | 0.39 | 0.44 | 1.07 | 0.75 | -0.13 | -0.18 | 1.77 |
| 0.96 | 1.31 | 1.42 | 1.03 | 1.31 | 1.44 | 1.36 | 1.23 | 0.37 | 0.75 | 0.85 | 0.39 | 1.04 | 0.63 | 0.62 | 2.06 |
| 2.24 | 3.75 | 4 | 2.38 | 4 | 2.63 | 2.52 | 3.85 | 1.86 | 3.22 | 3.67 | 2.1 | 2.8 | 1.24 | 1.27 | 3.82 |
| 0.98 | -0.41 | -0.4 | 1.05 | 1.15 | -0.45 | -0.48 | 1.08 | 0.75 | -0.76 | -0.88 | 0.83 | 2.87 | 0.7 | 0.7 | 3.89 |
| 2.14 | 1.85 | 1.99 | 2.27 | 0.51 | 2.93 | 2.81 | 0.46 | 0.5 | 1.44 | 1.64 | 0.54 | 0.89 | 0.67 | 0.67 | 1.91 |
| 0.47 | 0.47 | 0.53 | 0.51 | 0.81 | 1.31 | 1.23 | 0.75 | 1.65 | 0.39 | 0.43 | 1.86 | 0.23 | 0.3 | 0.27 | 1.25 |
| -0.78 | -0.71 | -0.72 | -0.82 | -0.85 | -1.4 | -1.4 | -0.87 | -1.37 | -1.79 | -2.06 | -1.6 | -0.3 | -1.66 | -1.8 | 0.72 |
| 0.47 | 0.47 | 0.53 | 0.51 | 0.68 | 1.31 | 1.23 | 0.62 | 1.65 | 0.39 | 0.43 | 1.86 | 0.68 | 0.3 | 0.27 | 1.7 |
| 0.68 | 1.01 | 1.1 | 0.73 | 0.96 | 2.15 | 2.05 | 0.89 | 1.08 | 1.3 | 1.48 | 1.21 | 1.33 | 0.44 | 0.42 | 2.35 |
| 0.17 | 0.17 | 0.21 | 0.19 | 0.38 | 0.32 | 0.27 | 0.33 | 0.38 | 0.9 | 1.02 | 0.4 | 0.5 | -1.07 | -1.17 | 1.52 |
| 0.3 | 0.11 | 0.15 | 0.33 | 0.27 | 0.71 | 0.65 | 0.22 | 0.52 | 0.61 | 0.69 | 0.56 | 1.02 | 0.23 | 0.2 | 2.04 |
| 0.09 | 0.55 | 0.61 | 0.11 | 0.46 | 0.72 | 0.66 | 0.41 | 0.97 | 0.59 | 0.66 | 1.08 | 0.69 | 0.22 | 0.19 | 1.71 |
| 0.26 | 0.46 | 0.52 | 0.29 | 0.07 | 0.85 | 0.79 | 0.03 | 0.15 | 0.92 | 1.04 | 0.14 | 0.89 | 0.36 | 0.34 | 1.91 |
| 0.24 | 1.07 | 1.16 | 0.26 | 0.53 | 0.89 | 0.83 | 0.48 | 0.47 | 0.9 | 1.02 | 0.51 | 1.12 | 0.49 | 0.47 | 2.14 |
| -0.02 | 1.91 | 2.05 | -0.01 | 0.63 | 0.21 | 0.17 | 0.57 | 0.81 | 0.6 | 0.68 | 0.9 | 1.01 | -0.15 | -0.2 | 2.03 |
| 0.32 | 0.3 | 0.35 | 0.35 | 0.9 | 0.69 | 0.63 | 0.83 | 1.1 | 0.34 | 0.38 | 1.23 | 0.38 | 0.77 | 0.77 | 1.4 |
| 0.76 | 0.3 | 0.35 | 0.81 | 0.77 | 0.73 | 0.67 | 0.71 | -0.62 | 1.09 | 1.23 | -0.74 | 0.79 | 0.32 | 0.3 | 1.81 |


| -0.77 | -1.26 | -1.3 | -0.8 | -0.5 | -1.02 | -1.03 | -0.53 | 0.29 | -1.13 | -1.31 | 0.3 | 1.18 | -1.16 | -1.27 | 2.2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2.14 | 0.27 | 0.32 | 2.27 | 4 | 0.34 | 0.29 | 3.85 | 1.67 | 0.22 | 0.24 | 1.88 | 0.36 | 0.31 | 0.29 | 1.38 |
| 0.15 | 0.21 | 0.25 | 0.17 | 0.57 | 0.84 | 0.78 | 0.51 | 0.74 | 0.66 | 0.74 | 0.82 | 0.74 | 0.78 | 0.78 | 1.76 |
| 1.25 | 0.76 | 0.84 | 1.33 | 0.98 | 1.25 | 1.18 | 0.91 | 0.96 | 0.81 | 0.92 | 1.07 | 0.68 | 0.39 | 0.37 | 1.7 |
| 0.53 | 0.45 | 0.51 | 0.57 | 0.48 | 1.39 | 1.31 | 0.43 | 0.19 | 0.6 | 0.67 | 0.19 | 1.23 | 0.81 | 0.81 | 2.25 |
| 0.91 | 1.29 | 1.4 | 0.97 | 0.3 | 1.2 | 1.13 | 0.25 | 0.61 | 0.53 | 0.59 | 0.67 | 1.48 | 0.52 | 0.51 | 2.5 |
| -1.16 | -0.56 | -0.56 | -1.22 | -0.61 | -1.47 | -1.47 | -0.63 | -0.8 | -1.76 | -2.03 | -0.95 | -0.3 | -1.3 | -1.42 | 0.72 |
| 0.04 | 0.53 | 0.59 | 0.05 | 0.89 | 0.64 | 0.58 | 0.83 | 1.01 | 1.01 | 1.14 | 1.13 | 0.91 | 0.82 | 0.82 | 1.93 |
| 0.36 | 0.37 | 0.42 | 0.39 | 0.83 | 0.89 | 0.83 | 0.77 | 0.02 | 0.84 | 0.95 | -0.01 | 0.6 | 0.96 | 0.97 | 1.62 |
| 0.5 | 0.07 | 0.11 | 0.54 | 0.38 | 0.37 | 0.32 | 0.33 | 0.42 | 0.29 | 0.32 | 0.45 | 0.91 | 0.39 | 0.37 | 1.93 |
| 0.77 | 0.41 | 0.47 | 0.82 | 0.26 | 0.75 | 0.69 | 0.21 | 0.62 | 0.44 | 0.49 | 0.68 | -0.06 | 0.46 | 0.44 | 0.96 |
| -0.37 | 0.12 | 0.16 | -0.38 | -0.49 | 0.8 | 0.74 | -0.52 | 0.69 | 0.24 | 0.26 | 0.76 | 0.34 | 0.42 | 0.4 | 1.36 |
| -0.12 | -0.93 | -0.95 | -0.12 | -2.38 | -1.38 | -1.38 | -2.35 | -2.23 | -0.44 | -0.51 | -2.58 | -0.42 | -0.2 | -0.25 | 0.6 |
| 0.35 | -0.68 | -0.69 | 0.38 | 0.21 | -0.32 | -0.35 | 0.16 | 0.7 | -1.25 | -1.44 | 0.77 | 1.83 | -1.15 | -1.26 | 2.85 |
| 0.41 | 0.5 | 0.56 | 0.44 | 0.87 | 0.64 | 0.58 | 0.81 | 0.75 | 0.8 | 0.9 | 0.83 | 1.04 | 0.45 | 0.43 | 2.06 |
| -0.32 | 0.05 | 0.08 | -0.33 | 0.79 | 1.3 | 1.22 | 0.73 | 0.94 | 0.42 | 0.47 | 1.05 | 0.82 | 0.59 | 0.58 | 1.84 |
| -0.32 | 0.47 | 0.53 | -0.33 | 0.59 | 0.49 | 0.44 | 0.53 | 0.59 | 0.8 | 0.9 | 0.64 | 0.78 | 0.65 | 0.64 | 1.8 |
| 0.52 | 0.41 | 0.47 | 0.56 | 0.55 | 0.23 | 0.18 | 0.49 | 1.01 | 0.83 | 0.94 | 1.13 | 0.89 | 0.81 | 0.81 | 1.91 |
| 0.35 | 0.04 | 0.07 | 0.38 | 0.68 | 0.68 | 0.62 | 0.62 | 1.29 | 0.93 | 1.05 | 1.45 | 0.64 | 0.85 | 0.86 | 1.66 |
| -3.29 | 0.35 | 0.4 | -3.47 | 0.34 | 1.15 | 1.08 | 0.29 | 0.61 | 0.81 | 0.91 | 0.67 | 0.63 | 0.62 | 0.61 | 1.65 |
| 1.74 | -0.87 | -0.89 | 1.85 | -1.89 | 0.03 | -0.01 | -1.88 | -1.3 | 0.09 | 0.09 | -1.52 | -1.53 | 1.13 | 1.15 | -0.51 |
| -0.87 | -0.98 | -1.01 | -0.91 | -1.29 | -0.05 | -0.09 | -1.29 | -1.37 | 0.05 | 0.05 | -1.6 | -1.31 | 1.23 | 1.26 | -0.29 |
| 1.5 | -0.87 | -0.89 | 1.6 | -1.04 | -1.79 | -1.78 | -1.05 | -0.23 | -0.6 | -0.7 | -0.29 | -1.51 | 0.96 | 0.97 | -0.49 |
| -0.75 | -0.87 | -0.89 | -0.78 | -1.33 | -0.05 | -0.09 | -1.33 | -1.3 | 0.09 | 0.09 | -1.52 | -1.42 | 1.29 | 1.32 | -0.4 |
| 0.04 | -0.98 | -1.01 | 0.05 | -1.83 | -0.12 | -0.16 | -1.82 | -1.3 | 0.13 | 0.14 | -1.52 | -1.53 | 1.39 | 1.43 | -0.51 |


| -0.13 | -0.67 | -0.68 | -0.13 | -2.21 | -0.12 | -0.16 | -2.19 | -1.22 | 0.04 | 0.03 | -1.43 | -1.35 | 1.38 | 1.42 | -0.33 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1.74 | -0.98 | -1.01 | 1.85 | -1.71 | -0.19 | -0.22 | -1.7 | -1.32 | -0.8 | -0.93 | -1.54 | -1.53 | 1.02 | 1.04 | -0.51 |
| 1.32 | -0.98 | -1.01 | 1.41 | -1.83 | -0.24 | -0.27 | -1.82 | -1.26 | -0.03 | -0.05 | -1.47 | -1.49 | -0.01 | -0.05 | -0.47 |
| 1.51 | -0.83 | -0.85 | 1.61 | -1.85 | 0.03 | -0.01 | -1.84 | -1.31 | 0.04 | 0.04 | -1.53 | -1.31 | 1.14 | 1.16 | -0.29 |
| 1.4 | -0.87 | -0.89 | 1.49 | -1.06 | 0.03 | -0.01 | -1.07 | -1.26 | 0.09 | 0.09 | -1.47 | -1.53 | 1.41 | 1.45 | -0.51 |
| 1.64 | -0.87 | -0.89 | 1.75 | -1.44 | -0.19 | -0.22 | -1.44 | -1.3 | 0.04 | 0.03 | -1.52 | -1.35 | 1.48 | 1.52 | -0.33 |
| 1.78 | -0.98 | -1.01 | 1.89 | -0.97 | -0.13 | -0.17 | -0.98 | -1.3 | -0.02 | -0.04 | -1.52 | -1.44 | 1.38 | 1.42 | -0.42 |
| 0.81 | -1.05 | -1.08 | 0.87 | -1.83 | -0.02 | -0.06 | -1.82 | -1.3 | 0.12 | 0.13 | -1.52 | -1.53 | 1.48 | 1.52 | -0.51 |
| 1.78 | -0.98 | -1.01 | 1.89 | -2.31 | 0.03 | -0.01 | -2.29 | 0.14 | 0.09 | 0.09 | 0.13 | -1.44 | 1.48 | 1.52 | -0.42 |
| 1.83 | -0.94 | -0.96 | 1.95 | -2.22 | -0.25 | -0.28 | -2.2 | -1.25 | -0.04 | -0.06 | -1.46 | -1.35 | 0.6 | 0.59 | -0.33 |
| 1.74 | -0.98 | -1.01 | 1.85 | -1.67 | -0.03 | -0.07 | -1.66 | -1.3 | -0.1 | -0.13 | -1.52 | -1.53 | 0.94 | 0.95 | -0.51 |
| 0.44 | -0.87 | -0.89 | 0.48 | -1.83 | -0.07 | -0.11 | -1.82 | -1.3 | 0.09 | 0.09 | -1.52 | -1.53 | 1.32 | 1.35 | -0.51 |
| 1.74 | -0.79 | -0.8 | 1.85 | -1.56 | -0.41 | -0.44 | -1.56 | -1.33 | 0.03 | 0.02 | -1.55 | -1.33 | -1.1 | -1.2 | -0.31 |
| 0.73 | -0.98 | -1.01 | 0.78 | -1.71 | -0.2 | -0.23 | -1.7 | -1.25 | 0.06 | 0.06 | -1.46 | -1.53 | 1.23 | 1.26 | -0.51 |
| 0.58 | -1.79 | -1.86 | 0.62 | -1.69 | -0.26 | -0.29 | -1.68 | -1.28 | -0.12 | -0.15 | -1.49 | -0.27 | 1.48 | 1.52 | 0.75 |
| 1.18 | -0.87 | -0.89 | 1.26 | -1.59 | -0.16 | -0.19 | -1.59 | -0.21 | 0.09 | 0.09 | -0.27 | -1.53 | 1.23 | 1.26 | -0.51 |
| 1.74 | -0.84 | -0.86 | 1.85 | -1.83 | -0.07 | -0.11 | -1.82 | -1.33 | 0.03 | 0.02 | -1.55 | -1.53 | 1.14 | 1.16 | -0.51 |
| 1.66 | -0.98 | -1.01 | 1.77 | -1.83 | -0.11 | -0.15 | -1.82 | -1.31 | 1.09 | 1.24 | -1.53 | -1.27 | 0.99 | 1 | -0.25 |
| 1.15 | 0.61 | 0.68 | 1.23 | -1.92 | -0.65 | -0.67 | -1.91 | -1.25 | 0.91 | 1.03 | -1.46 | -1.64 | -1.21 | -1.32 | -0.62 |
| -0.12 | 0.57 | 0.64 | -0.12 | -1.93 | -0.2 | -0.23 | -1.92 | -1.3 | 0.8 | 0.9 | -1.52 | -1.51 | 0.95 | 0.96 | -0.49 |
| 1.74 | -0.87 | -0.89 | 1.85 | -1.71 | 0.03 | -0.01 | -1.7 | -1.3 | 0.09 | 0.09 | -1.52 | -1.53 | 1.08 | 1.1 | -0.51 |
| 1.26 | 3.75 | 4 | 1.34 | -1.83 | 1.87 | 1.78 | -1.82 | -1.3 | 3.22 | 3.67 | -1.52 | -1.53 | 1.48 | 1.52 | -0.51 |
| 0.97 | -0.86 | -0.88 | 1.04 | -1.81 | -0.64 | -0.66 | -1.8 | -1.3 | 0.09 | 0.09 | -1.52 | -1.44 | 0.87 | 0.88 | -0.42 |
| 2.04 | 0.72 | 0.79 | 2.17 | -1.56 | 1.87 | 1.78 | -1.56 | -1.31 | 3.07 | 3.5 | -1.53 | -1.39 | 1.22 | 1.25 | -0.37 |
| 1.05 | -0.87 | -0.89 | 1.12 | -1.6 | 0.03 | -0.01 | -1.6 | -1.3 | 0.06 | 0.06 | -1.52 | -1.53 | 1.34 | 1.37 | -0.51 |


| 1.74 | -0.87 | -0.89 | 1.85 | -1.83 | -0.18 | -0.21 | -1.82 | -1.3 | -0.17 | -0.21 | -1.52 | -1.53 | 1.34 | 1.37 | -0.51 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1.53 | -0.98 | -1.01 | 1.63 | -1.83 | 0.03 | -0.01 | -1.82 | -1.3 | -0.03 | -0.05 | -1.52 | -1.57 | 1.34 | 1.37 | -0.55 |
| 1.59 | -0.48 | -0.48 | 1.69 | -0.69 | -0.03 | -0.07 | -0.71 | 0.48 | -0.17 | -0.21 | 0.52 | -0.14 | -0.19 | -0.24 | 0.88 |
| 2.02 | -0.41 | -0.4 | 2.15 | -0.58 | -0.18 | -0.21 | -0.6 | -0.22 | -0.3 | -0.35 | -0.28 | -0.28 | 0.4 | 0.38 | 0.74 |
| 0.84 | -0.59 | -0.59 | 0.9 | -1.47 | -0.53 | -0.55 | -1.47 | -2.43 | -0.24 | -0.29 | -2.81 | 0 | -0.29 | -0.35 | 1.02 |
| 1.66 | -0.81 | -0.83 | 1.77 | -0.18 | -0.25 | -0.28 | -0.21 | 0.42 | -0.6 | -0.7 | 0.45 | -0.6 | 0.21 | 0.18 | 0.42 |
| 1.78 | -0.29 | -0.28 | 1.89 | -1.91 | -1.54 | -1.54 | -1.9 | -1.05 | -0.34 | -0.4 | -1.23 | -0.36 | 0.4 | 0.38 | 0.66 |
| 1.57 | -0.48 | -0.48 | 1.67 | -0.62 | -0.19 | -0.22 | -0.64 | 0.04 | -0.3 | -0.36 | 0.02 | -0.43 | -0.3 | -0.36 | 0.59 |
| 1.78 | -0.29 | -0.28 | 1.89 | -1.96 | -1.54 | -1.54 | -1.95 | -1.39 | -0.34 | -0.4 | -1.62 | -1.53 | 0.4 | 0.38 | -0.51 |
| 1.66 | -0.81 | -0.83 | 1.77 | -0.18 | -0.25 | -0.28 | -0.21 | 0.42 | -0.6 | -0.7 | 0.45 | -0.6 | 0.21 | 0.18 | 0.42 |
| 2.13 | -0.41 | -0.4 | 2.26 | -2.61 | 0.02 | -0.02 | -2.58 | 0.24 | -0.43 | -0.5 | 0.24 | -0.45 | 0.4 | 0.38 | 0.57 |
| 1.74 | -0.56 | -0.56 | 1.85 | -0.5 | -0.86 | -0.87 | -0.53 | -0.34 | -0.23 | -0.27 | -0.42 | -0.29 | 1.56 | 1.61 | 0.73 |
| 1.27 | -0.26 | -0.24 | 1.35 | -1.5 | -0.84 | -0.86 | -1.5 | 0.24 | -0.1 | -0.13 | 0.24 | 0.05 | -0.05 | -0.09 | 1.07 |
| 1.66 | -0.46 | -0.45 | 1.77 | -0.18 | -2.48 | -2.45 | -0.21 | -0.28 | -0.67 | -0.78 | -0.35 | -1.51 | 0.06 | 0.02 | -0.49 |
| 1.58 | -0.34 | -0.33 | 1.68 | -1.09 | -1.36 | -1.36 | -1.1 | 0.25 | -0.17 | -0.21 | 0.26 | -0.53 | -0.12 | -0.17 | 0.49 |
| 1.74 | -0.63 | -0.63 | 1.85 | -0.16 | -0.9 | -0.91 | -0.2 | -1.26 | -0.58 | -0.67 | -1.47 | -0.89 | 0.21 | 0.18 | 0.13 |
| 1.59 | -0.91 | -0.93 | 1.69 | -0.28 | -1.39 | -1.39 | -0.31 | -0.06 | -0.67 | -0.78 | -0.1 | -1.32 | -0.76 | -0.85 | -0.3 |
| 1.54 | -0.51 | -0.51 | 1.64 | -1.09 | -0.31 | -0.34 | -1.1 | 0.25 | -1.23 | -1.42 | 0.26 | -0.53 | 0.34 | 0.32 | 0.49 |
| 1.74 | -0.66 | -0.67 | 1.85 | -0.17 | -0.75 | -0.77 | -0.21 | 0.02 | -0.44 | -0.52 | -0.01 | -1.45 | 0.24 | 0.21 | -0.43 |
| 0.66 | 0.59 | 0.66 | 0.71 | 0.87 | -0.19 | -0.22 | 0.81 | 0.42 | -0.22 | -0.26 | 0.45 | 0.98 | -0.29 | -0.35 | 2 |
| 0.25 | 0.85 | 0.93 | 0.27 | 0.76 | -1.65 | -1.64 | 0.7 | 0.59 | -0.09 | -0.12 | 0.64 | 0.87 | -1.12 | -1.22 | 1.89 |
| 1.27 | -0.49 | -0.49 | 1.35 | -1.48 | -0.38 | -0.41 | -1.48 | 0.51 | -0.23 | -0.28 | 0.55 | -0.08 | 0.56 | 0.55 | 0.94 |
| 0.09 | -0.73 | -0.74 | 0.11 | -0.28 | -1.36 | -1.36 | -0.31 | 0.69 | -0.59 | -0.69 | 0.76 | -0.7 | 0.24 | 0.21 | 0.32 |
| 0.73 | 0.88 | 0.96 | 0.78 | 0.47 | 0.58 | 0.52 | 0.42 | 1.01 | 0.68 | 0.77 | 1.13 | 1.42 | 0.49 | 0.47 | 2.44 |
| 0.62 | 0.67 | 0.74 | 0.67 | 0.33 | 0.6 | 0.54 | 0.28 | 0.92 | 0.68 | 0.77 | 1.02 | 1.34 | 0.36 | 0.34 | 2.36 |


| 1.77 | -0.29 | -0.28 | 1.88 | -0.06 | -1.39 | -1.39 | -0.1 | 0.52 | -0.72 | -0.84 | 0.56 | -0.49 | 0.4 | 0.38 | 0.53 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1.94 | -0.46 | -0.45 | 2.06 | -1.28 | -0.15 | -0.18 | -1.28 | 0.42 | -0.35 | -0.41 | 0.45 | -0.15 | 0.48 | 0.46 | 0.87 |
| 1.54 | -0.63 | -0.63 | 1.64 | -1.48 | 0.23 | 0.18 | -1.48 | 0.64 | -0.58 | -0.67 | 0.7 | -0.4 | -0.15 | -0.2 | 0.62 |
| 1.78 | -0.29 | -0.28 | 1.89 | -1.91 | -1.54 | -1.54 | -1.9 | -1.77 | -0.34 | -0.4 | -2.05 | -0.36 | 0.4 | 0.38 | 0.66 |
| 1.24 | 0.59 | 0.66 | 1.32 | 0.59 | 0.96 | 0.89 | 0.53 | 0.47 | -0.74 | -0.86 | 0.51 | 0.3 | 0.38 | 0.36 | 1.32 |
| 0.66 | 0.4 | 0.46 | 0.71 | 0.95 | 1.04 | 0.97 | 0.88 | -0.21 | -0.79 | -0.92 | -0.27 | 0.9 | -0.29 | -0.35 | 1.92 |
| 0.45 | 0.18 | 0.22 | 0.49 | 0.74 | 1.11 | 1.04 | 0.68 | 0.46 | 0.77 | 0.87 | 0.5 | -0.16 | 0.99 | 1 | 0.86 |
| 0.29 | 0.27 | 0.32 | 0.32 | 0.92 | 0.58 | 0.52 | 0.85 | 0.66 | 0.63 | 0.71 | 0.73 | 1.02 | 0.81 | 0.81 | 2.04 |
| -0.38 | -0.19 | -0.17 | -0.39 | 0.44 | 0.73 | 0.67 | 0.39 | -1.2 | 0.64 | 0.72 | -1.4 | 1.11 | -0.12 | -0.17 | 2.13 |
| 0.59 | 0.7 | 0.77 | 0.63 | 0.88 | 1.23 | 1.16 | 0.82 | 0.6 | 0.95 | 1.07 | 0.66 | 1.15 | 0.79 | 0.79 | 2.17 |
| 0.79 | 0.82 | 0.9 | 0.85 | 0.78 | 0.62 | 0.56 | 0.72 | 0.69 | 0.74 | 0.83 | 0.76 | -0.14 | 0.49 | 0.47 | 0.88 |
| 0.42 | 0.57 | 0.64 | 0.45 | 0.79 | -0.11 | -0.15 | 0.73 | 0.83 | -0.25 | -0.3 | 0.92 | 0.61 | 0.62 | 0.61 | 1.63 |
| 0.24 | 0.97 | 1.06 | 0.26 | 0.73 | 1.16 | 1.09 | 0.67 | 0.8 | 0.77 | 0.87 | 0.89 | 0.78 | 1.28 | 1.31 | 1.8 |
| 0.89 | -0.85 | -0.87 | 0.95 | -0.41 | -0.39 | -0.42 | -0.44 | 0.36 | -0.37 | -0.43 | 0.38 | -0.08 | -0.55 | -0.62 | 0.94 |
| 1.66 | -1.2 | -1.24 | 1.77 | -1.71 | -0.2 | -0.23 | -1.7 | -0.28 | -0.41 | -0.48 | -0.35 | -1.07 | 0.1 | 0.06 | -0.05 |
| 1.27 | -0.69 | -0.7 | 1.35 | -0.73 | -1.05 | -1.06 | -0.75 | 0.2 | -1.33 | -1.53 | 0.2 | -0.14 | -0.31 | -0.37 | 0.88 |
| 1.54 | -0.49 | -0.49 | 1.64 | -0.72 | -0.04 | -0.08 | -0.74 | 0.64 | -0.61 | -0.71 | 0.7 | -0.39 | 0.33 | 0.31 | 0.63 |
| 1.08 | -0.46 | -0.45 | 1.15 | -0.64 | -0.75 | -0.77 | -0.66 | 0.64 | -1 | -1.16 | 0.7 | -0.65 | 0.14 | 0.11 | 0.37 |
| -4 | -0.8 | -0.81 | -4.22 | -0.8 | -2.92 | -2.88 | -0.82 | -1.39 | -1.6 | -1.84 | -1.62 | -0.26 | 0.26 | 0.23 | 0.76 |
| 1.31 | -0.36 | -0.35 | 1.4 | -0.3 | -0.34 | -0.37 | -0.33 | 0.31 | -0.83 | -0.96 | 0.32 | 0.05 | 0 | -0.04 | 1.07 |
| 0.24 | -3.61 | -3.79 | 0.26 | -0.27 | -2.17 | -2.15 | -0.3 | 1.16 | 0.66 | 0.74 | 1.3 | -1.23 | 0.53 | 0.52 | -0.21 |
| 1.53 | -0.12 | -0.09 | 1.63 | 0.99 | 1.03 | 0.96 | 0.92 | 0.48 | 0.62 | 0.7 | 0.52 | 0.9 | 1.02 | 1.04 | 1.92 |
| -3.29 | 1.17 | 1.27 | -3.47 | -2.14 | 0.8 | 0.74 | -2.12 | -1.39 | 0.9 | 1.02 | -1.62 | -1.53 | -1.29 | -1.4 | -0.51 |
| 0.95 | 1.25 | 1.35 | 1.02 | 0.68 | 0.8 | 0.74 | 0.62 | 0.72 | 0.88 | 0.99 | 0.79 | 0.77 | 0.82 | 0.82 | 1.79 |
| 1.27 | -0.6 | -0.6 | 1.35 | -1.09 | -0.19 | -0.22 | -1.1 | 0.38 | -0.1 | -0.13 | 0.4 | -0.61 | 1.56 | 1.61 | 0.41 |


| 1.78 | -0.82 | -0.84 | 1.89 | -2.77 | 0.07 | 0.03 | -2.73 | -1.17 | -0.61 | -0.71 | -1.37 | -0.1 | -0.59 | -0.67 | 0.92 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.96 | -0.83 | -0.85 | 1.03 | -1.94 | -0.74 | -0.76 | -1.93 | -2.17 | -0.47 | -0.55 | -2.51 | -0.31 | -0.29 | -0.35 | 0.71 |
| 1.74 | -0.32 | -0.31 | 1.85 | -0.3 | -0.08 | -0.12 | -0.33 | -0.02 | -0.2 | -0.24 | -0.05 | -0.25 | 0.25 | 0.22 | 0.77 |
| 1.74 | -0.57 | -0.57 | 1.85 | -0.18 | -0.25 | -0.28 | -0.21 | -0.03 | -1.03 | -1.19 | -0.06 | -0.67 | 0.21 | 0.18 | 0.35 |
| 1.78 | -0.75 | -0.76 | 1.89 | -1.88 | -1.35 | -1.35 | -1.87 | -0.33 | -1.13 | -1.3 | -0.41 | -0.23 | 1.21 | 1.24 | 0.79 |
| 1.94 | -1.61 | -1.67 | 2.06 | -0.57 | -1.28 | -1.28 | -0.59 | 0.32 | -0.79 | -0.91 | 0.34 | -0.04 | 0.48 | 0.46 | 0.98 |
| 1.44 | -1.04 | -1.07 | 1.53 | -0.21 | -0.63 | -0.65 | -0.24 | -0.71 | -0.44 | -0.52 | -0.84 | -0.77 | -1.18 | -1.29 | 0.25 |
| 0.86 | -0.46 | -0.46 | 0.92 | -0.48 | -0.07 | -0.11 | -0.51 | -0.33 | -0.09 | -0.11 | -0.41 | 0.05 | -1.2 | -1.31 | 1.07 |
| 0.95 | 1.17 | 1.27 | 1.02 | 0.76 | 0.8 | 0.74 | 0.7 | 0.72 | 0.9 | 1.02 | 0.79 | 1.05 | 0.82 | 0.82 | 2.07 |
| 0.8 | -0.47 | -0.47 | 0.86 | -0.62 | -1.33 | -1.33 | -0.64 | 0.11 | -0.04 | -0.06 | 0.1 | -0.04 | -0.22 | -0.27 | 0.98 |
| 0.32 | -0.57 | -0.57 | 0.35 | -0.83 | -0.99 | -1 | -0.85 | -0.14 | -0.86 | -1 | -0.19 | -1.33 | 0.2 | 0.17 | -0.31 |
| 0.09 | -0.71 | -0.72 | 0.11 | -0.49 | -0.27 | -0.3 | -0.52 | -0.07 | -1.08 | -1.25 | -0.11 | -0.35 | 0.12 | 0.08 | 0.67 |
| -0.67 | -0.89 | -0.91 | -0.7 | -0.96 | 0.04 | 0 | -0.97 | 0.4 | 0.07 | 0.07 | 0.43 | -0.21 | 1.17 | 1.19 | 0.81 |
| 1.74 | -0.98 | -1.01 | 1.85 | -1.92 | -0.13 | -0.17 | -1.91 | 0.41 | -1.09 | -1.26 | 0.44 | -1.55 | 1.56 | 1.61 | -0.53 |
| 1.01 | 1.16 | 1.26 | 1.08 | -1.4 | -0.05 | -0.09 | -1.4 | -1.3 | -0.73 | -0.85 | -1.52 | -1.59 | 1.17 | 1.19 | -0.57 |
| 1.21 | -1.57 | -1.63 | 1.29 | -2.01 | -1.1 | -1.11 | -1.99 | 0.18 | -0.18 | -0.22 | 0.18 | -1.53 | -0.44 | -0.51 | -0.51 |
| 0.3 | 0.77 | 0.85 | 0.33 | 0.57 | 0.53 | 0.48 | 0.51 | 0.49 | 0.45 | 0.5 | 0.53 | -1.38 | 0.77 | 0.77 | -0.36 |
| 0.29 | 0.88 | 0.96 | 0.32 | 0.68 | 1 | 0.93 | 0.62 | 1.1 | 0.67 | 0.76 | 1.23 | 1.22 | -0.8 | -0.89 | 2.24 |
| -0.49 | -1.02 | -1.05 | -0.51 | -1.24 | -1.57 | -1.56 | -1.25 | -1.3 | 0.09 | 0.09 | -1.52 | -0.28 | -0.39 | -0.45 | 0.74 |
| 1.84 | 0.75 | 0.83 | 1.96 | -1.92 | -0.13 | -0.17 | -1.91 | 0.27 | 0.66 | 0.74 | 0.28 | -0.7 | 1.36 | 1.39 | 0.32 |
| 1.04 | 1.88 | 2.02 | 1.11 | -2.08 | -0.13 | -0.17 | -2.06 | -1.31 | 1.37 | 1.56 | -1.53 | -1.58 | 1.5 | 1.54 | -0.56 |
| 1.83 | 0.46 | 0.52 | 1.95 | -1.69 | 0.86 | 0.8 | -1.68 | 0.11 | -1.39 | -1.6 | 0.1 | 1.53 | 1.4 | 1.44 | 2.55 |
| 1.79 | 0.66 | 0.73 | 1.9 | -1.73 | -0.05 | -0.09 | -1.72 | -0.05 | 1.04 | 1.18 | -0.09 | 0.08 | 1.35 | 1.38 | 1.1 |
| 1.76 | 1.38 | 1.49 | 1.87 | -0.91 | -0.05 | -0.09 | -0.92 | -0.63 | 0.95 | 1.08 | -0.75 | -1.55 | 1.29 | 1.32 | -0.53 |
| 1.74 | -1.11 | -1.14 | 1.85 | -1.92 | 0.77 | 0.71 | -1.91 | -0.53 | 0.09 | 0.09 | -0.64 | -1.53 | 1.56 | 1.61 | -0.51 |


| 1.74 | 0.49 | 0.55 | 1.85 | -1.99 | -0.05 | -0.09 | -1.97 | 2.06 | 0.57 | 0.64 | 2.33 | -1.58 | 1.56 | 1.61 | -0.56 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2.04 | -0.98 | -1.01 | 2.17 | -1.6 | 0.79 | 0.73 | -1.6 | -1.3 | 0.03 | 0.02 | -1.52 | -1.39 | 1.14 | 1.16 | -0.37 |
| 1.11 | -0.98 | -1.01 | 1.18 | -1.89 | -0.5 | -0.52 | -1.88 | -1.31 | 0.09 | 0.09 | -1.53 | -1.47 | 0.95 | 0.96 | -0.45 |
| -3.89 | -0.98 | -1.01 | -4.11 | -0.75 | -0.25 | -0.28 | -0.77 | -1.3 | 0.01 | 0 | -1.52 | -1.53 | -0.66 | -0.74 | -0.51 |
| -1.87 | -1.22 | -1.26 | -1.97 | -1.68 | -3.08 | -3.03 | -1.67 | -1.3 | -1 | -1.15 | -1.52 | -0.88 | -1.19 | -1.3 | 0.14 |
| -0.58 | -0.83 | -0.85 | -0.6 | -1.17 | -0.15 | -0.18 | -1.18 | -1.29 | 0.02 | 0.01 | -1.51 | -0.5 | -0.38 | -0.44 | 0.52 |
| 0.21 | 0.48 | 0.54 | 0.23 | 0.45 | 1.23 | 1.16 | 0.4 | 0.81 | 0.22 | 0.24 | 0.9 | 0.93 | 0.52 | 0.51 | 1.95 |
| 0.3 | 0.67 | 0.74 | 0.33 | 0.57 | 0.68 | 0.62 | 0.51 | -0.14 | 0.05 | 0.05 | -0.19 | 0.87 | 0.86 | 0.87 | 1.89 |
| -0.43 | -0.22 | -0.2 | -0.44 | -0.65 | 0.12 | 0.08 | -0.67 | -1.28 | 0.09 | 0.09 | -1.49 | 0 | -1.51 | -1.64 | 1.02 |
| 2.21 | -0.34 | -0.33 | 2.35 | -1.22 | -0.07 | -0.11 | -1.23 | 0.93 | -1.19 | -1.37 | 1.03 | -0.11 | 1.37 | 1.4 | 0.91 |
| 1.79 | -0.98 | -1 | 1.9 | -1.83 | -1.57 | -1.56 | -1.82 | -0.62 | -1.03 | -1.19 | -0.74 | -0.77 | 1.5 | 1.54 | 0.25 |
| -0.7 | -0.6 | -0.6 | -0.73 | -0.36 | -0.95 | -0.96 | -0.39 | -1.3 | 0.31 | 0.34 | -1.52 | -0.71 | -1.31 | -1.43 | 0.31 |
| 0.49 | 0.76 | 0.84 | 0.53 | 0.67 | 1.02 | 0.95 | 0.61 | 0.46 | 0.25 | 0.27 | 0.5 | 0.96 | 0.71 | 0.71 | 1.98 |
| 2.56 | -0.71 | -0.72 | 2.72 | -2.3 | -0.62 | -0.64 | -2.28 | -0.13 | -0.56 | -0.65 | -0.18 | -0.23 | 0.24 | 0.21 | 0.79 |
| 2.28 | -1.25 | -1.29 | 2.42 | -2.01 | -1.45 | -1.45 | -1.99 | 4 | -0.29 | -0.34 | 4.55 | -0.16 | 1.11 | 1.13 | 0.86 |
| 1.79 | 0.46 | 0.52 | 1.9 | -1.73 | 0.8 | 0.74 | -1.72 | $-2.41$ | 1.31 | 1.49 | -2.79 | -1.61 | 1.38 | 1.42 | -0.59 |
| -0.87 | 0.86 | 0.94 | -0.91 | -0.37 | -1.23 | -1.23 | -0.4 | -1.3 | 0.09 | 0.09 | -1.52 | -1.34 | -0.87 | -0.96 | -0.32 |
| 1.74 | 1.19 | 1.29 | 1.85 | -1.92 | 0.75 | 0.69 | -1.91 | -1.33 | 0.09 | 0.09 | -1.55 | -1.03 | 1.46 | 1.5 | -0.01 |
| 1.79 | 0.74 | 0.81 | 1.9 | -1.85 | -0.05 | -0.09 | -1.84 | -0.48 | 0.87 | 0.98 | -0.58 | -1.53 | 1.56 | 1.61 | -0.51 |
| 1.74 | 1.47 | 1.59 | 1.85 | -1.92 | -0.05 | -0.09 | -1.91 | -1.21 | 0.86 | 0.97 | -1.41 | -1.26 | 1.33 | 1.36 | -0.24 |
| 1.18 | -1.29 | -1.33 | 1.26 | -1.8 | -0.22 | -0.25 | -1.79 | -0.03 | -0.51 | -0.59 | -0.06 | -1.53 | 1.18 | 1.2 | -0.51 |
| 1.4 | 2.12 | 2.27 | 1.49 | -1.82 | -0.5 | -0.52 | -1.81 | 0.07 | 0.09 | 0.09 | 0.05 | -0.02 | 0.93 | 0.94 | 1 |
| -0.92 | -0.98 | -1.01 | -0.96 | -0.49 | 0.31 | 0.26 | -0.52 | -1.3 | 0.02 | 0.01 | -1.52 | 0.14 | -0.92 | -1.01 | 1.16 |
| -0.33 | -0.69 | -0.7 | -0.34 | -1.09 | -0.58 | -0.6 | -1.1 | -1.3 | 0.18 | 0.19 | -1.52 | -0.85 | -0.36 | -0.42 | 0.17 |
| 0.7 | 0.64 | 0.71 | 0.75 | 0.6 | 1.17 | 1.1 | 0.54 | 1.6 | -0.49 | -0.57 | 1.8 | 1.42 | 0.88 | 0.89 | 2.44 |


| 1.74 | -0.52 | -0.52 | 1.85 | -1.94 | -1.07 | -1.08 | -1.93 | -0.71 | -3.38 | -3.88 | -0.84 | -1.53 | 0.03 | -0.01 | -0.51 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.32 | -0.78 | -0.79 | -0.33 | -1.79 | -0.05 | -0.09 | -1.78 | -0.88 | -0.22 | -0.26 | -1.04 | -1.61 | 1.41 | 1.45 | -0.59 |
| 1.79 | -0.09 | -0.06 | 1.9 | 0.27 | 0.9 | 0.84 | 0.22 | 1.01 | -0.75 | -0.87 | 1.13 | 0.14 | 0.16 | 0.13 | 1.16 |
| 1.35 | 0.31 | 0.36 | 1.44 | 0.47 | 1.07 | 1 | 0.42 | 0.79 | 0.32 | 0.36 | 0.87 | 0.88 | 0.36 | 0.34 | 1.9 |
| 0.47 | 0.51 | 0.57 | 0.51 | 0.86 | 0.51 | 0.46 | 0.8 | -1.3 | 0.02 | 0.01 | -1.52 | 1.12 | 0.74 | 0.74 | 2.14 |
| 0.04 | 0.45 | 0.51 | 0.05 | 0.75 | 0.53 | 0.48 | 0.69 | -1.29 | -1.17 | -1.35 | -1.51 | 1.3 | 0.55 | 0.54 | 2.32 |
| -0.96 | -0.91 | -0.93 | -1.01 | -1.12 | -2.04 | -2.02 | -1.13 | -1.68 | -0.03 | -0.05 | -1.95 | -0.96 | -0.86 | -0.95 | 0.06 |
| -0.99 | 0.04 | 0.07 | -1.04 | 0.65 | -0.02 | -0.06 | 0.59 | 0.1 | 0.05 | 0.05 | 0.08 | 0.23 | 0.13 | 0.09 | 1.25 |
| -1.29 | -0.37 | -0.36 | -1.35 | 0.3 | -0.07 | -0.11 | 0.25 | -0.36 | 0.25 | 0.28 | -0.44 | 0.07 | 0.38 | 0.36 | 1.09 |
| -1.16 | -0.88 | -0.9 | -1.22 | 0.34 | -0.72 | -0.74 | 0.29 | -0.3 | -0.64 | -0.74 | -0.37 | -1.53 | 0.18 | 0.15 | -0.51 |
| -0.7 | 0.04 | 0.07 | -0.73 | 0.17 | 0.78 | 0.72 | 0.13 | 0.13 | -0.67 | -0.78 | 0.12 | 0.24 | 0.13 | 0.09 | 1.26 |
| -0.86 | -0.14 | -0.12 | -0.9 | 0.24 | -0.59 | -0.61 | 0.19 | -0.02 | -1.61 | -1.85 | -0.05 | 0.22 | 0.16 | 0.13 | 1.24 |
| -0.93 | 0.07 | 0.11 | -0.97 | 0.31 | 0.24 | 0.19 | 0.26 | 0.24 | -1.09 | -1.26 | 0.24 | -0.23 | 0.01 | -0.03 | 0.79 |
| -0.06 | -0.74 | -0.75 | -0.05 | 0.34 | 0.05 | 0.01 | 0.29 | -0.02 | -1.14 | -1.32 | -0.05 | 0.52 | 0.54 | 0.53 | 1.54 |
| -0.69 | -0.29 | -0.27 | -0.72 | 0.27 | -0.08 | -0.12 | 0.22 | 0.3 | -2.26 | -2.6 | 0.31 | 0.01 | -0.76 | -0.84 | 1.03 |
| -1.16 | -0.04 | -0.01 | -1.22 | 0.65 | -0.02 | -0.06 | 0.59 | -0.11 | 0.05 | 0.05 | -0.16 | 0.24 | 0.13 | 0.09 | 1.26 |
| -0.99 | 0.05 | 0.08 | -1.04 | 0.63 | -0.02 | -0.06 | 0.57 | 0.1 | 0.09 | 0.09 | 0.08 | 0.31 | 0.13 | 0.09 | 1.33 |
| -0.65 | -0.74 | -0.75 | -0.68 | -0.62 | -2.92 | -2.88 | -0.64 | -1.56 | -2.07 | -2.38 | -1.81 | -0.07 | 0.35 | 0.33 | 0.95 |
| -0.9 | -0.39 | -0.38 | -0.94 | 0.47 | -0.16 | -0.19 | 0.42 | -0.36 | 0.25 | 0.28 | -0.44 | 0.2 | 0.13 | 0.09 | 1.22 |
| -1.03 | -0.07 | -0.04 | -1.08 | -0.93 | -0.12 | -0.16 | -0.94 | -0.21 | -1.21 | -1.4 | -0.27 | -1.29 | -1.52 | -1.65 | -0.27 |
| 2.18 | 0.14 | 0.18 | 2.32 | 0.2 | 1.02 | 0.95 | 0.15 | -0.79 | -0.37 | -0.44 | -0.93 | -0.07 | 0.67 | 0.67 | 0.95 |
| -0.69 | -0.85 | -0.87 | -0.72 | -1.73 | -0.44 | -0.47 | -1.72 | -0.38 | 0.35 | 0.39 | -0.46 | 0.04 | 0.13 | 0.09 | 1.06 |
| -1.16 | -0.26 | -0.24 | -1.22 | 0.64 | -0.26 | -0.29 | 0.58 | -0.27 | -0.09 | -0.11 | -0.34 | -1.11 | 0.13 | 0.09 | -0.09 |
| -0.06 | -0.75 | -0.76 | -0.05 | 0.34 | -1 | -1.01 | 0.29 | -0.19 | -1.11 | -1.28 | -0.25 | 0.52 | 0.49 | 0.48 | 1.54 |
| -1.16 | -0.38 | -0.37 | -1.22 | 0.3 | -0.55 | -0.57 | 0.25 | -0.51 | -0.55 | -0.64 | -0.61 | 0.52 | 0.13 | 0.1 | 1.54 |


| -0.9 | 0.05 | 0.08 | -0.94 | 0.23 | 0.76 | 0.7 | 0.18 | -0.2 | -1.8 | -2.07 | -0.26 | 0.29 | 0.16 | 0.13 | 1.31 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.9 | -0.28 | -0.26 | -0.94 | -0.02 | 0.65 | 0.59 | -0.06 | -0.21 | -1.8 | -2.07 | -0.27 | 0.29 | 0.14 | 0.11 | 1.31 |
| -1.16 | -0.29 | -0.27 | -1.22 | 0.35 | 0.04 | 0 | 0.3 | 0.32 | -0.16 | -0.2 | 0.34 | 0.14 | 0.18 | 0.15 | 1.16 |
| -0.69 | -0.29 | -0.27 | -0.72 | 0.48 | -0.08 | -0.12 | 0.43 | 0.32 | -1.07 | -1.24 | 0.34 | 0.07 | 0.31 | 0.29 | 1.09 |
| -0.9 | 0.03 | 0.06 | -0.94 | 0.23 | 0.78 | 0.72 | 0.18 | -1.27 | -0.24 | -0.29 | -1.48 | 0.29 | 0.16 | 0.13 | 1.31 |
| -1.16 | -0.29 | -0.27 | -1.22 | 0.34 | -0.08 | -0.12 | 0.29 | -0.4 | -0.72 | -0.84 | -0.49 | 0.09 | 0.18 | 0.15 | 1.11 |
| -0.69 | -0.37 | -0.36 | -0.72 | 0.48 | -0.07 | -0.11 | 0.43 | -0.36 | -0.62 | -0.72 | -0.44 | -0.07 | 0.35 | 0.33 | 0.95 |
| -1.13 | -0.35 | -0.34 | -1.19 | 0.62 | -0.07 | -0.11 | 0.56 | -0.63 | 0.15 | 0.16 | -0.75 | 0.28 | 0.13 | 0.09 | 1.3 |
| -1.37 | -0.37 | -0.36 | -1.44 | -0.31 | -0.07 | -0.11 | -0.34 | -0.37 | 0.24 | 0.26 | -0.45 | -1.48 | -1.04 | -1.14 | -0.46 |
| -0.95 | -0.35 | -0.34 | -1 | -0.78 | -0.02 | -0.06 | -0.8 | 0.28 | -0.64 | -0.74 | 0.29 | 0.29 | -1.04 | -1.14 | 1.31 |
| -0.71 | -1.05 | -1.08 | -0.74 | 0.18 | 0.7 | 0.64 | 0.13 | 0.28 | -1.57 | -1.81 | 0.29 | 0.29 | -1.04 | -1.14 | 1.31 |
| -0.95 | 0.04 | 0.07 | -1 | 0.36 | 0.6 | 0.54 | 0.31 | -0.01 | -0.87 | -1.01 | -0.04 | 0.34 | -1.14 | -1.25 | 1.36 |
| -0.06 | -0.14 | -0.12 | -0.05 | 0.37 | -0.83 | -0.85 | 0.32 | -1.39 | -0.86 | -0.99 | -1.62 | 0.22 | 0.55 | 0.54 | 1.24 |
| -1.16 | 0.07 | 0.11 | -1.22 | 0.49 | 0.24 | 0.19 | 0.44 | 0.24 | -0.97 | -1.12 | 0.24 | 0.29 | 0.13 | 0.09 | 1.31 |
| -1.16 | -0.13 | -0.11 | -1.22 | 0.61 | 0.18 | 0.14 | 0.55 | 0.25 | -1.57 | -1.81 | 0.26 | 0.16 | 0.19 | 0.16 | 1.18 |
| -1.16 | 0.07 | 0.11 | -1.22 | 0.65 | 0.24 | 0.19 | 0.59 | 0.25 | -1.57 | -1.81 | 0.26 | 0.29 | 0.13 | 0.09 | 1.31 |
| -1.15 | -0.55 | -0.55 | -1.21 | -0.32 | 0.24 | 0.19 | -0.35 | -1.39 | -1.13 | -1.3 | -1.62 | 0.29 | 0.15 | 0.12 | 1.31 |
| -1.16 | -0.81 | -0.83 | -1.22 | 0.65 | -3.17 | -3.12 | 0.59 | -0.18 | -0.62 | -0.72 | -0.24 | 0.24 | 0.13 | 0.09 | 1.26 |
| -1.16 | -0.46 | -0.46 | -1.22 | 0.65 | -0.31 | -0.34 | 0.59 | -0.41 | -0.03 | -0.05 | -0.5 | 0.27 | 0.13 | 0.09 | 1.29 |
| -0.7 | -1.3 | -1.34 | -0.73 | -1.87 | -0.42 | -0.45 | -1.86 | -2.14 | -1.15 | -1.33 | -2.48 | -1.21 | -0.29 | -0.35 | -0.19 |
| -0.38 | -0.78 | -0.79 | -0.39 | -1.27 | -0.65 | -0.67 | -1.27 | 0.21 | -0.05 | -0.07 | 0.21 | -0.58 | -1.14 | -1.25 | 0.44 |
| -3.29 | -0.91 | -0.93 | -3.47 | -2.14 | -0.35 | -0.38 | -2.12 | -1.39 | -0.33 | -0.39 | -1.62 | -1.53 | 0.13 | 0.1 | -0.51 |
| -1.4 | -0.88 | -0.9 | -1.47 | -0.88 | -0.86 | -0.87 | -0.9 | -0.82 | -0.22 | -0.26 | -0.97 | -0.4 | -0.29 | -0.35 | 0.62 |
| -0.54 | -0.57 | -0.57 | -0.56 | -0.6 | -0.71 | -0.73 | -0.62 | -1.93 | -0.74 | -0.86 | -2.24 | -0.24 | -1.15 | -1.26 | 0.78 |
| -0.66 | -0.81 | -0.83 | -0.69 | -0.37 | -0.26 | -0.29 | -0.4 | -2.16 | -1.11 | -1.28 | -2.5 | -0.96 | -0.88 | -0.97 | 0.06 |


| -3.29 | -0.88 | -0.9 | -3.47 | -2.14 | -0.65 | -0.67 | -2.12 | -1.39 | -0.16 | -0.19 | -1.62 | -1.53 | -0.95 | -1.05 | -0.51 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.54 | -1.01 | -1.04 | -0.56 | -0.91 | -1.87 | -1.86 | -0.92 | -0.4 | -1.48 | -1.7 | -0.49 | -1.29 | -0.39 | -0.45 | -0.27 |
| -0.89 | -1.07 | -1.1 | -0.93 | -1.19 | -0.97 | -0.98 | -1.2 | -0.27 | -0.44 | -0.52 | -0.34 | -0.07 | -3.5 | -3.74 | 0.95 |
| -0.7 | -1.06 | -1.09 | -0.73 | -1.63 | -0.05 | -0.09 | -1.62 | -0.28 | -0.58 | -0.67 | -0.35 | 0.14 | -0.77 | -0.86 | 1.16 |
| -0.74 | -0.78 | -0.79 | -0.77 | -1.42 | -1.14 | -1.15 | -1.42 | 0.09 | -0.39 | -0.46 | 0.07 | 0.1 | -0.7 | -0.78 | 1.12 |
| 0.05 | -0.87 | -0.89 | 0.06 | -0.95 | -0.65 | -0.67 | -0.96 | 0.68 | -0.24 | -0.29 | 0.75 | -1.28 | -0.95 | -1.05 | -0.26 |
| -0.84 | -0.82 | -0.84 | -0.88 | -2.64 | -1.21 | -1.21 | -2.61 | 0.16 | -0.05 | -0.07 | 0.15 | -0.53 | -0.42 | -0.49 | 0.49 |
| -1.1 | -0.73 | -0.74 | -1.15 | -0.64 | -0.72 | -0.74 | -0.66 | -0.21 | -0.71 | -0.82 | -0.27 | -0.31 | -0.45 | -0.52 | 0.71 |
| -1.35 | -0.41 | -0.4 | -1.42 | -1.01 | -0.01 | -0.05 | -1.02 | -1.6 | -0.9 | -1.04 | -1.86 | -0.24 | -0.27 | -0.33 | 0.78 |
| -1.75 | -1.01 | -1.04 | -1.84 | -2.03 | -0.83 | -0.85 | -2.01 | 0.23 | -0.62 | -0.72 | 0.23 | 0.03 | -1.17 | -1.28 | 1.05 |
| -2.75 | -1.11 | -1.14 | -2.9 | -1.86 | -2.14 | -2.12 | -1.85 | 0.24 | 0.01 | 0 | 0.24 | -0.29 | -0.59 | -0.66 | 0.73 |
| -1.43 | -0.63 | -0.63 | -1.5 | -0.31 | -0.75 | -0.77 | -0.34 | -0.58 | -0.98 | -1.13 | -0.69 | -1.04 | -0.34 | -0.4 | -0.02 |
| -0.21 | -0.68 | -0.69 | -0.21 | -0.86 | -1.21 | -1.21 | -0.88 | -0.37 | -0.65 | -0.76 | -0.45 | -0.7 | -0.63 | -0.71 | 0.32 |
| 0.04 | -0.56 | -0.56 | 0.05 | -0.98 | -2.4 | -2.37 | -0.99 | 0.25 | -1.12 | -1.29 | 0.26 | -0.23 | -1.22 | -1.33 | 0.79 |
| -1.23 | -0.4 | -0.39 | -1.29 | -1.05 | -1.37 | -1.37 | -1.06 | -1.04 | 0.14 | 0.15 | -1.22 | -0.34 | -0.76 | -0.85 | 0.68 |
| -1.61 | -0.48 | -0.48 | -1.69 | -1.89 | 0.03 | -0.01 | -1.88 | -1.31 | -0.96 | -1.11 | -1.53 | -0.51 | -0.23 | -0.28 | 0.51 |
| -0.71 | -0.8 | -0.82 | -0.74 | -0.37 | 0 | -0.04 | -0.4 | -1.8 | -0.13 | -0.16 | -2.09 | -0.09 | -1.16 | -1.27 | 0.93 |
| -3.29 | -0.75 | -0.76 | -3.47 | -0.99 | -1.54 | -1.54 | -1 | -2.12 | -0.52 | -0.61 | -2.46 | -0.21 | -1.2 | -1.31 | 0.81 |
| -0.83 | -0.8 | -0.82 | -0.87 | -1.18 | -1.13 | -1.14 | -1.19 | -0.93 | -0.27 | -0.32 | -1.09 | -0.31 | -0.76 | -0.85 | 0.71 |
| -0.08 | -0.72 | -0.73 | -0.07 | -0.72 | -0.63 | -0.65 | -0.74 | -0.25 | -0.5 | -0.58 | -0.32 | -0.48 | -0.37 | -0.43 | 0.54 |
| -0.94 | -0.73 | -0.74 | -0.98 | -0.9 | -1.59 | -1.58 | -0.91 | -0.38 | -0.24 | -0.29 | -0.46 | -0.77 | -1.87 | -2.02 | 0.25 |
| -1.11 | -1.23 | -1.27 | -1.16 | -0.78 | -0.86 | -0.87 | -0.8 | 0.14 | -0.09 | -0.11 | 0.13 | -0.34 | -1.15 | -1.26 | 0.68 |
| -1.23 | -0.59 | -0.59 | -1.29 | -1.87 | -1.14 | -1.15 | -1.86 | -0.79 | -0.15 | -0.18 | -0.93 | -1.28 | -1.12 | -1.23 | -0.26 |
| -1.15 | -0.45 | -0.44 | -1.21 | -2.17 | -0.21 | -0.24 | -2.15 | -1.51 | -0.02 | -0.04 | -1.76 | -0.4 | -0.83 | -0.92 | 0.62 |
| -0.87 | -0.92 | -0.94 | -0.91 | -1.79 | -0.23 | -0.26 | -1.78 | -0.48 | -0.25 | -0.3 | -0.58 | -1.32 | -0.46 | -0.53 | -0.3 |


| -0.8 | -0.9 | -0.92 | -0.84 | -1.59 | -0.01 | -0.05 | -1.59 | -0.15 | -1.77 | -2.04 | -0.2 | -0.87 | -1.17 | -1.28 | 0.15 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -1.09 | -0.44 | -0.43 | -1.14 | -0.82 | -1.07 | -1.08 | -0.84 | -1.72 | -0.51 | -0.59 | -2 | -0.64 | -1.18 | -1.29 | 0.38 |
| -1.24 | -0.8 | -0.81 | -1.3 | -1 | -0.66 | -0.68 | -1.01 | -2.73 | -2.57 | -2.95 | -3.15 | -0.71 | -2.77 | -2.97 | 0.31 |
| -1.58 | -0.62 | -0.62 | -1.66 | -0.05 | -2.17 | -2.15 | -0.09 | -2.36 | -0.23 | -0.27 | -2.73 | -0.09 | -1 | -1.1 | 0.93 |
| -2.6 | -0.38 | -0.37 | -2.74 | -0.62 | -0.77 | -0.79 | -0.64 | -0.26 | -1.31 | -1.51 | -0.33 | 0.09 | -1.35 | -1.47 | 1.11 |
| -0.2 | -0.69 | -0.7 | -0.2 | -1.09 | -1.47 | -1.47 | -1.1 | -0.83 | -0.33 | -0.39 | -0.98 | -1.04 | -0.67 | -0.75 | -0.02 |
| -1.3 | -1.24 | -1.28 | -1.37 | -1.04 | -0.35 | -0.38 | -1.05 | -1.56 | -0.47 | -0.55 | -1.81 | -1.28 | -1.13 | -1.24 | -0.26 |
| -2.45 | -2.75 | -2.88 | -2.58 | -1.91 | -0.96 | -0.97 | -1.9 | -1.07 | -0.76 | -0.88 | -1.25 | -0.96 | -0.41 | -0.47 | 0.06 |
| -0.8 | -1.17 | -1.21 | -0.84 | -1.87 | -2.23 | -2.21 | -1.86 | -1.06 | -1.11 | -1.28 | -1.24 | -0.28 | -0.47 | -0.54 | 0.74 |
| -2.2 | -0.8 | -0.81 | -2.32 | -0.73 | -1.1 | -1.11 | -0.75 | -1.9 | -0.62 | -0.72 | -2.2 | -0.96 | -0.53 | -0.6 | 0.06 |
| -1.59 | -0.65 | -0.66 | -1.67 | -0.66 | -0.66 | -0.68 | -0.68 | 0.2 | -0.85 | -0.98 | 0.2 | -0.37 | -1.01 | -1.11 | 0.65 |
| -0.59 | -0.98 | -1.01 | -0.61 | -0.52 | -2.24 | -2.22 | -0.55 | 0 | -2.13 | -2.45 | -0.03 | 0.1 | -1.14 | -1.25 | 1.12 |
| -0.82 | -0.86 | -0.88 | -0.86 | -0.92 | -1.3 | -1.3 | -0.93 | 0.05 | -1.69 | -1.95 | 0.03 | -0.11 | -0.96 | -1.06 | 0.91 |
| -1.01 | -1.05 | -1.08 | -1.06 | -0.82 | -0.65 | -0.67 | -0.84 | -1.14 | -1.1 | -1.27 | -1.33 | -0.82 | -1.12 | -1.22 | 0.2 |
| -0.35 | -2.75 | -2.88 | -0.36 | -0.7 | -0.43 | -0.46 | -0.72 | -2.65 | -0.84 | -0.97 | -3.06 | -0.03 | -1.17 | -1.28 | 0.99 |
| -3.29 | -0.9 | -0.92 | -3.47 | -2.14 | -1.54 | -1.54 | -2.12 | -1.39 | -1.21 | -1.4 | -1.62 | -1.53 | -1.29 | -1.4 | -0.51 |
| -1.45 | -1.09 | -1.12 | -1.52 | 0.05 | -0.48 | -0.51 | 0.01 | -1.71 | -0.44 | -0.52 | -1.99 | -0.64 | -2.18 | -2.34 | 0.38 |
| -2.81 | -0.92 | -0.94 | -2.96 | -0.76 | -0.78 | -0.8 | -0.78 | -0.31 | -0.37 | -0.43 | -0.38 | -0.29 | -0.74 | -0.82 | 0.73 |
| -0.93 | -0.8 | -0.81 | -0.97 | -1.47 | -0.51 | -0.53 | -1.47 | 0.46 | -0.1 | -0.13 | 0.5 | -1.31 | -1.17 | -1.28 | -0.29 |
| -1.16 | -0.59 | -0.59 | -1.22 | -1.43 | -0.56 | -0.58 | -1.43 | 0.72 | -1.51 | -1.74 | 0.79 | 1.68 | -0.36 | -0.42 | 2.7 |
| 0.53 | -0.8 | -0.82 | 0.57 | 0.61 | -1.09 | -1.1 | 0.55 | 1.13 | 1.2 | 1.36 | 1.26 | 1.01 | 0.72 | 0.72 | 2.03 |
| 0.31 | -0.72 | -0.73 | 0.34 | -1.55 | -2.31 | -2.28 | -1.55 | -0.16 | -0.35 | -0.41 | -0.21 | -0.07 | -0.15 | -0.2 | 0.95 |
| -1.1 | -0.81 | -0.83 | -1.15 | -0.63 | -0.37 | -0.4 | -0.65 | 0.21 | -1.11 | -1.28 | 0.21 | -0.97 | -1.11 | -1.21 | 0.05 |
| -1.13 | -1.19 | -1.23 | -1.19 | -1.86 | -0.45 | -0.48 | -1.85 | 0.19 | -0.34 | -0.4 | 0.19 | -0.19 | 0.27 | 0.24 | 0.83 |
| -0.68 | -1.07 | -1.1 | -0.71 | -1.17 | -0.43 | -0.46 | -1.18 | -0.21 | -1.07 | -1.24 | -0.27 | -0.73 | -0.68 | -0.76 | 0.29 |


| -0.47 | -0.67 | -0.68 | -0.49 | -1.43 | -0.58 | -0.6 | -1.43 | 0.38 | -0.16 | -0.19 | 0.4 | -0.35 | -1.16 | -1.27 | 0.67 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -1.46 | -2.75 | -2.88 | -1.53 | -0.7 | -0.18 | -0.21 | -0.72 | -2.65 | -0.79 | -0.92 | -3.06 | -0.03 | -1.16 | -1.27 | 0.99 |
| -1.91 | -0.88 | -0.9 | -2.01 | -0.7 | -0.65 | -0.67 | -0.72 | -2.65 | -1.62 | -1.86 | -3.06 | -0.03 | -0.95 | -1.05 | 0.99 |
| -3.26 | -0.54 | -0.54 | -3.44 | -1.86 | -0.52 | -0.54 | -1.85 | -1.31 | -0.34 | -0.4 | -1.53 | -0.41 | -0.21 | -0.26 | 0.61 |
| -1.79 | -0.53 | -0.53 | -1.88 | -0.34 | 0.13 | 0.09 | -0.37 | -0.82 | -0.86 | -1 | -0.97 | -1.38 | -0.93 | -1.02 | -0.36 |
| -1.02 | -0.98 | -1.01 | -1.07 | -0.71 | -0.82 | -0.84 | -0.73 | -2.06 | 0.21 | 0.23 | -2.39 | -0.52 | -3.08 | -3.29 | 0.5 |
| 0.05 | -1.75 | -1.82 | 0.06 | -0.44 | -0.15 | -0.18 | -0.47 | 0.15 | -1.09 | -1.26 | 0.14 | -0.53 | -1.18 | -1.29 | 0.49 |
| -1.25 | -0.6 | -0.6 | -1.31 | -0.54 | -1.02 | -1.03 | -0.56 | -1.07 | -0.37 | -0.44 | -1.25 | -0.96 | -0.95 | -1.05 | 0.06 |
| -1.64 | -0.84 | -0.86 | -1.73 | -1.86 | -1.1 | -1.11 | -1.85 | -0.4 | -0.36 | -0.42 | -0.49 | -0.24 | -0.17 | -0.22 | 0.78 |
| -0.54 | -0.35 | -0.34 | -0.56 | -0.88 | -0.9 | -0.91 | -0.9 | -1.38 | -1.08 | -1.25 | -1.61 | -0.24 | -1.17 | -1.28 | 0.78 |
| -1.03 | -0.89 | -0.91 | -1.08 | -0.86 | -0.56 | -0.58 | -0.88 | -2.02 | -0.23 | -0.28 | -2.34 | -1.73 | -0.86 | -0.95 | -0.71 |
| -0.57 | -1.13 | -1.16 | -0.59 | -0.49 | -0.42 | -0.45 | -0.52 | -0.27 | 0.09 | 0.09 | -0.34 | -0.12 | -1.19 | -1.3 | 0.9 |
| -1.3 | -0.63 | -0.64 | -1.37 | -1.86 | -0.17 | -0.2 | -1.85 | -0.83 | -0.54 | -0.63 | -0.98 | -0.2 | -1.22 | -1.33 | 0.82 |
| -1.31 | -0.88 | -0.9 | -1.38 | -0.82 | -1.34 | -1.34 | -0.84 | -1.34 | -0.31 | -0.37 | -1.56 | -0.3 | -1.21 | -1.32 | 0.72 |
| -1.32 | -0.5 | -0.5 | -1.39 | -2.06 | -0.06 | -0.1 | -2.04 | -0.63 | -0.5 | -0.58 | -0.75 | -0.71 | -0.65 | -0.73 | 0.31 |
| -0.82 | -1.43 | -1.48 | -0.86 | -1.28 | -0.65 | -0.67 | -1.28 | -2.65 | -0.16 | -0.19 | -3.06 | -0.29 | -0.95 | -1.05 | 0.73 |
| 0.47 | 0.29 | 0.34 | 0.51 | 0.69 | 0.46 | 0.41 | 0.63 | 0.22 | 0.33 | 0.37 | 0.22 | 1.18 | 0.63 | 0.62 | 2.2 |
| 0.5 | -0.22 | -0.2 | 0.54 | -0.15 | 0.66 | 0.6 | -0.19 | 0.96 | 0.88 | 1 | 1.07 | 1.63 | -0.3 | -0.36 | 2.65 |
| -0.14 | -0.06 | -0.03 | -0.14 | -0.63 | 0.52 | 0.47 | -0.65 | 0.73 | -0.52 | -0.61 | 0.81 | 0.81 | -1.02 | -1.12 | 1.83 |
| 0.07 | 0.16 | 0.2 | 0.08 | -0.07 | 0.27 | 0.22 | -0.11 | -0.51 | -2.18 | -2.5 | -0.61 | -0.69 | -1.12 | -1.22 | 0.33 |
| -0.72 | 0.08 | 0.12 | -0.75 | 0.31 | 1.16 | 1.09 | 0.26 | -0.37 | 0.52 | 0.58 | -0.45 | 1.12 | 0.47 | 0.45 | 2.14 |
| 0.06 | 0.35 | 0.4 | 0.07 | 0.76 | -0.58 | -0.6 | 0.7 | -0.7 | 0.69 | 0.78 | -0.83 | 1.28 | 0.9 | 0.91 | 2.3 |
| 0.58 | 0.31 | 0.36 | 0.62 | -0.1 | 0.78 | 0.72 | -0.14 | 0.8 | 0.69 | 0.78 | 0.89 | 0.32 | 0.52 | 0.51 | 1.34 |
| 1.11 | -0.44 | -0.43 | 1.18 | 0.73 | 1.15 | 1.08 | 0.67 | 0.72 | 0.68 | 0.77 | 0.79 | 0.57 | 0.48 | 0.46 | 1.59 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |


| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.61 | 1.49 | 1.53 | 1.63 |
| 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
| -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 0.98 | 0.93 | 0.94 | 2 |
| -0.87 | 0.34 | 0.39 | -0.91 | -0.18 | 0.59 | 0.53 | -0.21 | -1.03 | 0.91 | 1.03 | -1.21 | 0.32 | 0.74 | 0.74 | 1.34 |
| 0.15 | 0.4 | 0.45 | 0.17 | 0.35 | -0.46 | -0.49 | 0.3 | -0.16 | -0.15 | -0.18 | -0.21 | 0.42 | -1.11 | -1.21 | 1.44 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.23 | -0.09 | -0.12 | 1.38 | 1.08 | -1.12 | -1.22 | 2.1 |
| 0.31 | 0.79 | 0.87 | 0.34 | 0.38 | -1.38 | -1.38 | 0.33 | 1.05 | -0.09 | -0.12 | 1.17 | 0.83 | -1.12 | -1.22 | 1.85 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.21 | -0.09 | -0.12 | 1.35 | 0.8 | -0.8 | -0.89 | 1.82 |
| -1.36 | -0.8 | -0.82 | -1.43 | -0.9 | -1.36 | -1.36 | -0.91 | 0.46 | -1.13 | -1.3 | 0.5 | -0.1 | -1.19 | -1.3 | 0.92 |
| -1.87 | -0.98 | -1 | -1.97 | -0.5 | -0.41 | -0.44 | -0.53 | -0.87 | -0.89 | -1.03 | -1.03 | -0.87 | -1.3 | -1.41 | 0.15 |
| -0.48 | -1.19 | -1.23 | -0.5 | -0.56 | -1.17 | -1.18 | -0.58 | 0.3 | -1.08 | -1.25 | 0.31 | -1.36 | -0.22 | -0.27 | -0.34 |
| -1.89 | -0.7 | -0.71 | -1.99 | -2.73 | -1.39 | -1.39 | -2.69 | -1.56 | -0.89 | -1.03 | -1.81 | -0.86 | -0.57 | -0.64 | 0.16 |
| -1.82 | -0.59 | -0.59 | -1.92 | -0.71 | -1.42 | -1.42 | -0.73 | -1.84 | -0.73 | -0.85 | -2.13 | -0.21 | -1.38 | -1.5 | 0.81 |
| -1.41 | -1.09 | -1.12 | -1.48 | -0.61 | -1.53 | -1.53 | -0.63 | -1.65 | -1.11 | -1.28 | -1.92 | -1.96 | -1.17 | -1.28 | -0.94 |
| -1.78 | -0.43 | -0.42 | -1.87 | -0.74 | -0.81 | -0.83 | -0.76 | -1.2 | -0.72 | -0.83 | -1.4 | -0.35 | -1.2 | -1.31 | 0.67 |
| -0.73 | -0.95 | -0.97 | -0.76 | -0.25 | -1.13 | -1.14 | -0.28 | -1.65 | -1.8 | -2.07 | -1.92 | -0.66 | -1.19 | -1.3 | 0.36 |
| 0.91 | 1.09 | 1.18 | 0.97 | 0.18 | 0.61 | 0.55 | 0.13 | 0.37 | 0.78 | 0.88 | 0.39 | 0.99 | 1.18 | 1.2 | 2.01 |
| 0.65 | 1.15 | 1.25 | 0.7 | 0.16 | 1.24 | 1.17 | 0.12 | 1.57 | 1.35 | 1.53 | 1.77 | 0.89 | 0.75 | 0.75 | 1.91 |
| 0.67 | 1.55 | 1.67 | 0.72 | 1.13 | 1.24 | 1.17 | 1.06 | 1.58 | 1 | 1.13 | 1.78 | 1.4 | 0.89 | 0.9 | 2.42 |
| 0.04 | -0.04 | -0.01 | 0.05 | 0.41 | -0.27 | -0.3 | 0.36 | 0.99 | 0.25 | 0.27 | 1.1 | 1.32 | -0.83 | -0.92 | 2.34 |
| 0.93 | 1.2 | 1.3 | 0.99 | 0.76 | -0.24 | -0.27 | 0.7 | 1.08 | 0.91 | 1.03 | 1.21 | 0.9 | 1.02 | 1.04 | 1.92 |


| 0.1 | 0.83 | 0.91 | 0.12 | 1.58 | 1.62 | 1.54 | 1.5 | 1.4 | 1.2 | 1.36 | 1.57 | 1.55 | 1.02 | 1.03 | 2.57 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.3 | 0.91 | 0.99 | 0.33 | 0.65 | 0.73 | 0.67 | 0.59 | 0.27 | 0.71 | 0.8 | 0.28 | 0.48 | 0.9 | 0.91 | 1.5 |
| 1.75 | 0.58 | 0.65 | 1.86 | -0.15 | 0.78 | 0.72 | -0.19 | 0.39 | 0.59 | 0.66 | 0.42 | 1.39 | -1.18 | -1.29 | 2.41 |
| 1.85 | 0.53 | 0.59 | 1.97 | 0.34 | 1.14 | 1.07 | 0.29 | 0.82 | 0.98 | 1.11 | 0.91 | 0.58 | 0.75 | 0.75 | 1.6 |
| -0.18 | 1.85 | 1.99 | -0.18 | 1.08 | 1.21 | 1.14 | 1.01 | 1.54 | -0.51 | -0.59 | 1.73 | 1.35 | 1.28 | 1.31 | 2.37 |
| 0.47 | 1 | 1.09 | 0.51 | 0.63 | 2.63 | 2.52 | 0.57 | 0.36 | 0.95 | 1.07 | 0.38 | 0.73 | 0.84 | 0.84 | 1.75 |
| 0.09 | 0.99 | 1.08 | 0.11 | 1.13 | 0.65 | 0.59 | 1.06 | 1.4 | 0.74 | 0.84 | 1.57 | 1.21 | 0.74 | 0.74 | 2.23 |
| -0.08 | 1.52 | 1.64 | -0.07 | 0.67 | 0.61 | 0.55 | 0.61 | 0.75 | 0.72 | 0.81 | 0.83 | 0.7 | 0.81 | 0.81 | 1.72 |
| 0.37 | 0.74 | 0.81 | 0.4 | 0.7 | 0.28 | 0.23 | 0.64 | 0.86 | 1.16 | 1.32 | 0.95 | 0.83 | 0.8 | 0.8 | 1.85 |
| 0.02 | 0.7 | 0.77 | 0.03 | 0.69 | -0.84 | -0.86 | 0.63 | 0.64 | 0.78 | 0.88 | 0.7 | 0.86 | 0.28 | 0.25 | 1.88 |
| 0.62 | 0.51 | 0.57 | 0.67 | 0.57 | 0.77 | 0.71 | 0.51 | 0.93 | 0.47 | 0.53 | 1.03 | 0.97 | 0.11 | 0.07 | 1.99 |
| -0.45 | -0.34 | -0.33 | -0.47 | 0.61 | -0.77 | -0.79 | 0.55 | 1.1 | 0.66 | 0.74 | 1.23 | 1.14 | 0.62 | 0.61 | 2.16 |
| 0.63 | 0.56 | 0.62 | 0.68 | 1.05 | -1.79 | -1.78 | 0.98 | 0.94 | 0.99 | 1.12 | 1.05 | 0.41 | -0.35 | -0.41 | 1.43 |
| 0.47 | 0.92 | 1.01 | 0.51 | 0.62 | 1.21 | 1.14 | 0.56 | 0.77 | 0.91 | 1.03 | 0.85 | 0.53 | 0.85 | 0.86 | 1.55 |
| 0.92 | 0.75 | 0.83 | 0.98 | 0.48 | 0.51 | 0.46 | 0.43 | 0.78 | 0.9 | 1.02 | 0.86 | 0.55 | 0.65 | 0.64 | 1.57 |
| -0.15 | 0.26 | 0.31 | -0.15 | 1.32 | 0.89 | 0.83 | 1.24 | 0.87 | 1.18 | 1.34 | 0.97 | 0.69 | 0.96 | 0.97 | 1.71 |
| 0.08 | 0.97 | 1.06 | 0.09 | 0.08 | 0.86 | 0.8 | 0.04 | 0.67 | 0.47 | 0.53 | 0.74 | -0.32 | 0.49 | 0.47 | 0.7 |
| 0.7 | 0.72 | 0.79 | 0.75 | 0.81 | 0.89 | 0.83 | 0.75 | 0.68 | 0.53 | 0.6 | 0.75 | 0.49 | 0.72 | 0.72 | 1.51 |
| 0.45 | 0.07 | 0.11 | 0.49 | 0.57 | -0.41 | -0.44 | 0.51 | 0.36 | 0.85 | 0.96 | 0.38 | 0.68 | 0.1 | 0.06 | 1.7 |
| 1.31 | 3.19 | 3.41 | 1.4 | 0.05 | 0.71 | 0.65 | 0.01 | 1.61 | 0.61 | 0.69 | 1.81 | 1.72 | 0.88 | 0.89 | 2.74 |
| 0.92 | 0.78 | 0.86 | 0.98 | 1.29 | 1.11 | 1.04 | 1.21 | 1.17 | 0.73 | 0.82 | 1.31 | 1.34 | 0.92 | 0.93 | 2.36 |
| -0.55 | -0.05 | -0.02 | -0.57 | 0.49 | 0 | -0.04 | 0.44 | -0.31 | 0.75 | 0.85 | -0.38 | 0.49 | 0.48 | 0.46 | 1.51 |
| 0.45 | -0.12 | -0.1 | 0.49 | 0.5 | 0.33 | 0.28 | 0.45 | 0.92 | 0.95 | 1.08 | 1.02 | -0.1 | 0.34 | 0.32 | 0.92 |
| -0.22 | 0.27 | 0.32 | -0.22 | -0.93 | 0.23 | 0.18 | -0.94 | 0.26 | 0.47 | 0.53 | 0.27 | -0.2 | 0.37 | 0.35 | 0.82 |
| -0.45 | -0.25 | -0.23 | -0.47 | -0.5 | 0.94 | 0.87 | -0.53 | -0.85 | 0.7 | 0.79 | -1 | 0.53 | 0.24 | 0.21 | 1.55 |


| 0.14 | 0.52 | 0.58 | 0.16 | 0.33 | 1.55 | 1.47 | 0.28 | 0.77 | 0.6 | 0.68 | 0.85 | 0.48 | 0.8 | 0.8 | 1.5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.02 | -1.17 | -1.28 | 2.04 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.56 | 0.29 | 0.26 | 1.58 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.61 | 0.58 | 0.57 | 1.63 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.56 | 1.49 | 1.53 | 1.58 |
| 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
| -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 1.02 | 0.93 | 0.94 | 2.04 |
| -0.87 | 0.34 | 0.39 | -0.91 | -0.18 | 0.59 | 0.53 | -0.21 | -1.03 | 0.91 | 1.03 | -1.21 | 0.32 | 0.74 | 0.74 | 1.34 |
| 0.15 | 0.4 | 0.45 | 0.17 | 0.35 | -0.46 | -0.49 | 0.3 | -0.16 | -0.15 | -0.18 | -0.21 | 0.42 | -1.11 | -1.21 | 1.44 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.23 | -0.09 | -0.12 | 1.38 | 1.03 | -1.12 | -1.22 | 2.05 |
| 0.31 | 0.79 | 0.87 | 0.34 | 0.38 | -1.38 | -1.38 | 0.33 | 1.05 | -0.09 | -0.12 | 1.17 | 0.79 | -1.12 | -1.22 | 1.81 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.21 | -0.09 | -0.12 | 1.35 | 0.76 | -0.8 | -0.89 | 1.78 |
| -1.36 | -0.8 | -0.82 | -1.43 | -0.9 | -1.36 | -1.36 | -0.91 | 0.46 | -1.13 | -1.3 | 0.5 | -0.1 | -1.19 | -1.3 | 0.92 |
| -1.87 | -0.98 | -1 | -1.97 | -0.5 | -0.41 | -0.44 | -0.53 | -0.87 | -0.89 | -1.03 | -1.03 | -0.87 | -1.3 | -1.41 | 0.15 |
| -0.48 | -1.19 | -1.23 | -0.5 | -0.56 | -1.17 | -1.18 | -0.58 | 0.3 | -1.08 | -1.25 | 0.31 | -0.96 | -0.22 | -0.27 | 0.06 |
| -1.89 | -0.7 | -0.71 | -1.99 | -2.73 | -1.39 | -1.39 | -2.69 | -1.56 | -0.89 | -1.03 | -1.81 | -0.86 | -0.57 | -0.64 | 0.16 |
| 0.1 | -0.32 | -0.31 | 0.12 | -0.12 | -1.07 | -1.08 | -0.16 | 0.25 | -3.14 | -3.6 | 0.26 | -1.12 | -0.9 | -0.99 | -0.1 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.02 | -1.17 | -1.28 | 2.04 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.71 | 0.67 | 0.66 | 1.73 |
| -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.56 | 1.49 | 1.53 | 1.58 |


| 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 0.98 | 0.93 | 0.94 | 2 |
| -0.87 | 0.34 | 0.39 | -0.91 | -0.18 | 0.59 | 0.53 | -0.21 | -1.03 | 0.91 | 1.03 | -1.21 | 0.38 | 0.74 | 0.74 | 1.4 |
| 0.15 | 0.4 | 0.45 | 0.17 | 0.35 | -0.46 | -0.49 | 0.3 | -0.16 | -0.15 | -0.18 | -0.21 | 0.42 | -1.11 | -1.21 | 1.44 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.23 | -0.09 | -0.12 | 1.38 | 1.08 | -1.12 | -1.22 | 2.1 |
| 0.31 | 0.79 | 0.87 | 0.34 | 0.38 | -1.38 | -1.38 | 0.33 | 1.05 | -0.09 | -0.12 | 1.17 | 0.83 | -1.12 | -1.22 | 1.85 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.21 | -0.09 | -0.12 | 1.35 | 0.76 | -0.8 | -0.89 | 1.78 |
| -1.36 | -0.8 | -0.82 | -1.43 | -0.9 | -1.36 | -1.36 | -0.91 | 0.46 | -1.13 | -1.3 | 0.5 | -0.1 | -1.19 | -1.3 | 0.92 |
| -1.87 | -0.98 | -1 | -1.97 | -0.5 | -0.41 | -0.44 | -0.53 | -0.87 | -0.89 | -1.03 | -1.03 | -0.87 | -1.3 | -1.41 | 0.15 |
| -0.48 | -1.19 | -1.23 | -0.5 | -0.56 | -1.17 | -1.18 | -0.58 | 0.3 | -1.08 | -1.25 | 0.31 | -1.36 | -0.22 | -0.27 | -0.34 |
| -1.89 | -0.7 | -0.71 | -1.99 | -2.73 | -1.39 | -1.39 | -2.69 | -1.56 | -0.89 | -1.03 | -1.81 | -0.86 | -0.57 | -0.64 | 0.16 |
| 0.1 | -0.32 | -0.31 | 0.12 | -0.12 | -1.07 | -1.08 | -0.16 | 0.25 | -3.14 | -3.6 | 0.26 | -1.55 | -0.9 | -0.99 | -0.53 |
| 0.37 | 0.34 | 0.39 | 0.4 | 1.19 | 0.85 | 0.79 | 1.12 | 0.43 | 1.11 | 1.26 | 0.46 | 0.94 | 0.77 | 0.77 | 1.96 |
| 1.27 | -3.61 | -3.79 | 1.35 | 0.44 | -1.54 | -1.54 | 0.39 | 0.49 | 1.18 | 1.34 | 0.53 | 0.82 | 0.51 | 0.5 | 1.84 |
| 0.38 | 0.58 | 0.65 | 0.41 | 0.38 | 0.09 | 0.05 | 0.33 | 0.63 | 0.74 | 0.84 | 0.69 | 0.17 | 0.04 | 0 | 1.19 |
| 0.51 | 0.44 | 0.5 | 0.55 | 0.73 | 1.03 | 0.96 | 0.67 | 1.08 | 0.77 | 0.87 | 1.21 | 0.71 | 0.76 | 0.76 | 1.73 |
| -0.84 | 0.76 | 0.84 | -0.88 | 0.74 | 0.95 | 0.88 | 0.68 | 0.49 | 0.51 | 0.57 | 0.53 | 0.23 | 0.21 | 0.18 | 1.25 |
| 0.09 | 0.77 | 0.85 | 0.11 | 1.39 | 0.63 | 0.57 | 1.31 | 0.69 | 0.88 | 1 | 0.76 | 0.77 | 0.39 | 0.37 | 1.79 |
| 0.52 | 0.95 | 1.04 | 0.56 | 1.72 | 0.73 | 0.67 | 1.63 | -0.27 | 1.87 | 2.13 | -0.34 | 0.27 | 0.96 | 0.97 | 1.29 |
| 0.42 | -0.33 | -0.32 | 0.45 | 1.11 | 0.84 | 0.78 | 1.04 | 0.49 | 0.92 | 1.04 | 0.53 | 1.23 | 0.14 | 0.11 | 2.25 |
| 0.31 | 1.03 | 1.12 | 0.34 | 0.75 | 1.37 | 1.29 | 0.69 | 0.81 | 0.98 | 1.11 | 0.9 | 0.94 | 0.6 | 0.59 | 1.96 |
| 0.61 | 0.73 | 0.8 | 0.66 | 0.92 | 0.45 | 0.4 | 0.85 | 0.25 | 0.49 | 0.55 | 0.26 | -1.55 | 0.36 | 0.34 | -0.53 |
| -0.63 | -1.07 | -1.1 | -0.66 | 0.32 | -1.76 | -1.75 | 0.27 | 0.1 | -0.51 | -0.59 | 0.08 | -1.32 | -0.2 | -0.25 | -0.3 |
| -0.74 | -0.32 | -0.31 | -0.77 | -0.44 | -2.55 | -2.52 | -0.47 | 0.54 | 0.18 | 0.2 | 0.59 | -1.51 | 0.07 | 0.03 | -0.49 |
| -0.53 | -1.11 | -1.14 | -0.55 | 0.2 | -1.71 | -1.7 | 0.15 | -0.11 | 0.09 | 0.09 | -0.16 | -0.39 | 0.09 | 0.05 | 0.63 |


| -0.48 | -1.15 | -1.18 | -0.5 | -0.4 | -1.36 | -1.36 | -0.43 | 0.4 | -0.37 | -0.43 | 0.43 | -1.13 | -0.26 | -0.32 | -0.11 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.72 | -0.95 | -0.97 | -0.75 | -0.94 | -1.11 | -1.12 | -0.95 | -0.29 | -0.3 | -0.35 | -0.36 | -1.33 | -0.47 | -0.54 | -0.31 |
| -1.01 | -1.05 | -1.08 | -1.06 | -0.37 | -1.37 | -1.37 | -0.4 | 0.21 | -0.53 | -0.62 | 0.21 | -0.58 | -0.27 | -0.33 | 0.44 |
| -0.17 | -1.06 | -1.09 | -0.17 | -0.62 | -1.31 | -1.31 | -0.64 | 0.21 | 0.56 | 0.63 | 0.21 | -1.54 | -1.32 | -1.44 | -0.52 |
| -0.58 | -0.87 | -0.89 | -0.6 | -1.86 | -0.42 | -0.45 | -1.85 | 0.01 | -2.01 | -2.31 | -0.02 | -1.17 | -0.2 | -0.25 | -0.15 |
| -0.02 | -0.77 | -0.78 | -0.01 | 0.88 | -1.4 | -1.4 | 0.82 | -1.33 | -0.4 | -0.47 | -1.55 | -1.31 | -0.29 | -0.35 | -0.29 |
| 0.2 | -1.28 | -1.32 | 0.22 | -0.58 | -3.11 | -3.06 | -0.6 | 0.28 | 0.32 | 0.36 | 0.29 | 0.35 | 0.27 | 0.24 | 1.37 |
| 0.36 | -0.96 | -0.98 | 0.39 | -0.81 | 0.18 | 0.14 | -0.83 | 0.49 | 0.74 | 0.83 | 0.53 | 0.07 | 0.3 | 0.27 | 1.09 |
| -0.53 | -0.06 | -0.03 | -0.55 | -0.79 | -0.17 | -0.2 | -0.81 | 0.3 | 0.69 | 0.78 | 0.31 | -0.82 | 0.98 | 0.99 | 0.2 |
| -0.27 | 1.5 | 1.62 | -0.28 | 0.24 | 1.02 | 0.95 | 0.19 | -0.26 | 1.17 | 1.33 | -0.33 | 0.14 | 0.78 | 0.78 | 1.16 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.61 | 1.49 | 1.53 | 1.63 |
| 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
| -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 0.98 | 0.93 | 0.94 | 2 |
| -0.87 | 0.34 | 0.39 | -0.91 | -0.18 | 0.59 | 0.53 | -0.21 | -1.03 | 0.91 | 1.03 | -1.21 | 0.32 | 0.74 | 0.74 | 1.34 |
| 0.15 | 0.4 | 0.45 | 0.17 | 0.35 | -0.46 | -0.49 | 0.3 | -0.16 | -0.15 | -0.18 | -0.21 | 0.42 | -1.11 | -1.21 | 1.44 |
| 1.04 | 0.25 | 0.3 | 1.11 | 1.08 | 0.73 | 0.67 | 1.01 | 0.75 | 0.51 | 0.57 | 0.83 | 0.14 | 0.79 | 0.79 | 1.16 |
| -0.93 | 0.67 | 0.74 | -0.97 | 0.11 | -0.5 | -0.52 | 0.07 | 0.2 | -0.23 | -0.28 | 0.2 | 0.9 | -0.3 | -0.36 | 1.92 |
| 0.81 | 1.02 | 1.11 | 0.87 | 0.87 | 0.78 | 0.72 | 0.81 | -0.86 | -0.18 | -0.22 | -1.01 | -0.54 | 1.04 | 1.06 | 0.48 |
| 0.19 | 0.75 | 0.83 | 0.21 | 0.91 | 1.08 | 1.01 | 0.84 | 0.33 | 0.83 | 0.94 | 0.35 | 0.8 | 0.74 | 0.74 | 1.82 |
| -0.05 | 0.74 | 0.81 | -0.04 | 0.92 | 0.21 | 0.17 | 0.85 | -1.61 | 0.74 | 0.84 | -1.87 | 0.56 | 0.66 | 0.65 | 1.58 |


| -0.54 | 0.04 | 0.07 | -0.56 | 0.82 | 0.88 | 0.82 | 0.76 | 0.48 | 0.77 | 0.87 | 0.52 | 0.39 | 0.62 | 0.61 | 1.41 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | -0.29 | -0.27 | 0.01 | 0.72 | -0.05 | -0.09 | 0.66 | -1.17 | 0.55 | 0.62 | -1.37 | 0.81 | 0.72 | 0.72 | 1.83 |
| 0.77 | 0.58 | 0.65 | 0.82 | 0.72 | 0.77 | 0.71 | 0.66 | 0.6 | 0.02 | 0.01 | 0.66 | 1.12 | 0.58 | 0.57 | 2.14 |
| -0.14 | 0.77 | 0.85 | -0.14 | 0.46 | 0.87 | 0.81 | 0.41 | 0.9 | 0.7 | 0.79 | 1 | 1.26 | 0.55 | 0.54 | 2.28 |
| 0.94 | 0.61 | 0.68 | 1 | 0.89 | 1.03 | 0.96 | 0.83 | 0.68 | 0.83 | 0.94 | 0.75 | 0.84 | 0.13 | 0.1 | 1.86 |
| -0.12 | 1.2 | 1.3 | -0.12 | 0.74 | 0.31 | 0.26 | 0.68 | 1.49 | 0.91 | 1.03 | 1.67 | 1.11 | 0.83 | 0.83 | 2.13 |
| 0.61 | 0.73 | 0.8 | 0.66 | 0.92 | 0.45 | 0.4 | 0.85 | 0.25 | 0.49 | 0.55 | 0.26 | -1.55 | 0.36 | 0.34 | -0.53 |
| -0.63 | -1.07 | -1.1 | -0.66 | 0.32 | -1.76 | -1.75 | 0.27 | 0.1 | -0.51 | -0.59 | 0.08 | -1.32 | -0.2 | -0.25 | -0.3 |
| -0.74 | -0.32 | -0.31 | -0.77 | -0.44 | -2.55 | -2.52 | -0.47 | 0.54 | 0.18 | 0.2 | 0.59 | -1.51 | 0.07 | 0.03 | -0.49 |
| -0.53 | -1.11 | -1.14 | -0.55 | 0.2 | -1.71 | -1.7 | 0.15 | -0.11 | 0.09 | 0.09 | -0.16 | -0.39 | 0.09 | 0.05 | 0.63 |
| -0.48 | -1.15 | -1.18 | -0.5 | -0.4 | -1.36 | -1.36 | -0.43 | 0.4 | -0.37 | -0.43 | 0.43 | -1.13 | -0.26 | -0.32 | -0.11 |
| -0.72 | -0.95 | -0.97 | -0.75 | -0.94 | -1.11 | -1.12 | -0.95 | -0.29 | -0.3 | -0.35 | -0.36 | -1.33 | -0.47 | -0.54 | -0.31 |
| -1.01 | -1.05 | -1.08 | -1.06 | -0.37 | -1.37 | -1.37 | -0.4 | 0.21 | -0.53 | -0.62 | 0.21 | -0.58 | -0.27 | -0.33 | 0.44 |
| -0.17 | -1.06 | -1.09 | -0.17 | -0.62 | -1.31 | -1.31 | -0.64 | 0.21 | 0.56 | 0.63 | 0.21 | -1.54 | -1.32 | -1.44 | -0.52 |
| -0.58 | -0.87 | -0.89 | -0.6 | -1.86 | -0.42 | -0.45 | -1.85 | 0.01 | -2.01 | -2.31 | -0.02 | -1.17 | -0.2 | -0.25 | -0.15 |
| -0.02 | -0.77 | -0.78 | -0.01 | 0.88 | -1.4 | -1.4 | 0.82 | -1.33 | -0.4 | -0.47 | -1.55 | -1.31 | -0.29 | -0.35 | -0.29 |
| 0.2 | -1.28 | -1.32 | 0.22 | -0.58 | -3.11 | -3.06 | -0.6 | 0.28 | 0.32 | 0.36 | 0.29 | 0.35 | 0.27 | 0.24 | 1.37 |
| 0.36 | -0.96 | -0.98 | 0.39 | -0.81 | 0.18 | 0.14 | -0.83 | 0.49 | 0.74 | 0.83 | 0.53 | 0.07 | 0.3 | 0.27 | 1.09 |
| -0.53 | -0.06 | -0.03 | -0.55 | -0.79 | -0.33 | -0.36 | -0.81 | 0.3 | 0.69 | 0.78 | 0.31 | -0.82 | 0.98 | 0.99 | 0.2 |
| -0.27 | 1.5 | 1.62 | -0.28 | 0.24 | 1.02 | 0.95 | 0.19 | -0.26 | 1.17 | 1.33 | -0.33 | 0.14 | 0.78 | 0.78 | 1.16 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |


| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.61 | 1.49 | 1.53 | 1.63 |
| 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
| -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 0.98 | 0.93 | 0.94 | 2 |
| -0.87 | 0.34 | 0.39 | -0.91 | -0.18 | 0.59 | 0.53 | -0.21 | -1.03 | 0.91 | 1.03 | -1.21 | 0.32 | 0.74 | 0.74 | 1.34 |
| 0.15 | 0.4 | 0.45 | 0.17 | 0.35 | -0.46 | -0.49 | 0.3 | -0.16 | -0.15 | -0.18 | -0.21 | 0.42 | -1.11 | -1.21 | 1.44 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.23 | -0.09 | -0.12 | 1.38 | 1.08 | -1.12 | -1.22 | 2.1 |
| 0.31 | 0.79 | 0.87 | 0.34 | 0.38 | -1.38 | -1.38 | 0.33 | 1.05 | -0.09 | -0.12 | 1.17 | 0.83 | -1.12 | -1.22 | 1.85 |
| -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.21 | -0.09 | -0.12 | 1.35 | 0.8 | -0.8 | -0.89 | 1.82 |
| -1.36 | -0.8 | -0.82 | -1.43 | -0.9 | -1.36 | -1.36 | -0.91 | 0.46 | -1.13 | -1.3 | 0.5 | -0.1 | -1.19 | -1.3 | 0.92 |
| -1.87 | -0.98 | -1 | -1.97 | -0.5 | -0.41 | -0.44 | -0.53 | -0.87 | -0.89 | -1.03 | -1.03 | -0.87 | -1.3 | -1.41 | 0.15 |
| -0.48 | -1.19 | -1.23 | -0.5 | -0.56 | -1.17 | -1.18 | -0.58 | 0.3 | -1.08 | -1.25 | 0.31 | -1.36 | -0.22 | -0.27 | -0.34 |
| -1.89 | -0.7 | -0.71 | -1.99 | -2.73 | -1.39 | -1.39 | -2.69 | -1.56 | -0.89 | -1.03 | -1.81 | -0.86 | -0.57 | -0.64 | 0.16 |
| -1.82 | -0.59 | -0.59 | -1.92 | -0.71 | -1.42 | -1.42 | -0.73 | -1.84 | -0.73 | -0.85 | -2.13 | -0.21 | -1.38 | -1.5 | 0.81 |
| -0.32 | -0.72 | -0.73 | -0.33 | 0.52 | -0.04 | -0.08 | 0.47 | 0.04 | -0.9 | -1.04 | 0.02 | -0.41 | -0.57 | -0.64 | 0.61 |
| -0.4 | -1.42 | -1.47 | -0.41 | -0.18 | -0.8 | -0.82 | -0.21 | 0.07 | -0.52 | -0.61 | 0.05 | -0.08 | -0.68 | -0.76 | 0.94 |
| -1.01 | -0.7 | -0.71 | -1.06 | 0.91 | -1.88 | -1.87 | 0.84 | -1.13 | -1.01 | -1.17 | -1.32 | -0.3 | -0.84 | -0.93 | 0.72 |
| -0.83 | -0.63 | -0.63 | -0.87 | 0.3 | -1.88 | -1.87 | 0.25 | -0.58 | -0.86 | -1 | -0.69 | -0.47 | -0.84 | -0.93 | 0.55 |
| -1.14 | -0.3 | -0.29 | -1.2 | 0.8 | -1.16 | -1.17 | 0.74 | -0.29 | -0.4 | -0.47 | -0.36 | -0.23 | -0.65 | -0.73 | 0.79 |
| -0.69 | -0.78 | -0.79 | -0.72 | 0.71 | -0.24 | -0.27 | 0.65 | -1 | -1.09 | -1.26 | -1.17 | -0.32 | -0.31 | -0.37 | 0.7 |
| -0.85 | -0.83 | -0.85 | -0.89 | 0.54 | -1.37 | -1.37 | 0.48 | -1.7 | -0.19 | -0.23 | -1.97 | -0.4 | -1.42 | -1.54 | 0.62 |
| -0.31 | -0.66 | -0.67 | -0.32 | -0.26 | -0.61 | -0.63 | -0.29 | -0.23 | -0.59 | -0.69 | -0.29 | -0.26 | -1.19 | -1.3 | 0.76 |
| -0.93 | -0.49 | -0.49 | -0.97 | 0.42 | -0.52 | -0.54 | 0.37 | -0.78 | -1.62 | -1.86 | -0.92 | -0.51 | -1.4 | -1.52 | 0.51 |
| -0.68 | -0.95 | -0.97 | -0.71 | 0.41 | -0.98 | -0.99 | 0.36 | 0.07 | -0.54 | -0.63 | 0.05 | -0.4 | -0.82 | -0.91 | 0.62 |
| -0.49 | -0.95 | -0.97 | -0.51 | 0.49 | -0.58 | -0.6 | 0.44 | -0.9 | -0.55 | -0.64 | -1.06 | -0.06 | -0.96 | -1.06 | 0.96 |


| -0.61 | -0.49 | -0.49 | -0.64 | -0.06 | -0.16 | -0.19 | -0.1 | -0.2 | -1.13 | -1.3 | -0.26 | -0.64 | -1.04 | -1.14 | 0.38 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -1.28 | -0.39 | -0.38 | -1.34 | -0.09 | -1.38 | -1.38 | -0.13 | -1.51 | -0.94 | -1.09 | -1.76 | -0.69 | -0.86 | -0.95 | 0.33 |
| -1.61 | -0.97 | -0.99 | -1.69 | 0.38 | -0.24 | -0.27 | 0.33 | -0.21 | -1.06 | -1.22 | -0.27 | -0.14 | -1.14 | -1.25 | 0.88 |
| -0.75 | -0.3 | -0.29 | -0.78 | 0.24 | -1.16 | -1.17 | 0.19 | -0.38 | -0.4 | -0.47 | -0.46 | -0.55 | -0.65 | -0.73 | 0.47 |
| -1.58 | -0.63 | -0.64 | -1.66 | -0.02 | -1.93 | -1.91 | -0.06 | -1.22 | -0.33 | -0.39 | -1.43 | -0.25 | -1.44 | -1.56 | 0.77 |
| -0.98 | -0.69 | -0.7 | -1.03 | 0.69 | -0.76 | -0.78 | 0.63 | -0.78 | -0.01 | -0.02 | -0.92 | 0 | -0.96 | -1.06 | 1.02 |
| -1.03 | -0.58 | -0.58 | -1.08 | 0.33 | -0.45 | -0.48 | 0.28 | -0.37 | -1.02 | -1.18 | -0.45 | -0.14 | -0.94 | -1.03 | 0.88 |
| -0.33 | -0.09 | -0.06 | -0.34 | 0.8 | -0.29 | -0.32 | 0.74 | 0.19 | -0.34 | -0.4 | 0.19 | -0.39 | -0.75 | -0.83 | 0.63 |
| -0.79 | -0.57 | -0.57 | -0.83 | -0.4 | -0.84 | -0.86 | -0.43 | -0.71 | -0.27 | -0.32 | -0.84 | -0.12 | -1.05 | -1.15 | 0.9 |
| -1.12 | -0.3 | -0.29 | -1.17 | 0.41 | -0.97 | -0.98 | 0.36 | -1.12 | -0.13 | -0.16 | -1.31 | -0.16 | -0.65 | -0.73 | 0.86 |
| -0.72 | -0.86 | -0.88 | -0.75 | -0.14 | -1.18 | -1.19 | -0.18 | 0.01 | -1.08 | -1.25 | -0.02 | -0.25 | -1.07 | -1.17 | 0.77 |
| -0.66 | -0.95 | -0.97 | -0.69 | -0.14 | -1.47 | -1.47 | -0.18 | -1.37 | -0.04 | -0.06 | -1.6 | -0.26 | -1.19 | -1.3 | 0.76 |
| -0.52 | -1.35 | -1.4 | -0.54 | 0.32 | -0.95 | -0.96 | 0.27 | -0.96 | -1.93 | -2.22 | -1.13 | -0.55 | -0.87 | -0.96 | 0.47 |
| -0.49 | -0.3 | -0.29 | -0.51 | -0.6 | -1.16 | -1.17 | -0.62 | -0.68 | -0.4 | -0.47 | -0.81 | -0.55 | -0.65 | -0.73 | 0.47 |
| -0.32 | -0.29 | -0.28 | -0.33 | 0.58 | -0.36 | -0.39 | 0.52 | -0.66 | -0.37 | -0.43 | -0.79 | -0.73 | -1.16 | -1.27 | 0.29 |
| -0.25 | -0.63 | -0.63 | -0.25 | 0.24 | -2.65 | -2.61 | 0.19 | 0.01 | -0.49 | -0.57 | -0.02 | -1.3 | -1.18 | -1.29 | -0.28 |
| -0.59 | -0.59 | -0.59 | -0.61 | 0.13 | -0.44 | -0.47 | 0.09 | -0.41 | -0.58 | -0.68 | -0.5 | -0.63 | -0.91 | -1 | 0.39 |
| -0.99 | -0.35 | -0.34 | -1.04 | 0.2 | -0.29 | -0.32 | 0.15 | -1.32 | -2.01 | -2.31 | -1.54 | -0.2 | -0.78 | -0.87 | 0.82 |
| -0.53 | -0.84 | -0.86 | -0.55 | 0.22 | -1.49 | -1.49 | 0.17 | -0.27 | -0.03 | -0.05 | -0.34 | -0.29 | -0.67 | -0.75 | 0.73 |
| -0.43 | -0.69 | -0.7 | -0.44 | 0.46 | -0.37 | -0.4 | 0.41 | -0.35 | -0.16 | -0.19 | -0.43 | -0.48 | -0.86 | -0.95 | 0.54 |
| -0.98 | -0.33 | -0.32 | -1.03 | 0.23 | -0.29 | -0.32 | 0.18 | -1.09 | -1.01 | -1.17 | -1.28 | -0.29 | -0.68 | -0.76 | 0.73 |
| -0.36 | -0.56 | -0.56 | -0.37 | 0.59 | -0.79 | -0.81 | 0.53 | -0.14 | -0.97 | -1.12 | -0.19 | -0.24 | -1.16 | -1.27 | 0.78 |
| -0.55 | -0.57 | -0.57 | -0.57 | -0.13 | -1.22 | -1.22 | -0.17 | -0.8 | -0.47 | -0.55 | -0.95 | -0.16 | -1.12 | -1.23 | 0.86 |
| -0.96 | -0.3 | -0.29 | -1.01 | 0.89 | -1.16 | -1.17 | 0.83 | -0.5 | -0.4 | -0.47 | -0.6 | -0.54 | -0.65 | -0.73 | 0.48 |
| -1.26 | -0.34 | -0.33 | -1.32 | 0.58 | -0.66 | -0.68 | 0.52 | -0.93 | -2.01 | -2.31 | -1.09 | -0.22 | 0.15 | 0.12 | 0.8 |


| -1.33 | -0.49 | -0.49 | -1.4 | -0.49 | -0.76 | -0.78 | -0.52 | -1.56 | -1 | -1.15 | -1.81 | -0.22 | -2.01 | -2.17 | 0.8 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.84 | -0.91 | -0.93 | -0.88 | 0.18 | -1.46 | -1.46 | 0.13 | -0.61 | -1.11 | -1.28 | -0.73 | -0.48 | 0.13 | 0.09 | 0.54 |
| -0.5 | -0.8 | -0.81 | -0.52 | -0.35 | -1.73 | -1.72 | -0.38 | 0.09 | -1.14 | -1.32 | 0.07 | -0.31 | -0.94 | -1.03 | 0.71 |
| -0.2 | -0.98 | -1 | -0.2 | 0.01 | -1.5 | -1.5 | -0.03 | -0.06 | -2.58 | -2.96 | -0.1 | -0.65 | -1.53 | -1.66 | 0.37 |
| -0.47 | -0.74 | -0.75 | -0.49 | 0.53 | -1.18 | -1.19 | 0.48 | -0.36 | -1.48 | -1.71 | -0.44 | -0.08 | -1.07 | -1.17 | 0.94 |
| -0.92 | -1.19 | -1.23 | -0.96 | 0.01 | -1.21 | -1.21 | -0.03 | -0.76 | -0.57 | -0.66 | -0.9 | -0.03 | -0.75 | -0.83 | 0.99 |
| -0.4 | -1.35 | -1.4 | -0.41 | -0.48 | -0.95 | -0.96 | -0.51 | 0.04 | -1.93 | -2.22 | 0.02 | -0.34 | -0.87 | -0.96 | 0.68 |
| -1.08 | -0.3 | -0.29 | -1.13 | -0.9 | -0.97 | -0.98 | -0.91 | -0.28 | -0.13 | -0.16 | -0.35 | 0.05 | -0.65 | -0.73 | 1.07 |
| -0.55 | -0.3 | -0.29 | -0.57 | -0.3 | -1.28 | -1.28 | -0.33 | 0.5 | -1.13 | -1.31 | 0.54 | -0.72 | -0.25 | -0.31 | 0.3 |
| -0.56 | -0.57 | -0.57 | -0.58 | -0.25 | -1.22 | -1.22 | -0.28 | -0.12 | -0.47 | -0.55 | -0.17 | -0.2 | -1.12 | -1.23 | 0.82 |
| 0.17 | -0.39 | -0.38 | 0.19 | 0.27 | -1.22 | -1.22 | 0.22 | 0.12 | -0.55 | -0.64 | 0.11 | -0.01 | -1.12 | -1.23 | 1.01 |
| -0.08 | -0.24 | -0.22 | -0.07 | 0.23 | -1.16 | -1.17 | 0.18 | 0.07 | -0.5 | -0.58 | 0.05 | -0.26 | -0.27 | -0.33 | 0.76 |
| -0.78 | -0.28 | -0.26 | -0.82 | 0.12 | -1.38 | -1.38 | 0.08 | -1.53 | -0.52 | -0.61 | -1.78 | -0.27 | -0.64 | -0.72 | 0.75 |
| 0.09 | -0.41 | -0.4 | 0.11 | -0.47 | -1.16 | -1.17 | -0.5 | 0.19 | -0.36 | -0.42 | 0.19 | -0.5 | -1.17 | -1.28 | 0.52 |
| -0.78 | -0.62 | -0.62 | -0.82 | 0.95 | -1.93 | -1.91 | 0.88 | -1.07 | -2.26 | -2.6 | -1.25 | -1.05 | -0.52 | -0.59 | -0.03 |
| -0.74 | -1.48 | -1.53 | -0.77 | -0.39 | -1.21 | -1.21 | -0.42 | -0.18 | -0.4 | -0.47 | -0.24 | -0.6 | -0.75 | -0.83 | 0.42 |
| -0.69 | -0.58 | -0.58 | -0.72 | 0.77 | -1.18 | -1.19 | 0.71 | -0.55 | -1.33 | -1.53 | -0.66 | -0.7 | -1.15 | -1.26 | 0.32 |
| -0.72 | -0.74 | -0.75 | -0.75 | 0.9 | -1.25 | -1.25 | 0.83 | -0.84 | -1.48 | -1.71 | -0.99 | 0.02 | -1.07 | -1.17 | 1.04 |
| -1.17 | -0.3 | -0.29 | -1.23 | -0.72 | -1.16 | -1.17 | -0.74 | -0.21 | -0.4 | -0.47 | -0.27 | -0.22 | -0.65 | -0.73 | 0.8 |
| -0.99 | -0.63 | -0.63 | -1.04 | -1.05 | -1.18 | -1.19 | -1.06 | -0.66 | -1.28 | -1.48 | -0.79 | -0.42 | -1.07 | -1.17 | 0.6 |
| -1.17 | -0.75 | -0.76 | -1.23 | -0.21 | -1.45 | -1.45 | -0.24 | 0.01 | 0.02 | 0.01 | -0.02 | -0.34 | -0.33 | -0.39 | 0.68 |
| -0.36 | -0.34 | -0.33 | -0.37 | 0.62 | -0.66 | -0.68 | 0.56 | -1.04 | -0.43 | -0.5 | -1.22 | -0.12 | -0.13 | -0.18 | 0.9 |
| -1.17 | -0.97 | -0.99 | -1.23 | 0.45 | -0.3 | -0.33 | 0.4 | -2.02 | -0.25 | -0.3 | -2.34 | -0.3 | -3.09 | -3.3 | 0.72 |
| -0.14 | -0.55 | -0.55 | -0.14 | 0.3 | -0.72 | -0.74 | 0.25 | -0.07 | 0.35 | 0.39 | -0.11 | -1.3 | -1.15 | -1.26 | -0.28 |
| -0.96 | -0.42 | -0.41 | -1.01 | 0.07 | -3.14 | -3.09 | 0.03 | -1.87 | -0.25 | -0.3 | -2.17 | -0.3 | -0.6 | -0.68 | 0.72 |


| -0.75 | -0.83 | -0.85 | -0.78 | 0.95 | 0.05 | 0.01 | 0.88 | -0.89 | -0.31 | -0.37 | -1.05 | -0.68 | -0.57 | -0.64 | 0.34 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.42 | -0.79 | -0.8 | -0.43 | -0.31 | -0.73 | -0.75 | -0.34 | 0.47 | -0.43 | -0.5 | 0.51 | 0.1 | -0.9 | -0.99 | 1.12 |
| -1.35 | -1.26 | -1.3 | -1.42 | 0.05 | -0.97 | -0.98 | 0.01 | -0.91 | -0.07 | -0.09 | -1.07 | -0.29 | -0.67 | -0.75 | 0.73 |
| -1.7 | -0.3 | -0.29 | -1.79 | -0.39 | -1.16 | -1.17 | -0.42 | -0.94 | -0.18 | -0.22 | -1.11 | -0.66 | -0.58 | -0.65 | 0.36 |
| -0.4 | -0.74 | -0.75 | -0.41 | -0.53 | -1.18 | -1.19 | -0.56 | 0.17 | -0.02 | -0.04 | 0.16 | -1.04 | -1.19 | -1.3 | -0.02 |
| -0.75 | -0.89 | -0.91 | -0.78 | 0.67 | -0.33 | -0.36 | 0.61 | -1.6 | -0.79 | -0.92 | -1.86 | -0.23 | -1.18 | -1.29 | 0.79 |
| -2.09 | -0.7 | -0.71 | -2.2 | 0.61 | 0.92 | 0.86 | 0.55 | -2.18 | -1.03 | -1.19 | -2.52 | -0.13 | -0.57 | -0.64 | 0.89 |
| -0.86 | -0.77 | -0.78 | -0.9 | 0.33 | 0.6 | 0.54 | 0.28 | -0.94 | -1.17 | -1.35 | -1.11 | -0.23 | -0.95 | -1.05 | 0.79 |
| -1.63 | -0.6 | -0.6 | -1.71 | 0.04 | 0.7 | 0.64 | 0 | -1.39 | -1.14 | -1.32 | -1.62 | -0.15 | -0.68 | -0.76 | 0.87 |
| -0.73 | -0.46 | -0.45 | -0.76 | 0.54 | -0.24 | -0.27 | 0.48 | -0.01 | -1.1 | -1.27 | -0.04 | -1.3 | -0.82 | -0.91 | -0.28 |
| -0.64 | -0.63 | -0.64 | -0.67 | 0.74 | 0.55 | 0.5 | 0.68 | -0.73 | -0.6 | -0.7 | -0.87 | -0.25 | -1.21 | -1.32 | 0.77 |
| -0.43 | -0.77 | -0.78 | -0.44 | 0.53 | 0.06 | 0.02 | 0.48 | -1.35 | -0.34 | -0.4 | -1.57 | -1.06 | -1.17 | -1.28 | -0.04 |
| -0.98 | -0.66 | -0.67 | -1.03 | 0.11 | 0.15 | 0.11 | 0.07 | -0.83 | -1 | -1.15 | -0.98 | -0.52 | -1.21 | -1.32 | 0.5 |
| -1.03 | -0.93 | -0.95 | -1.08 | 0.73 | 0.14 | 0.1 | 0.67 | -0.44 | -0.62 | -0.72 | -0.53 | 0.07 | -0.6 | -0.68 | 1.09 |
| -0.78 | -0.84 | -0.86 | -0.82 | -0.07 | 0.3 | 0.25 | -0.11 | -1.25 | -0.58 | -0.67 | -1.46 | -1.3 | -0.43 | -0.5 | -0.28 |
| -1.24 | -0.59 | -0.59 | -1.3 | 0.51 | 0.46 | 0.41 | 0.46 | -1.81 | -0.66 | -0.77 | -2.1 | 0.03 | -0.73 | -0.81 | 1.05 |
| -0.4 | -0.97 | -0.99 | -0.41 | 0.92 | 0.19 | 0.15 | 0.85 | -0.55 | -1.37 | -1.58 | -0.66 | -0.18 | -0.44 | -0.51 | 0.84 |
| -0.42 | -0.46 | -0.45 | -0.43 | 0.62 | 0.63 | 0.57 | 0.56 | -0.68 | -0.72 | -0.84 | -0.81 | -0.02 | -0.59 | -0.66 | 1 |
| -0.58 | -0.94 | -0.96 | -0.6 | 0.64 | 0.48 | 0.43 | 0.58 | -0.91 | -0.76 | -0.88 | -1.07 | -1 | -1.12 | -1.23 | 0.02 |
| -0.43 | -0.6 | -0.6 | -0.44 | 0.75 | -0.27 | -0.3 | 0.69 | -0.8 | -0.49 | -0.57 | -0.95 | -1.28 | -0.98 | -1.08 | -0.26 |
| -0.2 | -0.52 | -0.52 | -0.2 | 1 | 0.2 | 0.16 | 0.93 | 0.07 | -1.76 | -2.02 | 0.05 | -0.32 | -2.52 | -2.7 | 0.7 |
| -1.06 | -0.35 | -0.34 | -1.11 | -0.03 | 0.47 | 0.42 | -0.07 | -0.37 | -1.04 | -1.2 | -0.45 | -0.01 | -0.66 | -0.74 | 1.01 |
| -1.05 | -0.75 | -0.76 | -1.1 | 0.44 | -0.12 | -0.16 | 0.39 | -0.83 | -0.47 | -0.55 | -0.98 | -0.2 | -0.65 | -0.73 | 0.82 |
| -0.26 | -0.5 | -0.5 | -0.27 | 0.8 | 0.61 | 0.55 | 0.74 | 0.07 | -0.42 | -0.49 | 0.05 | -0.64 | -1.15 | -1.26 | 0.38 |
| -0.4 | -1.19 | -1.23 | -0.41 | 0.94 | 0.09 | 0.05 | 0.87 | -0.75 | -0.92 | -1.06 | -0.89 | -0.17 | -1.47 | -1.59 | 0.85 |


| -0.52 | -0.51 | -0.51 | -0.54 | 0.89 | 0.29 | 0.24 | 0.83 | -0.33 | -0.46 | -0.54 | -0.41 | -0.01 | -0.55 | -0.62 | 1.01 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.96 | -0.46 | -0.46 | -1.01 | 0.74 | 0.61 | 0.55 | 0.68 | -1.71 | -1.4 | -1.61 | -1.99 | -0.37 | -0.35 | -0.41 | 0.65 |
| -0.49 | -1.3 | -1.34 | -0.51 | 0.45 | 0.39 | 0.34 | 0.4 | -1 | -0.48 | -0.56 | -1.17 | -0.49 | -0.45 | -0.52 | 0.53 |
| -0.61 | -0.9 | -0.92 | -0.64 | 0.97 | 0.23 | 0.18 | 0.9 | -0.66 | -0.05 | -0.07 | -0.79 | -0.39 | -0.87 | -0.96 | 0.63 |
| -1.56 | -1.09 | -1.12 | -1.64 | 0.46 | 0.64 | 0.58 | 0.41 | -1.75 | -1.16 | -1.34 | -2.03 | -0.51 | -0.53 | -0.6 | 0.51 |
| -0.57 | -0.85 | -0.87 | -0.59 | -0.21 | -0.36 | -0.39 | -0.24 | -0.69 | -1.1 | -1.27 | -0.82 | -0.3 | -0.33 | -0.39 | 0.72 |
| -0.79 | -0.6 | -0.6 | -0.83 | 0.65 | -0.38 | -0.41 | 0.59 | -1 | -0.77 | -0.89 | -1.17 | -0.25 | -1.17 | -1.28 | 0.77 |
| -1.09 | -0.69 | -0.7 | -1.14 | 0.06 | 0.21 | 0.17 | 0.02 | -1.31 | -0.77 | -0.89 | -1.53 | -0.02 | -1.16 | -1.27 | 1 |
| -0.79 | -0.64 | -0.65 | -0.83 | 0.75 | 0.73 | 0.67 | 0.69 | -0.32 | -1.08 | -1.25 | -0.4 | -1.5 | -0.54 | -0.61 | -0.48 |
| -1.68 | -0.95 | -0.97 | -1.77 | 0.41 | 0.42 | 0.37 | 0.36 | -0.74 | -0.12 | -0.15 | -0.88 | -0.31 | -0.01 | -0.05 | 0.71 |
| -0.52 | -0.75 | -0.76 | -0.54 | 0.13 | 0.48 | 0.43 | 0.09 | -0.15 | -0.4 | -0.47 | -0.2 | -0.57 | -1.08 | -1.18 | 0.45 |
| -0.49 | -0.8 | -0.82 | -0.51 | 0.03 | 0.52 | 0.47 | -0.01 | 0 | -1.1 | -1.27 | -0.03 | -0.07 | -0.79 | -0.88 | 0.95 |
| -0.23 | -0.62 | -0.62 | -0.23 | -0.3 | 0.8 | 0.74 | -0.33 | -0.52 | -0.94 | -1.09 | -0.62 | -0.05 | -0.49 | -0.56 | 0.97 |
| -1.38 | -0.29 | -0.27 | -1.45 | 0.31 | 0.02 | -0.02 | 0.26 | -1.73 | -1.34 | -1.54 | -2.01 | -0.01 | -0.88 | -0.97 | 1.01 |
| -0.58 | -0.5 | -0.5 | -0.6 | 0.28 | 0.69 | 0.63 | 0.23 | -0.28 | -0.93 | -1.08 | -0.35 | -0.66 | -0.03 | -0.07 | 0.36 |
| -0.59 | -0.7 | -0.71 | -0.61 | 0.18 | 0.59 | 0.53 | 0.13 | 0.16 | -1.27 | -1.46 | 0.15 | -0.86 | -0.66 | -0.74 | 0.16 |
| -1.23 | -0.8 | -0.82 | -1.29 | 0.15 | 0.77 | 0.71 | 0.11 | -0.73 | -0.84 | -0.97 | -0.87 | -0.14 | -0.16 | -0.21 | 0.88 |
| -1.28 | -1.12 | -1.15 | -1.34 | 0.39 | 1.18 | 1.11 | 0.34 | -2.18 | -1.07 | -1.23 | -2.52 | -0.35 | -0.48 | -0.55 | 0.67 |
| -0.41 | -0.67 | -0.68 | -0.42 | -0.2 | 0.21 | 0.17 | -0.23 | 0.01 | -1.11 | -1.28 | -0.02 | -0.42 | -0.57 | -0.64 | 0.6 |
| -0.14 | -0.06 | -0.03 | -0.14 | -0.63 | 0.52 | 0.47 | -0.65 | 0.73 | -0.52 | -0.61 | 0.81 | 0.81 | -1.02 | -1.12 | 1.83 |
| 0.07 | 0.16 | 0.2 | 0.08 | -0.07 | 0.27 | 0.22 | -0.11 | -0.51 | -2.18 | -2.5 | -0.61 | -0.69 | -1.12 | -1.22 | 0.33 |
| -0.72 | 0.08 | 0.12 | -0.75 | 0.31 | 1.16 | 1.09 | 0.26 | -0.37 | 0.52 | 0.58 | -0.45 | 1.12 | 0.47 | 0.45 | 2.14 |
| 0.06 | 0.35 | 0.4 | 0.07 | 0.76 | -1.19 | -1.2 | 0.7 | -0.7 | 0.69 | 0.78 | -0.83 | 1.28 | 0.9 | 0.91 | 2.3 |
| 0.58 | 0.31 | 0.36 | 0.62 | -0.1 | 0.78 | 0.72 | -0.14 | 0.8 | 0.69 | 0.78 | 0.89 | 0.32 | 0.52 | 0.51 | 1.34 |
| 1.11 | -0.44 | -0.43 | 1.18 | 0.73 | 1.15 | 1.08 | 0.67 | 0.72 | 0.68 | 0.77 | 0.79 | 0.57 | 0.48 | 0.46 | 1.59 |


| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.61 | 0.73 | 0.8 | 0.66 | 0.92 | 0.45 | 0.4 | 0.85 | 0.25 | 0.49 | 0.55 | 0.26 | -1.55 | 0.36 | 0.34 | -0.53 |
| -0.63 | -1.07 | -1.1 | -0.66 | 0.32 | -1.76 | -1.75 | 0.27 | 0.1 | -0.51 | -0.59 | 0.08 | -1.32 | -0.2 | -0.25 | -0.3 |
| -0.74 | -0.32 | -0.31 | -0.77 | -0.44 | -2.55 | -2.52 | -0.47 | 0.54 | 0.18 | 0.2 | 0.59 | -1.51 | 0.07 | 0.03 | -0.49 |
| -0.53 | -1.11 | -1.14 | -0.55 | 0.2 | -1.71 | -1.7 | 0.15 | -0.11 | 0.09 | 0.09 | -0.16 | -0.39 | 0.09 | 0.05 | 0.63 |
| -0.48 | -1.15 | -1.18 | -0.5 | -0.4 | -1.36 | -1.36 | -0.43 | 0.4 | -0.37 | -0.43 | 0.43 | -1.13 | -0.26 | -0.32 | -0.11 |
| -0.72 | -0.95 | -0.97 | -0.75 | -0.94 | -1.11 | -1.12 | -0.95 | -0.29 | -0.3 | -0.35 | -0.36 | -1.33 | -0.47 | -0.54 | -0.31 |
| -1.01 | -1.05 | -1.08 | -1.06 | -0.37 | -1.37 | -1.37 | -0.4 | 0.21 | -0.53 | -0.62 | 0.21 | -0.58 | -0.27 | -0.33 | 0.44 |
| -0.17 | -1.06 | -1.09 | -0.17 | -0.62 | -1.31 | -1.31 | -0.64 | 0.21 | 0.56 | 0.63 | 0.21 | -1.54 | -1.32 | -1.44 | -0.52 |
| -0.58 | -0.87 | -0.89 | -0.6 | -1.86 | -0.42 | -0.45 | -1.85 | 0.01 | -2.01 | -2.31 | -0.02 | -1.17 | -0.2 | -0.25 | -0.15 |
| -0.02 | -0.77 | -0.78 | -0.01 | 0.88 | -1.4 | -1.4 | 0.82 | -1.33 | -0.4 | -0.47 | -1.55 | -1.31 | -0.29 | -0.35 | -0.29 |
| 0.2 | -1.28 | -1.32 | 0.22 | -0.58 | -3.11 | -3.06 | -0.6 | 0.28 | 0.32 | 0.36 | 0.29 | 0.35 | 0.27 | 0.24 | 1.37 |
| 0.36 | -0.96 | -0.98 | 0.39 | -0.81 | 0.18 | 0.14 | -0.83 | 0.49 | 0.74 | 0.83 | 0.53 | 0.07 | 0.3 | 0.27 | 1.09 |
| -0.53 | -0.06 | -0.03 | -0.55 | -0.79 | -0.33 | -0.36 | -0.81 | 0.3 | 0.69 | 0.78 | 0.31 | -0.82 | 0.98 | 0.99 | 0.2 |
| -0.27 | 1.5 | 1.62 | -0.28 | 0.24 | 1.02 | 0.95 | 0.19 | -0.26 | 1.17 | 1.33 | -0.33 | 0.14 | 0.78 | 0.78 | 1.16 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| -0.14 | -0.06 | -0.03 | -0.14 | -0.63 | 0.52 | 0.47 | -0.65 | 0.73 | -0.52 | -0.61 | 0.81 | 0.81 | -1.02 | -1.12 | 1.83 |
| 0.07 | 0.16 | 0.2 | 0.08 | -0.07 | 0.27 | 0.22 | -0.11 | -0.51 | -2.18 | -2.5 | -0.61 | -0.69 | -1.12 | -1.22 | 0.33 |
| -0.72 | 0.08 | 0.12 | -0.75 | 0.31 | 1.16 | 1.09 | 0.26 | -0.37 | 0.52 | 0.58 | -0.45 | 1.12 | 0.47 | 0.45 | 2.14 |


| 0.06 | 0.35 | 0.4 | 0.07 | 0.76 | -1.19 | -1.2 | 0.7 | -0.7 | 0.69 | 0.78 | -0.83 | 1.28 | 0.9 | 0.91 | 2.3 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.58 | 0.31 | 0.36 | 0.62 | -0.1 | 0.78 | 0.72 | -0.14 | 0.8 | 0.69 | 0.78 | 0.89 | 0.32 | 0.52 | 0.51 | 1.34 |
| 1.11 | -0.44 | -0.43 | 1.18 | 0.73 | 1.15 | 1.08 | 0.67 | 0.72 | 0.68 | 0.77 | 0.79 | 0.57 | 0.48 | 0.46 | 1.59 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.61 | 0.73 | 0.8 | 0.66 | 0.92 | 0.45 | 0.4 | 0.85 | 0.25 | 0.49 | 0.55 | 0.26 | -1.55 | 0.36 | 0.34 | -0.53 |
| -0.63 | -1.07 | -1.1 | -0.66 | 0.32 | -1.76 | -1.75 | 0.27 | 0.1 | -0.51 | -0.59 | 0.08 | -1.32 | -0.2 | -0.25 | -0.3 |
| -0.74 | -0.32 | -0.31 | -0.77 | -0.44 | -2.55 | -2.52 | -0.47 | 0.54 | 0.18 | 0.2 | 0.59 | -1.51 | 0.07 | 0.03 | -0.49 |
| -0.53 | -1.11 | -1.14 | -0.55 | 0.2 | -1.71 | -1.7 | 0.15 | -0.11 | 0.09 | 0.09 | -0.16 | -0.39 | 0.09 | 0.05 | 0.63 |
| -0.48 | -1.15 | -1.18 | -0.5 | -0.4 | -1.36 | -1.36 | -0.43 | 0.4 | -0.37 | -0.43 | 0.43 | -1.13 | -0.26 | -0.32 | -0.11 |
| -0.72 | -0.95 | -0.97 | -0.75 | -0.94 | -1.11 | -1.12 | -0.95 | -0.29 | -0.3 | -0.35 | -0.36 | -1.33 | -0.47 | -0.54 | -0.31 |
| -1.01 | -1.05 | -1.08 | -1.06 | -0.37 | -1.37 | -1.37 | -0.4 | 0.21 | -0.53 | -0.62 | 0.21 | -0.58 | -0.27 | -0.33 | 0.44 |
| -0.17 | -1.06 | -1.09 | -0.17 | -0.62 | -1.31 | -1.31 | -0.64 | 0.21 | 0.56 | 0.63 | 0.21 | -1.54 | -1.32 | -1.44 | -0.52 |
| -0.58 | -0.87 | -0.89 | -0.6 | -1.86 | -0.42 | -0.45 | -1.85 | 0.01 | -2.01 | -2.31 | -0.02 | -1.17 | -0.2 | -0.25 | -0.15 |
| -0.02 | -0.77 | -0.78 | -0.01 | 0.88 | -1.4 | -1.4 | 0.82 | -1.33 | -0.4 | -0.47 | -1.55 | -1.31 | -0.29 | -0.35 | -0.29 |
| 0.2 | -1.28 | -1.32 | 0.22 | -0.58 | -3.11 | -3.06 | -0.6 | 0.28 | 0.32 | 0.36 | 0.29 | 0.35 | 0.27 | 0.24 | 1.37 |
| 0.36 | -0.96 | -0.98 | 0.39 | -0.81 | 0.18 | 0.14 | -0.83 | 0.49 | 0.74 | 0.83 | 0.53 | 0.07 | 0.3 | 0.27 | 1.09 |
| -0.53 | -0.06 | -0.03 | -0.55 | -0.79 | -0.33 | -0.36 | -0.81 | 0.3 | 0.69 | 0.78 | 0.31 | -0.82 | 0.98 | 0.99 | 0.2 |
| -0.27 | 1.5 | 1.62 | -0.28 | 0.24 | 1.02 | 0.95 | 0.19 | -0.26 | 1.17 | 1.33 | -0.33 | 0.14 | 0.78 | 0.78 | 1.16 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |


| -0.14 | -0.06 | -0.03 | -0.14 | -0.63 | 0.52 | 0.47 | -0.65 | 0.73 | -0.52 | -0.61 | 0.81 | 0.81 | -1.02 | -1.12 | 1.83 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.07 | 0.16 | 0.2 | 0.08 | -0.07 | 0.27 | 0.22 | -0.11 | -0.51 | -2.18 | -2.5 | -0.61 | -0.69 | -1.12 | -1.22 | 0.33 |
| -0.72 | 0.08 | 0.12 | -0.75 | 0.31 | 1.16 | 1.09 | 0.26 | -0.37 | 0.52 | 0.58 | -0.45 | 1.12 | 0.47 | 0.45 | 2.14 |
| 0.06 | 0.35 | 0.4 | 0.07 | 0.76 | -1.19 | -1.2 | 0.7 | -0.7 | 0.69 | 0.78 | -0.83 | 1.28 | 0.9 | 0.91 | 2.3 |
| 0.58 | 0.31 | 0.36 | 0.62 | -0.1 | 0.78 | 0.72 | -0.14 | 0.8 | 0.69 | 0.78 | 0.89 | 0.32 | 0.52 | 0.51 | 1.34 |
| 1.11 | -0.44 | -0.43 | 1.18 | 0.73 | 1.15 | 1.08 | 0.67 | 0.72 | 0.68 | 0.77 | 0.79 | 0.57 | 0.48 | 0.46 | 1.59 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.61 | 1.49 | 1.53 | 1.63 |
| 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
| -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 0.98 | 0.93 | 0.94 | 2 |
| 0.61 | 0.73 | 0.8 | 0.66 | 0.92 | 0.45 | 0.4 | 0.85 | 0.25 | 0.49 | 0.55 | 0.26 | -1.55 | 0.36 | 0.34 | -0.53 |
| -0.63 | -1.07 | -1.1 | -0.66 | 0.32 | -1.76 | -1.75 | 0.27 | 0.1 | -0.51 | -0.59 | 0.08 | -1.32 | -0.2 | -0.25 | -0.3 |
| -0.74 | -0.32 | -0.31 | -0.77 | -0.44 | -2.55 | -2.52 | -0.47 | 0.54 | 0.18 | 0.2 | 0.59 | 0 | 0.07 | 0.03 | 1.02 |
| -0.53 | -1.11 | -1.14 | -0.55 | 0.2 | -1.71 | -1.7 | 0.15 | -0.11 | 0.09 | 0.09 | -0.16 | -0.39 | 0.09 | 0.05 | 0.63 |
| -0.48 | -1.15 | -1.18 | -0.5 | -0.4 | -1.36 | -1.36 | -0.43 | 0.4 | -0.37 | -0.43 | 0.43 | -1.13 | -0.26 | -0.32 | -0.11 |
| -0.72 | -0.95 | -0.97 | -0.75 | -0.94 | -1.11 | -1.12 | -0.95 | -0.29 | -0.3 | -0.35 | -0.36 | -1.33 | -0.47 | -0.54 | -0.31 |
| -1.01 | -1.05 | -1.08 | -1.06 | -0.37 | -1.37 | -1.37 | -0.4 | 0.21 | -0.53 | -0.62 | 0.21 | -0.58 | -0.27 | -0.33 | 0.44 |
| -0.17 | -1.06 | -1.09 | -0.17 | -0.62 | -1.31 | -1.31 | -0.64 | 0.21 | 0.56 | 0.63 | 0.21 | -1.54 | -1.32 | -1.44 | -0.52 |
| -0.58 | -0.87 | -0.89 | -0.6 | -1.86 | -0.42 | -0.45 | -1.85 | 0.01 | -2.01 | -2.31 | -0.02 | -1.17 | -0.2 | -0.25 | -0.15 |
| -0.02 | -0.77 | -0.78 | -0.01 | 0.88 | -1.4 | -1.4 | 0.82 | -1.33 | -0.4 | -0.47 | -1.55 | -1.31 | -0.29 | -0.35 | -0.29 |
| 0.2 | -1.28 | -1.32 | 0.22 | -0.58 | -3.11 | -3.06 | -0.6 | 0.28 | 0.32 | 0.36 | 0.29 | 0.35 | 0.27 | 0.24 | 1.37 |


| 0.36 | -0.96 | -0.98 | 0.39 | -0.81 | 0.18 | 0.14 | -0.83 | 0.49 | 0.74 | 0.83 | 0.53 | 0.07 | 0.3 | 0.27 | 1.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -0.53 | -0.06 | -0.03 | -0.55 | -0.79 | -0.33 | -0.36 | -0.81 | 0.3 | 0.69 | 0.78 | 0.31 | -0.82 | 0.98 | 0.99 | 0.2 |
| -0.27 | 1.5 | 1.62 | -0.28 | 0.24 | 1.02 | 0.95 | 0.19 | -0.26 | 1.17 | 1.33 | -0.33 | 0.14 | 0.78 | 0.78 | 1.16 |
| -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
| 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
| -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
| 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
| -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
| -2.48 | -0.65 | -0.66 | -2.61 | 0.25 | 0.94 | 0.87 | 0.2 | -2.08 | -1.08 | -1.25 | -2.41 | -0.42 | -1.14 | -1.25 | 0.6 |
| -1.47 | -0.99 | -1.02 | -1.55 | 0.23 | 1.12 | 1.05 | 0.18 | -1.23 | -0.84 | -0.97 | -1.44 | -0.39 | -0.01 | -0.05 | 0.63 |
| -0.09 | -0.64 | -0.65 | -0.09 | 0.18 | 0.42 | 0.37 | 0.13 | 0.33 | -0.3 | -0.36 | 0.35 | -0.09 | -1.14 | -1.25 | 0.93 |
| -0.71 | -0.8 | -0.81 | -0.74 | 0.56 | 1.03 | 0.96 | 0.5 | -0.47 | -1.07 | -1.24 | -0.57 | -0.24 | -1.19 | -1.3 | 0.78 |
| -0.53 | -0.7 | -0.71 | -0.55 | -0.68 | 0.1 | 0.06 | -0.7 | 0.01 | -0.79 | -0.92 | -0.02 | -0.17 | -1.29 | -1.4 | 0.85 |
| -0.1 | -0.57 | -0.57 | -0.1 | -0.04 | -0.07 | -0.11 | -0.08 | 0.3 | -1.15 | -1.33 | 0.31 | -0.41 | -0.71 | -0.79 | 0.61 |
| -0.21 | -0.68 | -0.69 | -0.21 | 0.25 | 0.03 | -0.01 | 0.2 | 0.31 | -0.6 | -0.7 | 0.32 | 0.03 | -1.93 | -2.08 | 1.05 |
| -0.62 | -0.69 | -0.7 | -0.65 | 0.46 | 0.45 | 0.4 | 0.41 | 0.1 | -0.65 | -0.75 | 0.08 | -0.15 | -0.73 | -0.81 | 0.87 |
| -1.37 | -0.87 | -0.89 | -1.44 | 0.22 | 0.32 | 0.27 | 0.17 | -1.67 | -1 | -1.16 | -1.94 | -0.21 | -0.63 | -0.71 | 0.81 |
| -0.51 | -0.47 | -0.47 | -0.53 | -0.55 | 0.1 | 0.06 | -0.57 | 0.16 | -0.88 | -1.02 | 0.15 | -0.42 | -0.72 | -0.8 | 0.6 |
| -0.39 | -0.51 | -0.51 | -0.4 | 0.46 | 0.74 | 0.68 | 0.41 | -0.95 | -0.43 | -0.5 | -1.12 | -0.03 | -0.67 | -0.75 | 0.99 |
| -0.44 | -0.56 | -0.56 | -0.46 | 0.15 | 0.61 | 0.55 | 0.11 | -0.24 | -0.98 | -1.13 | -0.3 | -1.12 | -1.14 | -1.25 | -0.1 |
| -0.54 | -0.8 | -0.82 | -0.56 | 0.57 | 0.89 | 0.83 | 0.51 | -0.7 | -0.99 | -1.14 | -0.83 | -0.42 | -0.01 | -0.05 | 0.6 |
| -0.68 | -1.22 | -1.26 | -0.71 | -0.27 | -1.25 | -1.25 | -0.3 | -0.82 | -0.5 | -0.58 | -0.97 | 0.04 | -1.14 | -1.25 | 1.06 |
| -1.2 | -1.13 | -1.16 | -1.26 | -0.46 | 0.62 | 0.56 | -0.49 | -0.24 | -1.48 | -1.7 | -0.3 | -0.13 | -1.19 | -1.3 | 0.89 |
| 0.09 | -1.25 | -1.29 | 0.11 | 0.22 | 0.72 | 0.66 | 0.17 | 0.34 | -1.14 | -1.32 | 0.36 | -0.15 | -1.29 | -1.4 | 0.87 |
| -0.6 | -0.55 | -0.55 | -0.62 | 0.74 | 0.59 | 0.53 | 0.68 | -0.59 | -1.13 | -1.3 | -0.7 | -0.32 | -0.71 | -0.79 | 0.7 |


|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| -0.29 | -0.8 | -0.81 | -0.3 | -0.57 | 0.65 | 0.59 | -0.59 | 0.54 | -0.19 | -0.23 | 0.59 | -0.38 | -1.93 | -2.08 |  |
| -0.66 | -1.15 | -1.18 | -0.69 | 0.36 | 0.42 | 0.37 | 0.31 | -0.16 | -0.23 | -0.27 | -0.21 | -0.49 | -0.73 | -0.81 |  |
| -0.38 | -0.7 | -0.71 | -0.39 | 0 | 0.51 | 0.46 | -0.04 | -0.09 | -1.2 | -1.38 | -0.13 | -0.04 | -0.63 | -0.71 | 0.93 |
| -1.43 | -0.65 | -0.66 | -1.5 | -0.24 | 1.03 | 0.96 | -0.27 | -0.21 | -1.13 | -1.31 | -0.27 | -0.12 | -0.72 | -0.8 |  |
| -0.5 | -0.67 | -0.68 | -0.52 | 0.45 | 0.68 | 0.62 | 0.4 | -0.52 | -2.05 | -2.36 | -0.62 | -0.08 | -0.67 | -0.75 | 0.94 |
| -1.46 | -0.56 | -0.56 | -1.53 | -0.66 | 0.47 | 0.42 | -0.68 | -1.23 | -1.13 | -1.3 | -1.44 | -0.01 | -1.14 | -1.25 | 1.01 |
| -1.15 | -1.18 | -1.22 | -1.21 | -0.15 | 1.05 | 0.98 | -0.19 | -1.32 | -1.25 | -1.44 | -1.54 | -0.65 | -0.01 | -0.05 |  |
| -0.86 | -0.77 | -0.78 | -0.9 | -0.26 | 1.02 | 0.95 | -0.29 | 0.17 | -1.06 | -1.22 | 0.16 | 0.1 | -1.14 | -1.25 |  |
| -0.69 | -0.78 | -0.79 | -0.72 | 0.22 | 0.32 | 0.27 | 0.17 | -0.47 | -1.71 | -1.97 | -0.57 | -0.31 | -1.19 | -1.3 |  |
| -0.16 | -0.97 | -0.99 | -0.16 | 0.28 | 1.09 | 1.02 | 0.23 | 0.14 | -1.18 | -1.36 | 0.13 | -0.06 | -1.29 | -1.4 |  |
| -0.41 | -0.71 | -0.72 | -0.42 | 0.46 | 1.21 | 1.14 | 0.41 | 0.24 | -1 | -1.16 | 0.24 | -0.62 | -0.71 | -0.79 | 0.96 |
| -0.38 | -1.01 | -1.04 | -0.39 | 0.74 | 1.21 | 1.14 | 0.68 | -0.25 | -1 | -1.16 | -0.32 | -0.25 | -1.93 | -2.08 |  |
| -0.74 | -0.81 | -0.83 | -0.77 | 0.22 | 0.26 | 0.21 | 0.17 | -0.81 | -1.58 | -1.82 | -0.96 | -0.24 | -0.73 | -0.81 | 0.77 |
| -0.53 | -0.58 | -0.58 | -0.55 | -0.22 | 0.84 | 0.78 | -0.25 | -0.1 | -1.28 | -1.48 | -0.14 | -0.27 | -0.63 | -0.71 | 0.78 |
| -0.37 | -0.79 | -0.8 | -0.38 | 0.37 | 0.42 | 0.37 | 0.32 | 0 | -1.27 | -1.46 | -0.03 | -0.51 | -0.72 | -0.8 | 0.51 |
| -0.95 | -0.8 | -0.82 | -1 | 0.13 | 0.82 | 0.76 | 0.09 | -0.53 | -1.05 | -1.21 | -0.64 | 0.03 | -0.67 | -0.75 | 1.05 |
| -0.6 | -0.73 | -0.74 | -0.62 | 0.12 | 1.1 | 1.03 | 0.08 | -0.24 | -1 | -1.16 | -0.3 | -0.36 | -1.14 | -1.25 | 0.66 |
| -1.36 | -0.79 | -0.8 | -1.43 | 0.32 | 0.45 | 0.4 | 0.27 | -1.53 | -0.71 | -0.82 | -1.78 | -0.17 | -0.01 | -0.05 | 0.85 |
| -0.35 | -0.74 | -0.75 | -0.36 | -0.15 | 0.65 | 0.59 | -0.19 | -0.7 | -1.41 | -1.62 | -0.83 | -0.21 | -1.14 | -1.25 | 0.81 |
| -0.46 | -0.6 | -0.6 | -0.48 | 0.18 | 0.59 | 0.53 | 0.13 | -0.4 | -0.56 | -0.65 | -0.49 | -0.09 | -1.19 | -1.3 | 0.93 |
| -0.18 | -0.71 | -0.72 | -0.18 | 0.31 | 0.7 | 0.64 | 0.26 | -0.21 | -1 | -1.15 | -0.27 | -0.14 | -1.29 | -1.4 | 0.88 |
| -0.14 | -0.06 | -0.03 | -0.14 | -0.63 | 0.52 | 0.47 | -0.65 | 0.73 | -0.52 | -0.61 | 0.81 | 0.81 | -0.71 | -0.79 | 1.83 |
| 0.07 | 0.16 | 0.2 | 0.08 | -0.07 | 0.27 | 0.22 | -0.11 | -0.51 | -2.18 | -2.5 | -0.61 | -0.69 | -1.93 | -2.08 | 0.33 |
| -0.72 | 0.08 | 0.12 | -0.75 | 0.31 | 1.16 | 1.09 | 0.26 | -0.37 | 0.52 | 0.58 | -0.45 | 1.12 | -0.73 | -0.81 | 2.14 |
| 0.06 | 0.35 | 0.4 | 0.07 | 0.76 | -1.19 | -1.2 | 0.7 | -0.7 | 0.69 | 0.78 | -0.83 | 1.28 | -0.63 | -0.71 | 2.3 |


|  | 0.58 | 0.31 | 0.36 | 0.62 | -0.1 | 0.78 | 0.72 | -0.14 | 0.8 | 0.69 | 0.78 | 0.89 | 0.32 | -0.72 | -0.8 | 1.34 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1.11 | -0.44 | -0.43 | 1.18 | 0.73 | 1.15 | 1.08 | 0.67 | 0.72 | 0.68 | 0.77 | 0.79 | 0.57 | -0.67 | -0.75 | 1.59 |
|  | -0.01 | 0.05 | 0.08 | 0 | 0.5 | 1.2 | 1.13 | 0.45 | 1.14 | 0.48 | 0.54 | 1.27 | 1.05 | -1.17 | -1.28 | 2.07 |
|  | 0.85 | 0.27 | 0.32 | 0.91 | 1.05 | 1.02 | 0.95 | 0.98 | 0.31 | 0.97 | 1.1 | 0.32 | 1.4 | 1.16 | 1.18 | 2.42 |
|  | -0.71 | 0.72 | 0.79 | -0.74 | 0.3 | 0.51 | 0.46 | 0.25 | 0.87 | 1.24 | 1.41 | 0.97 | 0.61 | 0.29 | 0.26 | 1.63 |
|  | 0.69 | 1.39 | 1.5 | 0.74 | 1.2 | 0.6 | 0.54 | 1.13 | 0.69 | 0.7 | 0.79 | 0.76 | 0.66 | 0.58 | 0.57 | 1.68 |
|  | -0.8 | -0.06 | -0.03 | -0.84 | -0.18 | 0.87 | 0.81 | -0.21 | 0.89 | 0.9 | 1.02 | 0.99 | 0.66 | 0.67 | 0.66 | 1.68 |
|  | -0.44 | 1.35 | 1.46 | -0.46 | 0.65 | 0.31 | 0.26 | 0.59 | 0.63 | 0.84 | 0.95 | 0.69 | 0.61 | 1.49 | 1.53 | 1.63 |
|  | 0.98 | 1.37 | 1.48 | 1.05 | 0.23 | 1.55 | 1.47 | 0.18 | 0.49 | -1.08 | -1.25 | 0.53 | 1.11 | 0.31 | 0.28 | 2.13 |
|  | -0.89 | 0.52 | 0.58 | -0.93 | 0.67 | -0.27 | -0.3 | 0.61 | 1.3 | 0.47 | 0.53 | 1.46 | 0.98 | 0.93 | 0.94 | 2 |
|  | -0.87 | 0.34 | 0.39 | -0.91 | -0.18 | 0.59 | 0.53 | -0.21 | -1.03 | 0.91 | 1.03 | -1.21 | 0.32 | 0.74 | 0.74 | 1.34 |
|  | 0.15 | 0.4 | 0.45 | 0.17 | 0.35 | -0.46 | -0.49 | 0.3 | -0.16 | -0.15 | -0.18 | -0.21 | 0.42 | -1.11 | -1.21 | 1.44 |
|  | -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.23 | -0.09 | -0.12 | 1.38 | 1.08 | -1.12 | -1.22 | 2.1 |
|  | 0.31 | 0.79 | 0.87 | 0.34 | 0.38 | -1.38 | -1.38 | 0.33 | 1.05 | -0.09 | -0.12 | 1.17 | 0.83 | -1.12 | -1.22 | 1.85 |
|  | -0.48 | 0.85 | 0.93 | -0.5 | 0.39 | -1.65 | -1.64 | 0.34 | 1.21 | -0.09 | -0.12 | 1.35 | 0.8 | -0.8 | -0.89 | 1.82 |
|  | -1.36 | -0.8 | -0.82 | -1.43 | -0.9 | -1.36 | -1.36 | -0.91 | 0.46 | -1.13 | -1.3 | 0.5 | -0.1 | -1.19 | -1.3 | 0.92 |
|  | -1.87 | -0.98 | -1 | -1.97 | -0.5 | -0.41 | -0.44 | -0.53 | -0.87 | -0.89 | -1.03 | -1.03 | -0.87 | -1.3 | -1.41 | 0.15 |
|  | -0.48 | -1.19 | -1.23 | -0.5 | -0.56 | -1.17 | -1.18 | -0.58 | 0.3 | -1.08 | -1.25 | 0.31 | -1.36 | -0.22 | -0.27 | -0.34 |
|  | -1.89 | -0.7 | -0.71 | -1.99 | -2.73 | -1.39 | -1.39 | -2.69 | -1.56 | -0.89 | -1.03 | -1.81 | -0.86 | -0.57 | -0.64 | 0.16 |
| 1040 | -1.82 | -0.59 | -0.59 | -1.92 | -0.71 | -1.42 | -1.42 | -0.73 | -1.84 | -0.73 | -0.85 | -2.13 | -0.21 | -1.38 | -1.5 | 0.81 |
| 1041 | -0.55 | -0.71 | -0.72 | -0.57 | -0.43 | 0.3 | 0.25 | -0.46 | -0.87 | -0.09 | -0.12 | -1.03 | -0.7 | -1.14 | -1.25 | 0.32 |
| 1042 | -1.05 | -0.65 | -0.66 | -1.1 | 0.29 | 0.81 | 0.75 | 0.24 | -1.82 | -1.09 | -1.26 | -2.11 | -0.13 | -0.01 | -0.05 | 0.89 |
| 1043 | -1.32 | -0.91 | -0.93 | -1.39 | -0.5 | 0.6 | 0.54 | -0.53 | -0.75 | -0.16 | -0.19 | -0.89 | -0.54 | -1.14 | -1.25 | 0.48 |
| 1044 | -0.3 | -0.49 | -0.49 | -0.31 | 0.19 | -0.15 | -0.18 | 0.14 | -0.15 | -1.14 | -1.32 | -0.2 | -0.02 | -1.19 | -1.3 | 1 |
| 1045 | -0.85 | -0.83 | -0.85 | -0.89 | -0.1 | 0.76 | 0.7 | -0.14 | -0.42 | -1.01 | -1.17 | -0.51 | 0.15 | -1.29 | -1.4 | 1.17 |


| 1046 | -0.36 | -1.16 | -1.2 | -0.37 | 0.24 | 1.39 | 1.31 | 0.19 | -0.54 | -0.88 | -1.02 | -0.65 | -0.34 | -0.71 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1047 | -0.27 | -0.12 | -0.1 | -0.28 | -0.23 | 0.3 | 0.25 | -0.26 | -0.32 | -1.09 | -1.26 | -0.4 | -0.16 | -1.93 |
| 1048 | -0.82 | -0.29 | -0.27 | -0.86 | -0.74 | 0.44 | 0.39 | -0.76 | 0.27 | -1.11 | -1.28 | 0.28 | -0.36 | -0.73 |
| 1049 | -0.46 | -0.78 | -0.79 | -0.48 | -0.41 | 0.06 | 0.02 | -0.44 | -0.2 | -2.15 | -2.47 | -0.26 | -1.18 | -0.63 |
| 1050 | -0.94 | -0.96 | -0.98 | -0.98 | -0.42 | 0.86 | 0.8 | -0.45 | 0.15 | -1.37 | -1.58 | 0.14 | -0.16 | -0.72 |
| 1051 | -0.8 | -0.57 | -0.57 | -0.84 | 0.78 | 0.34 | 0.29 | 0.72 | -1.61 | -1.78 | -2.05 | -1.87 | -0.56 | -0.67 |

Appendix 10
Computation of transformation equations
For year 2011
$\mu_{N}=-0.01, \mu_{W}=0.02, \sigma_{N}=1.065, \sigma_{W}=1.127$
On substitution of these values into equation 4.1, we have

$$
x_{N}=\frac{1.065}{1.127} x_{W}+\left(-0.01-\frac{1.065}{1.127} 0.02\right)
$$

$x_{N}=\left(\operatorname{Tr} x_{N}\right)=0.945 x_{W}-0.03------------e q n 4.3$
And on substitution of the means and standard deviationsof examinees ability scores on NECO and WAEC tests into eqn 4.2 we have
$x_{W}=\frac{1.127}{1.065} x_{N}+\left(0.02-\frac{1.127}{1.065}(-0.01)\right)$
$x_{W}=\left(\operatorname{Tr} x_{W}\right)=1.058 x_{N}+0.01------------$ eqn 4.4
For year 2012
$\mu_{N}=0.03, \mu_{W}=-0.01, \sigma_{N}=1.103, \sigma_{W}=1.012$
On substitution of these values into equation 4.1 we have
$x_{N}=\frac{1.103}{1.072} x_{W}+\left[0.03-\frac{1.103}{1.072}(-0.01)\right]$
$x_{N}=\left(\operatorname{Tr} x_{N}\right)=1.029 x_{W}+0.04-------$ eqn 4.5
And on substitution of the values into eqn 4.2 we have,
$x_{W}=\frac{1.072}{1.103} x_{N}+\left(-0.01-\frac{1.072}{1.103} 0.03\right)$
$x_{W}=\left(\operatorname{Tr} x_{W}\right)=0.972 x_{N}-0.04----------$ eqn 4.6
For year 2013
$\mu_{N}=0.02, \mu_{W}=-0.01, \sigma_{N}=0.973, \sigma_{W}=1.113$, on substitution of these values into eqn 4.1 we have,

$$
x_{N}=\frac{0.973}{1.113} x_{W}+\left[0.02-\frac{0.973}{1.113}(-0.01)\right]
$$

$x_{N}=\left(\operatorname{Tr} x_{N}\right)=0.874 x_{W}+0.01------------$ eqn 4.7
And substitution of the means and standard deviations of examinees ability scores on NECO and WAEC mathematics tests into eqn 4.2 we have;
$x_{W}=\frac{1.113}{0.973} x_{N}+\left(-0.01-\frac{1.113}{0.973} 0.02\right)$
$x_{W}=\left(\operatorname{Tr} x_{W}\right)=1.144 x_{N}-0.03---------$ eqn 4.8, and
for year 2014
$\mu_{N}=0.05, \mu_{W}=0.01, \sigma_{N}=0.988, \sigma_{W}=1.043$, on substitution of these values into eqn 4.1 we have;

$$
x_{N}=\frac{0.988}{1.043} x_{W}+\left(0.05-\frac{0.988}{1.043} 0.01\right)
$$

$x_{N}=\left(\operatorname{Tr} x_{N}\right)=0.947 x_{W}+0.04----------$ eqn 4.9
And on substitution of the means and standard deviations of examinees ability scores on NECO and WAEC mathematics tests into eqn 4.2 we have;
$x_{W}=\frac{1.043}{0.988} x_{N}+\left(0.01-\frac{1.043}{0.988} 0.05\right)$
$x_{W}=\left(\operatorname{Tr} x_{W}\right)=1.056 x_{N}-0.04---------$ eqn4.10
Where; $\left(\operatorname{Tr} x_{N}\right)$ is the ability score of an examinee on NECO mathematics test transformed to the scale of WAEC Mathematics test, and $\left(\operatorname{Tr} x_{W}\right)$ is the ability score of an examinee on WAEC Mathematics test transformed to the scale of NECO Mathematics test

Equations 4.3, 4.5, 4.7, and 4.9 represent the functions used for placing ability scores of examinees on NECO mathematics test of 2011, 2012, 2013, and 2014 unto the scale of the examinees ability scores on WAEC mathematics test of 2011, 2012, 2013, and 2014 respectively. While equations 4.4, 4.6, 4.8, and 4.10 represent the functions used for placing ability scores of examinees on WAEC Mathematics test of years 2011, 2012, 2013, and 2014 unto the scale of the ability scores of examinees on NECO Mathematics test of 2011, 2012, 2013, and 2104 respectively.

